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Abstract

Research on the effect of candidate sex on voting behaviour has found mixed results. In some countries, in some elections, candidate sex has been found to influence voting behaviour but the mechanisms underpinning this relationship are not fully understood. In the British context the issue of candidate sex has become politicised by party strategies relating to the selection of women candidates. Controversy over mechanisms to improve women’s representation raises the salience of candidate sex in the public mind and divides opinion on whether women better represent women’s interests. Accordingly, British Conservative party leader David Cameron’s attempts to improve the representation of women in the Parliamentary Conservative party in 2010 made the issue of candidate sex particularly electorally salient and also made it the first election where both main parties fielded substantial numbers of women candidates in winnable seats. Given this increase in the supply and salience of women candidates, 2010 provides a unique opportunity in the British case to model the effect of candidate sex on vote choice. We show that attitudes towards the descriptive representation of women had a significant impact on whether women supported women candidates in 2010. However, the group of women who are motivated by feminist concerns to vote for women candidates is relatively small (9%) which explains why a difference between men and women in support for women candidates is not evident at the aggregate level.
Introduction

A growing body of work on candidate traits shows that people with a given social characteristic tend to prefer candidates or leaders who share that characteristic (Campbell and Cowley 2014; Cutler 2002). However, the existing evidence for whether women vote for women is mixed. For example, Kathleen Dolan found that candidate sex was a driver of voting behaviour for the US House of Representatives in 1992, but not in 1994 or 1996 (Dolan 1998; Dolan 2001; Dolan 2004). Eric Smith and Richard Fox used pooled US data from 1988 to 1992 and found that well-educated women were more inclined to support women candidates in House but not Senate races (Smith and Fox 2001), and others have found that women are more likely to vote for women candidates only when they are perceived as being pro-feminist (Plutzer and Zipp 1996). By contrast Fulton (2014) found that women are not more likely to vote for women candidates in the US, but that male Independents are somewhat less likely to vote for them. Others have found little evidence whatsoever of an association between candidate gender and vote choice (McElroy and Marsh 2010).

These findings suggest that the link – if any – between candidate sex and voting behaviour is not straightforward, and is also highly context specific. It is not clear from previous research why candidate sex should matter in some elections but not others, and why women may differ in their support for women candidates. Moreover, previous research has largely been limited to the United States. In this article we build on this emerging literature by examining the impact of candidate sex in the British context. We do so with reference to key theories about political representation (Pitkin 1967).

Hannah Pitkin famously described a four-part typology of political representation; comprising of formal, descriptive, substantive and symbolic representation. Feminist research has largely focused on two components of Pitkin’s typology: descriptive and substantive
representation. Descriptive representation occurs where the representative stands for a group by virtue of sharing similar characteristics and substantive representation occurs where the representative seeks to advance a group’s policy preferences. If descriptive representation is important to voters then, given the underrepresentation of women in the British parliament (where just 22% of MPs elected in 2010 were women), we might expect women in the electorate to demand more women MPs. Equally we might expect women to demand better descriptive representation (by supporting increasing the number of women in the legislature) because they believe that women MPs will deliver substantive representation (by promoting women’s issues, preferences and interests). Alternatively women may have a preference for the greater descriptive representation of women simply from the position of fairness - as 52% of the adult population are women - women may believe any lesser figure in the legislature is simply not just (Phillips 1995). In fact previous research has shown that women in Britain would like to see more women MPs in the House of Commons, although they are not concerned about the sex of their own MP (Cowley 2013). However, it is unlikely that all women would share the desire for more women parliamentarians and many may be indifferent to the proportion of women in the legislature. Whereas some women may be very much in favour of having more women in politics; others may not place such a high value on increasing women’s representation. We would therefore not expect women who are unconcerned about the descriptive representation of women to trade off voting for their preferred party on the basis of candidate sex. But do women who hold strong feminist views about the importance of getting more women into politics (descriptive representation) put these views in to practice when they have the chance by voting for women candidates?

We suggest that the 2010 British election provides a particularly instructive case to test our hypotheses since (1) the availability of women candidates, in a range of constituencies, was much higher than in previous years, and (2) the issue of women’s
political representation was also more salient than in previous elections, largely as a result of
the intra party debate about all women shortlists - a debate which was present to a greater or
less extent in all three major parties. This combination of both supply and salience provides
the most likely contemporary context for candidate sex effects to matter in Britain. It also
provides us with an opportunity to link political discourse about descriptive representation at
the party level to public attitudes towards descriptive representation, and investigate the
impact on voting behaviour.

Candidate sex and voting behaviour

Over the last few years a growing body of research has examined the impact of candidate sex
on voters, and has addressed the question of whether women are more likely (and men less
likely) to support women candidates  – and if so why. The most straightforward possibility
relates to identity politics, where voters have a baseline preference for candidates of their own
sex (Sanbonmatsu 2002). As Johnston et al. (1992) argue "the more an agent resembles oneself
the more he or she might be expected reflexively to understand and act on one's own interests".
This can provide a powerful motivation to vote for candidates or leaders who embody our own
demographic characteristics. Evidence to support this contention has been found in other
contexts, specifically with respect to ethnicity - in the US (Sigelman et al. 1995; Tate 1993;
Terkildsen 1992)), Europe (Bergh and Bjørklund 2011; Teney et al. 2010), and India (Heath,
Verniers and Kumar Forthcoming) - but evidence with respect to gender has been far more
mixed, and candidate sex has been found to influence voting behaviour in some elections but
not others  (Dolan 1998; Dolan 2001; Dolan 2004; Dolan and Lynch 2014).

In addition to the possibility of a straightforward identity politics effect, where voters
prefer candidates of their own sex, the literature suggests preference for, or hostility to, women
candidates might be influenced by the perceived feminism of the candidate. In their analysis of
exit polls from the 1992 elections Eric Plutzer and John Zipp found evidence to suggest that a
simple pattern of identity politics, whereby women vote for women is not a sufficient explanation for who votes for women candidates (Plutzer and Zipp 1996). According to their evidence, it would seem that in a number of cases women are indeed more likely to vote for women, but this effect is most pronounced when the woman candidate is perceived as being pro-feminist. When candidates are perceived as "acting for," rather than merely "standing for," women, they stimulate voting based on gender identity (p34).

This line of work has been developed in a number of different directions to examine the character traits and perceptions of male and female candidates. Recent research suggests that voters tend to assess men and women candidates differently by using stereotypes about their perceived strengths and weaknesses (Cutler 2002; Huddy and Terkildsen 1993a; McDermott 1997; McDermott 1998; Schwindt-Bayer, Crisp and Malecki 2008). Gender stereotypes portray women as approachable and empathetic and men as assertive leaders (Alexander and Andersen 1993; Burrell 2008; Huddy and Terkildsen 1993b; Kahn 1996; King and Matland 2003; Lawless 2004; Leeper 1991; Paul and Smith 2008; Sapiro 1981/1982). Voters tend to use these gender stereotypes to assess men and women candidates’ policy competencies differently, with women seen as more competent on matters of health and education and men seen as more competent on matters of the economy and terrorism (Dolan 2010). Perhaps as a result of gender stereotypes equating women with higher levels of compassion and empathy studies have found that women candidates are perceived as more liberal than men running for election (Koch 2000; Koch 2002; McDermott 1997) and voters who are worried about levels of honesty and integrity in government have been more likely to vote for women candidates (Dolan 2004; McDermott 1998). Although women candidates often believe they have to be better than men to succeed (Lawless and Fox 2010) this is a misperception and as candidates become known to voters gender stereotypes play far less of a
role and recent evidence suggests that gender stereotypes in candidate evaluations have no
effect on vote choice (Brooks 2013; Dolan 2014a; Dolan 2014b; Dolan and Lynch 2014).

Studies in the US have also tended to find stronger patterns of candidate gender voting
among independents, who do not have such strong partisan ties and so may be more influenced
by candidate traits (e.g. Fulton 2014). Other studies have found that candidate gender voting is
shaped by partisan affiliation, and that people are less likely to vote for Republican women
suggests that this is because issue and trait stereotypes of women and of Democrats are
generally consistent with each other, while stereotypes of women and of Republicans are more
at odds with each other.

One area that has received less attention though is to do with public opinion towards
the question of women’s representation in elected office (Dolan 2010: 70). Previous work in
this area has examined public attitudes towards descriptive representation (Campbell, Childs
and Lovenduski 2010) and how these attitudes shape evaluations of women candidates
(Dolan 2010; Rosenthal 1995), but as yet there has been little scholarly work on how these
attitudes directly influence actual behaviour. In this article we contribute to this emerging
literature by examining whether public attitudes towards the descriptive representation of
women condition the extent to which candidate sex influences voting behaviour. We expect
voters who support increasing the descriptive representation of women to be more likely to
vote for women candidates than those who do not, but we expect the effect to be more
pronounced among women than men- as the issue is likely to be more salient to women.

Gender and the 2010 British General Election

In order to test our research question we focus on the 2010 British election. As set out above
previous research has found that the extent to which gender is politically salient depends
upon the political context. For example, Kathleen Dolan (2014: 24) argues that the role
candidate sex plays in voter calculations varies with the electoral context and is more important in some contests than others. She found that women were more likely to vote for women candidates in 1992 than other election years because of the high salience of gender in the 1992 election, which was commonly referred to as the 'year of the woman' (Dolan 2001; Dolan 2004). Plutzer and Zipp also single out 1992, because of the salience of gender in the election, thus they thought that should women ever vote for women then they would certainly do so in this year.

Although it is hard to think of a similar example where gender has played such a significant role in election reporting in the British case, the 2010 provides us with the best opportunity in recent times to do so. From 1997 the Labour party was easily the most feminist oriented political party in Britain but after his election as Conservative leader in 2005 David Cameron sought to challenge the Labour party on this terrain. Cameron made improving the representation of women in parliament a significant part of his bid to ‘modernise’ or decontaminate the Conservative brand. He needed to shake off the ‘nasty party’ image and embracing gender equality, along with environmentalism and other liberal values, was part of the strategy (Childs and Webb 2012). In his first speech as party leader Cameron stressed the importance of improving the representation of women in politics and talked about work-life balance as one of the most pressing issues facing contemporary British families. His controversial A-list policy –that required that women made up 50 per cent of candidates on the party’s preferred candidate list- was substantially watered down in response to a backlash from party members. However these internal divisions kept the issue on the media’s agenda.

In addition to controversies over the selection of women candidates the issue of women voters, or more specifically mothers’ voting preferences became a highly salient feature of the media coverage of the election. In the long campaign all three party leaders
participated in web-chats on the Mumsnet website at their own request and the newsframe ‘the mumsnet election’ was directly referred to in 43 national newspaper articles (starting with the Times 17.11.09). In fact Labour’s election campaign coordinator (Douglas Alexander) was on record saying ‘Labour needs to win back middle-income female voters with children in marginal seats’ (Guardian, 19.2.10). Each of the three main parties devoted considerable space in their manifestos to flexible working hours, parental leave, and pre-school childcare provision. The Conservatives pledged to make Britain the most family friendly country in Europe, and the Labour and Liberal Democrat manifestos also made specific reference to sex discrimination in the work place, and gender disparities in pay. The salience of ‘family issues’ and other traditional values is measured by the Manifestos Research Group over time (Volkens et al 2014). As Figure 1 shows, these issues were far more salient in 2010 than they had been previously.

FIGURE 1 ABOUT HERE

The 2010 election saw more women candidates and more women winners than ever before, creating a step-change in the partisan balance of women’s representation in parliament. As a result of selecting more women candidates in winnable seats the Conservatives more than doubled the number of Conservative women MPs, from 17 in 2005 to 48 women in 2010 (Campbell and Childs 2010). The effect of these changes in the supply of women candidates has important methodological implications. When only the Labour party put forward women candidates with any serious chance of winning, it was only possible to examine whether voters would offer support (or not) to the Labour party on the basis of the candidate’s sex. In 2010, for the first time parties other than Labour put forward women candidates in winnable seats, making it possible to try and untangle the relationship between candidate sex, political party and vote choice to a much greater extent than previously. Regardless of the salience of gender as an electoral issue, it would simply not be possible to sensibly carry out this type of
analysis on earlier British elections since too many of the women candidates only stood for one of the major parties.

Despite the controversy surrounding the selection of so many women as Conservative candidates, by and large those women who did stand for election did just as well as their male counterparts. Although male Conservative candidates tended to do slightly better than female Conservative candidates, both in those seats where the Tories were the incumbent (51.4% vs 49.7%) and in those seats where they were the challenger (28.7% vs. 26.3%), the differences are relatively small and may just reflect differences in the winnability of the seats for which men and women were selected. If we compare the average swing for men and women Conservative candidates (which controls for the winnability of the seats) then the differences are even smaller and do not reach conventional levels of statistical significance. Where women stood as the Conservative candidate the party did 3.4 percentage points better than they did in 2005, and where men stood the improvement was 3.9 percentage points. Thus there does not appear to be any evidence to suggest that putting forward so many women candidates cost the Conservatives votes. Similarly, there is little to suggest that women candidates for Labour did any worse than men. Women candidates for the Liberal Democrats did slightly worse than men in the seats which the party held in 2005: where women stood there was a 3.4 percentage point swing against the Lib Dems, and where men stood there was a 0.4 percentage point swing towards the Lib Dems but again this difference was not statistically significant and we should probably treat these results with caution since there were only 12 women candidates standing for the Lib Dems in seats they won in 2005.

**TABLE ONE ABOUT HERE**
These findings show that on average men and women candidates performed more or less equally well for each of the three main parties, which suggests that the over-representation of men in Parliament owes more to inequalities in the selection process than it does to discrimination by voters. However, as Dolan (2010) cautions, just because real-world election results suggest that women do not suffer disproportionately at the ballot box, because of their sex, we should not assume that this means that voter attitudes about gender are irrelevant to politics, or even irrelevant to vote choice. We can only infer so much from aggregate level data about individual behaviour, and individual’s attitudes towards descriptive representation and gender stereotypes may shape an individual’s vote choice decision when faced with a woman candidate.

**Materials and Methods**

Previous studies on candidate sex effects have been hampered by the small number of women candidates that run for office in winnable seats. This creates problem at the analysis stage, since inferences about whether or not women are more likely to support women candidates are often based on very small samples of women candidates. This increases the likelihood of outliers - or unique personalities – driving the results, and just as we would be unwise to infer too much about gender voting from whether someone supported Mrs Thatcher or not, we should also treat with caution the results of analysis based on relatively few women candidates, since prominent female politicians may in a sense transcend gender. The increased supply of women candidates in winnable seats at the 2010 election helps us to overcome this problem, and provides a valuable opportunity to study the effect of gender away from the glare of the media spot light on one or two key personalities, and examine how candidate gender influences voting behaviour across a range of different political contexts.

In order to test our hypothesis about the effect of candidate sex on voting behaviour we link data on candidate sex and constituency characteristics to cross-sectional survey data.
The individual-level data on voters come from the 2010 British Election Study. The survey was administered to a random sample of 3075 respondents in 205 constituencies. After linking the survey data to the candidate data our sample includes 154 women candidates (54 Labour; 58 Conservative; and 42 Liberal Democrat). This provides us with a robust sample to test the impact of candidate sex on voting behaviour.

**Attitudes towards descriptive representation**

Our key theoretical variable of interest are attitudes towards the descriptive representation of women. To measure these attitudes we draw on a special module in the British Election Study, which includes the following four items on gender equality:

1. Women are better representatives of women’s interests than men
2. Most men are better suited to politics than are most women
3. Women need to get more involved in politics to solve problems that concern them
4. A husband’s job is to earn the money; a wife’s job is to look after the home and family.

For each item respondents could agree strongly, agree, neither agree nor disagree, disagree, disagree strongly or refuse to give an answer. Descriptive statistics are shown in Table 2, which reports the mean for each item on a scale of 1 to 5, where 1 is strongly agree and 5 is strongly disagree. With respect to attitudes about descriptive representation, both men and women were more likely to agree than disagree with the statement that women are better representatives of women’s interests than men, and women need to get more involved in politics to solve problems that concern them (as illustrated by mean values greater than 3, where 3 indicates neither agree nor disagree). Women were significantly more likely to agree with these statements than men, particularly on the issue of women getting more involved in politics; where 67% of women agreed with the statement compared to 53% of men.
The other two items are less directly related to issues of descriptive representation – and tap more into attitudes about traditional gender roles. Both men and women were more likely to disagree than agree with the statements that men are better suited to politics than women, and a husband’s job is to earn the money (as illustrated by mean values less than 3). Once again women were significantly more likely than men to disagree with these statements. These results reflect the findings of Kathleen Dolan’s 2013 US study; she found that “women are more likely than men to have a baseline preference for women candidates, to want to see more women in office, to support gender parity in elected office and to support same-sex representation” (Dolan 2014b: 69)

**TABLE TWO ABOUT HERE**

In order to check whether these items do indeed tap into attitudes about descriptive representation and traditional gender roles specifically, rather than attitudes towards gender equality in general we undertake a principal components analysis of the responses to all four items (Table 3). These items load on to two clear dimensions (Campbell, Childs and Lovenduski 2010). As such rather than simply combining items relating to gender equality in the BES data we create two additive scales. The first combines items one and three to create a scale measuring attitudes to the descriptive representation of women. The second scale, measuring attitudes to traditional gender roles, combines items two and four. Each scale includes a range of valid responses from zero to eight, where eight represents a more feminist position: either high levels of support for the descriptive representation of women or opposition to traditional gender roles.

**TABLE THREE ABOUT HERE**

The mean score for women on the descriptive representation scale is 4.79 compared to a mean score of 4.49 for men (the sex difference is significant at the 0.01 level). The mean score for traditional gender roles is 5.34 for women and 5.03 for men (the sex difference is
significant at the 0.01 level). Our key hypothesis concerns attitudes towards descriptive representation: we expect voters who are in favour of the descriptive representation of women to be more likely to vote for women candidates, but we expect the effect have a more pronounced effect among women than men- as the issue is likely to be more salient to women. By way of comparison we also examine public attitudes towards traditional gender roles, which according to prior theory should be empirically unrelated to vote for women candidates (Brooks 2013; Dolan 2014a; Dolan 2014b; Dolan and Lynch 2014).

**Modelling the impact of candidate sex on vote choice**

Our main hypothesis is to investigate the link between attitudes towards descriptive representation and support for women candidates. Do women who hold strong feminist views about the importance of getting more women into politics put these views into practice when they have the chance by voting for women candidates? Or do partisan ties over-ride feminist principles?

There are a number of challenges involved in estimating the effect of candidate traits on voting behaviour in multi-candidate elections, which make common econometric choice models, such as multinomial logit, inadvisable (see Alvarez and Nagler 1998). One solution is the conditional logit model, which is “conditional on the characteristics of the choices” (Alvarez and Nagler 1998: 56) and so can take into account candidate characteristics in a parsimonious manner. In order to estimate the effects of candidate sex on voting behaviour we therefore specify a conditional logit model (Long and Freese 2005). The model uses stacked data, with three rows for each respondent, one for each party. The conditional logit model allows us to fit how the choice between party alternatives is affected by the characteristics of the candidates. In formal terms Long and Freese (2005: 307) specify the model as:

$$
\Pr(y_i = p | x_i, z_i) = \frac{exp(z_{ip}\gamma + x_i\beta_p)}{\sum_{j=1}^{J} exp(z_{ij}\gamma + x_i\beta_j)}
$$
Where $z_{im}$ contains values of the alternative specific variables for alternative $p$ and case $i$. In our model there are three alternatives for party $p$: Labour, Conservative and Liberal Democrat. Then $\Upsilon$ is a parameter indicating the effect of the alternative specific variables on the probability of voting for one party over another. In general, for each variable $z_k$, there are $J$ values of the variable for each case, but only the single parameter $\Upsilon_k$. This enables a single coefficient for the effects of candidate sex to be estimated regardless of the candidate’s party. We also tested for evidence of variation in these coefficients across parties. As in a multinomial logit model, $x_i$ contains case specific independent variables for case $i$, and $\beta_p$ contains coefficients for the effects on alternative $p$ relative to the base alternative (Labour).

The control variables present in the models for both Men and Women are Party Like Score (an eleven-point scale from 0 = strongly dislike to 10 = strongly like the party) which captures both positive and negative disposition to each of the parties; Contacted by party (1 if the respondent was contacted by the party during the campaign, or 0 otherwise) which captures mobilisation and campaign effects; Incumbency which captures incumbency advantage, and Distance from Contention (for third and lower placed parties this is the difference between the 2005 share of the vote for the party and that for the second placed party, and 0 if the party came first or second in the constituency) which captures supply side factors to do with the allocation of tickets to women candidates. In particular women selected for Labour from all-women-shortlists (AWS) tend to be selected in competitive seats more often than women candidates from other parties. Finally we control for Candidate sex (1 if the candidate is a woman, or 0 if a man) which is our key theoretical variable of interest. Descriptive statistics are shown in the appendix. All of these controls are alternative-specific, so their values can vary across the party options for a given respondent. This enables a single coefficient for each variable to be estimated regardless of party. In practice this is equivalent to assuming that the magnitude of the effect for each of these variables is constant across party options. We tested the validity of
this assumption by specifying party-interactions (not shown), none of which were significant.

The model also allows case specific variables as predictors. The attitudinal variables that we examine as a potential moderator of candidate sex effects are attitudes towards descriptive representation and attitudes towards traditional gender roles discussed above. Both of these variables are measured on a 0-8 scale, where higher values indicate a more feminist position. Where a control variable is case-specific (say X) two coefficients are estimated (Con*X and LD*X), for the effects on voting Conservative and Liberal Democrat respectively, each relative to the baseline Labour (just as a regular multinomial logit/probit model would have two such coefficients).\textsuperscript{vii} In order to test our main hypotheses we also specify an interaction effect between candidate sex and attitudes towards gender.

**Results**

The conditional logit model of vote choice for Women (Model 1) and Men (Model 2) who voted for the Conservatives, Labour or Liberal Democrats is shown in table five. The model includes controls for the main effects of candidate sex and attitudes towards descriptive representation and an interaction term between these attitudes and candidate sex.

The coefficients for party are alternative-specific intercepts. In both models the main effects of the control variables for competiveness, incumbency status, mobilisation and party attachment are significant, in the expected direction, and of a similar magnitude for both men and women. The probability of an individual voting for a party declines the further that party is from contention; increases with direct personal contact with a party worker; and – not surprisingly – increases the more positively predisposed someone feels towards the party in general. We also tested for party specific interactions with these variables. None were significant. This indicates that the magnitude of the effect for these variables is similar across party options.

The main effects for attitudes towards descriptive representation and traditional gender
roles are specified as party interactions (since these variables are case-specific). None of the main effects are significant. Over and above the factors already discussed, people who support descriptive representation or are opposed to traditional gender roles are not more likely to vote for one party rather than another. However, of more interest from a theoretical point of view are the interaction terms with candidate sex. Are women who think women MPs better represent women’s interests more likely to vote for women candidates? If women act on these views in the ballot box then they should differ in their propensity to vote for women. The interaction term between attitudes towards descriptive representation and candidate sex is statistically significant and in the expected direction for women (Model 1). This indicates that women who support the descriptive representation of women in politics are more likely to vote for a party when it puts forward a woman candidate than women who are not concerned about descriptive representation. This signifies a process whereby those women who think that women MPs better represent women’s interests actually put this into practice and, if given the chance, are more likely to vote for women candidates.

**TABLE FOUR ABOUT HERE**

By contrast, men’s attitudes towards the descriptive representation of women are unrelated to vote choice and unrelated to whether or not they vote for women candidates. Although some men hold feminist principles and think more women should be represented in parliament, they do not act on these principles in the ballot box and are not any more likely to vote for women candidates than men who are unconcerned about descriptive representation. But by the same token there is also little evidence that men punish women candidates, even when they hold strong views about traditional gender roles. From the relationship between attitudes towards traditional gender roles and support for women candidates, we can see that neither the main effects nor the interaction term are significant. This indicates that gender stereotypes do not affect whether or not men vote for women candidates, and even those men who think
that women should stay at home do not sacrifice their preferred party by voting for a man rather than a woman. This is consistent with recent research in the US (Brooks 2013; Dolan 2014a; Dolan 2014b; Dolan and Lynch 2014).

We also investigated the possibility that men reacted differently to women candidates from different parties, and reacted differently to women candidates by the incumbency status of the party or the competitiveness of the party. In neither scenario did we find evidence of any significant differences, for either men or women. Unlike in the USA, where some studies have found that people are less likely to vote for Republican women candidates than Democrats (King and Matland 2003; Lawless and Person 2008), we do not find any evidence to suggest that people are less likely to vote for Conservative women candidates than Labour. We also tried a number of alternative specifications of the model, where distance from contention was replaced by a measure of marginality but this did not affect the substance of our findings either. We also examined the impact of party identification in addition to and in place of party affect, but this did not make any difference. Unlike in the USA, those people who do not identify with a particular political party are not more likely to vote for (or against) a candidate based on sex. Lastly, it could be argued that our controls are too strong, and that candidate sex may have an indirect effect on vote choice via our measures of party affect. Perhaps if, for example, the Conservatives put forward a woman candidate, individuals who hold traditional views on gender roles would become disaffected with the party and vote for an alternative. In this scenario the inclusion of party affect in our models may wash out the effects related to candidate sex. In order to consider this possibility we ran the models again without the party affect variable, but the main results of our analysis did not change.

These findings lend support to our main hypothesis that women’s attitudes towards descriptive representation are linked to whether or not they support women candidates. These findings are robust to a number of different specifications. We can get a clearer idea of the
magnitude of these effects by inspecting how the likelihood of voting for women candidates varies according to a woman’s level of support for descriptive representation. Generating marginal effects of interaction terms in non-linear models is not straightforward, and an alternative, and perhaps more meaningful approach, is to exponentiate the coefficients and interpret the interaction in terms of the natural metric of the model (Buis 2010). For conditional logit models these are the odds ratios. The odds ratio for candidate sex is 0.14, which means that the odds of voting for a party is 0.14 times lower when the party puts forward a female candidate than when it puts forward a male candidate. Since there is an interaction effect between candidate sex and attitudes towards descriptive representation, this effect of candidate sex refers to women who are completely opposed to descriptive representation (that is those women who scored zero on the 0-8 scale, of which there are relatively few). Nonetheless, the magnitude of this effect is quite substantial. Within this category we expect to find 1 woman voting for a party when it puts forward a woman candidate for every seven women voting for that party when it puts forward a man.

The interaction effect tells us by how much the effect of candidate sex on vote choice differs by women’s level of support for descriptive representation, but does so in multiplicative terms. The results also show that this interaction is significant. The odds ratio for the interaction term is 1.36 which means that the odds of voting for a party if it puts forward a woman candidate is 1.36 times higher for every point increase on the descriptive representation scale. So, for a woman who holds roughly the mean level of support for descriptive representation (five on the 0-8 scale) the odds of voting for a party if it puts forward a woman candidate are 0.65 times less than when it puts forward a male candidate. This means that within this category we expect to find about two women voting for a party when it puts forward a woman candidate for every three women who vote for that party when it puts forward a man. However, for women who exhibit a high level of support for descriptive representation (that is women who
scored eight on the scale) the odds of voting for a party are 1.6 times higher if it puts forward a woman candidate than a male candidate. This means that within this category we expect to find about three women voting for a party when it puts forward a woman candidate for every two women who vote for that party when it puts forward a man.

In substantive terms, the likelihood of voting for women candidates varies quite considerably by level of support for descriptive representation. Women with strong feminist principles about the representation of women are much more likely to vote for a party that puts forward a woman candidate than women who do not really care about the issue. However, even among feminist women, the effect of candidate sex on vote choice is not particularly strong, and relatively few women are prepared to compromise on their preferred party simply due to the sex of the candidate. Moreover, the size of the feminist group to which this positive effect of candidate sex refers is also relatively small. Only about 9% of women hold the sort of strong feminist principles that are associated with the greater likelihood of voting for women candidates. Part of the reason then that simplistic women-voting-for-women effects are not found is therefore because the effect only applies to a (relatively) small group of voters.

Discussion

Previous research in the USA shows candidate sex effects vary from election to election. We suggest that part of the explanation for these inconsistent results is to do with the salience of gender (and perhaps more specifically candidate gender) at the time of the election. In this respect 2010 represents something of an unusual context by British standards. There were not only many more women candidates than in previous elections, but the salience of candidate gender was also greater than in the past. Intra party dispute about the selection of women candidates bought the issue to the fore, and perhaps activated the political salience of attitudes towards descriptive representation.
Indeed, preliminary analysis of past elections in Britain – where gender was not such a salient feature of the political landscape – confirms this point. With respect to 2005 and 2001 we do not find any evidence of candidate sex effects on voting, although we should perhaps treat the results with a degree of caution since the number of women candidates was relatively small, particularly on the side of the Conservative and Liberal Democrats. In this respect the recently held 2015 election offers a valuable opportunity to further study these issues, as both the number of women candidates and the number of women elected to office were historically high. Future research may therefore wish to explore the different ways in gender is salient as a political issue – and how the politicisation of gender translates into voting preferences.

Thus by British standards 2010 was quite an unusual election, as the two main parties combined fielded more women candidates in winnable seats than ever before. In 2010 the political rhetoric about gender was closely linked to issues of descriptive representation, and how women felt about this issue appears to have exerted quite a strong impact on their voting behaviour. Our results from a series of conditional logit models show that in 2010 women are more likely to vote for women candidates, but only when they hold positive views about the value of descriptive representation of women. However, the group of women who held these views was relatively small (just 9%) and we suggest that this might explain why there tends to be little evidence of women voting for women at the aggregate level. Exceptional circumstances where issues of gender equality become highly politicized and extremely electorally salient, such as the 1992 elections in the United States, may yield a situation where more women are motivated to vote by feminist concerns but our research suggests that in a typical election this group will be small. By contrast, men’s attitudes towards descriptive representation are unrelated to whether or not they vote for women candidates. For both sexes, gender stereotypes are not related to vote choice and are not related to support for
women candidates. These results are robust to a number of different specifications and are not mediated by political party.

The implications of these findings are three-fold. When gender is politicised at the party level it has the potential to influence how voters participate in the political process. But it does not have a uniform impact. Women do not react in a homogenous way to the sex of the candidate and the assumption that women are more likely to vote for women candidates is therefore too simplistic and assumes a uniformity of motivations for doing so which are not present in practice. This goes against a straightforward identity politics effect. Secondly, women who hold strong feminist principles act on these principles in the ballot box, and are more likely to vote for a woman candidate when given the chance to do so. There is thus a link between how women think about the issue of descriptive representation and how they actually behave in practice. Those women who think there should be more women in politics go out of their way to make it happen, even if it means compromising on their preferred party. We suggest that the salience of gender at the party level bought these issues to the fore. Perhaps if it had been framed in a different way we would see different mechanisms at work.

Thirdly, there is no evidence that women candidates are discriminated against at the ballot box. Even men with strong traditional views, who think that women’s place is in the home, do not trade off voting for their preferred party when faced with a woman candidate, whatever party that woman stands for. Related to this last point there is little evidence that support for women candidates is affected by the party for which they stand. In the USA a number of studies have found that Republican women candidates tend to face a backlash from voters who hold traditional gender stereotypes. However, no such trend is evident in the UK, and Conservative women are no more or less likely to get punished at the ballot box than women candidates from either of the other two main parties. One potential reason for this is that in the UK each of the main parties are relatively close to each other on gender equality
issues, and so there is not the same sort of positional difference between the parties that there is in the USA (see Dolan 2004).
Figures and Tables

Figure 1: Salience of ‘family issues’ and other traditional values in party manifestos

Source: Manifesto Research Group. Item per603.
<table>
<thead>
<tr>
<th>Party</th>
<th>Incumbency</th>
<th>Candidate sex</th>
<th>% vote for party</th>
<th>Change in % vote for party 2005-2010</th>
<th>p-values&lt;sup&gt;viii&lt;/sup&gt;</th>
</tr>
</thead>
<tbody>
<tr>
<td>Conservative</td>
<td>Lost in 2005</td>
<td>women</td>
<td>26.35</td>
<td>3.42</td>
<td>.16</td>
</tr>
<tr>
<td></td>
<td></td>
<td>men</td>
<td>28.69</td>
<td>3.94</td>
<td></td>
</tr>
<tr>
<td></td>
<td>Won in 2005</td>
<td>women</td>
<td>49.79</td>
<td>3.98</td>
<td>.80</td>
</tr>
<tr>
<td></td>
<td></td>
<td>men</td>
<td>51.35</td>
<td>3.80</td>
<td></td>
</tr>
<tr>
<td>Labour</td>
<td>Lost in 2005</td>
<td>women</td>
<td>19.21</td>
<td>-7.35</td>
<td>.86</td>
</tr>
<tr>
<td></td>
<td></td>
<td>men</td>
<td>15.08</td>
<td>-7.45</td>
<td></td>
</tr>
<tr>
<td></td>
<td>Won in 2005</td>
<td>women</td>
<td>41.59</td>
<td>-6.36</td>
<td>.21</td>
</tr>
<tr>
<td></td>
<td></td>
<td>men</td>
<td>43.65</td>
<td>-5.46</td>
<td></td>
</tr>
<tr>
<td>Liberal Democrat</td>
<td>Lost in 2005</td>
<td>women</td>
<td>21.60</td>
<td>1.13</td>
<td>.69</td>
</tr>
<tr>
<td></td>
<td></td>
<td>men</td>
<td>20.52</td>
<td>0.96</td>
<td></td>
</tr>
<tr>
<td></td>
<td>Won in 2005</td>
<td>women</td>
<td>41.14</td>
<td>-3.11</td>
<td>.10</td>
</tr>
<tr>
<td></td>
<td></td>
<td>men</td>
<td>46.56</td>
<td>0.41</td>
<td></td>
</tr>
</tbody>
</table>
Table 2: Attitudes towards gender equality

<table>
<thead>
<tr>
<th></th>
<th>Men</th>
<th>Women</th>
<th>T-test</th>
</tr>
</thead>
<tbody>
<tr>
<td>Women are better representatives of women’s interests than men</td>
<td>2.95</td>
<td>2.86</td>
<td>2.13*</td>
</tr>
<tr>
<td>Most men are better suited to politics than are most women</td>
<td>3.38</td>
<td>3.57</td>
<td>-4.24***</td>
</tr>
<tr>
<td>Women need to get more involved in politics to solve problems that concern them</td>
<td>2.55</td>
<td>2.34</td>
<td>5.52***</td>
</tr>
<tr>
<td>A husband’s job is to earn the money; a wife’s job is to look after the home and family.</td>
<td>3.65</td>
<td>3.80</td>
<td>-2.95***</td>
</tr>
</tbody>
</table>

Notes: Items are measured on a 1-5 scale; where 1 = strongly agree and 5 = strongly disagree. * denotes p <0.05; ** p<0.01; *** p<0.005
Table 3: Principal Components Analysis of attitudes towards gender equality

<table>
<thead>
<tr>
<th></th>
<th>Component 1</th>
<th>Component 2</th>
</tr>
</thead>
<tbody>
<tr>
<td>Women MPs represent women’s interests</td>
<td>.282</td>
<td>.778</td>
</tr>
<tr>
<td>Men better suited politics than women</td>
<td>.842</td>
<td>-.145</td>
</tr>
<tr>
<td>Women get more involved in politics</td>
<td>.026</td>
<td>.831</td>
</tr>
<tr>
<td>husband earn wife care home &amp; family</td>
<td>.828</td>
<td>-.144</td>
</tr>
<tr>
<td></td>
<td>Model 1: Women</td>
<td>Model 2: Men</td>
</tr>
<tr>
<td>------------------------</td>
<td>---------------</td>
<td>--------------</td>
</tr>
<tr>
<td><strong>Party (baseline=Labour)</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Conservative</td>
<td>0.97 (0.99)</td>
<td>4.68 (4.99)</td>
</tr>
<tr>
<td>Liberal Democrats</td>
<td>0.53 (0.54)</td>
<td>1.10 (1.30)</td>
</tr>
<tr>
<td>Incumbent</td>
<td>1.66*** (0.27)</td>
<td>2.59*** (0.49)</td>
</tr>
<tr>
<td>Distance from contention</td>
<td>0.95*** (0.01)</td>
<td>0.96*** (0.01)</td>
</tr>
<tr>
<td>Party like</td>
<td>3.08*** (0.29)</td>
<td>3.30*** (0.29)</td>
</tr>
<tr>
<td>Contact</td>
<td>1.93* (0.57)</td>
<td>2.00 (0.79)</td>
</tr>
<tr>
<td>Female candidate</td>
<td>0.14 (0.15)</td>
<td>1.73 (2.35)</td>
</tr>
<tr>
<td>Descriptive representation*Con</td>
<td>0.99 (0.13)</td>
<td>0.76 (0.12)</td>
</tr>
<tr>
<td>Descriptive representation *LD</td>
<td>0.99 (0.13)</td>
<td>0.89 (0.14)</td>
</tr>
<tr>
<td>Descriptive representation *Female candidate</td>
<td>1.36* (0.19)</td>
<td>1.10 (0.19)</td>
</tr>
<tr>
<td>Gender role*Con</td>
<td>1.04 (0.11)</td>
<td>1.01 (0.10)</td>
</tr>
<tr>
<td>Gender role*LD</td>
<td>1.16 (0.12)</td>
<td>1.18 (0.13)</td>
</tr>
<tr>
<td>Gender role*Female candidate</td>
<td>1.03 (0.11)</td>
<td>0.90 (0.12)</td>
</tr>
<tr>
<td><strong>Log Likelihood</strong></td>
<td>-273</td>
<td>-217</td>
</tr>
</tbody>
</table>

**Notes:** Models 1 contains 2098 observations and Model 2 contains 1765 observations. Standard error in parenthesis. * denotes p <0.05; ** p<0.01; *** p<0.005
References


Heath, Oliver, Giles Verniers, and Sanjay Kumar. Forthcoming. "Do Muslim voters prefer Muslim candidates? Co-religiosity and political support in India." Electoral Studies.


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i Mumsnet is a UK website and discussion forum [http://www.mumsnet.com/](http://www.mumsnet.com/). The typical mumsnet user is a professional mother who is in paid employment.

ii The MRG measures are based on content analyses of the programmes of the main political parties at every post-war election. The policy statements in each ('quasi-') sentence are classified into fifty-six policy categories over seven policy domains. The data shown in Figure 2 is simply the sum of the percentage of sentences mentioning family values and other traditional values (per603) across each of the parties. This policy category includes favourable mentions of traditional moral values; maintenance and stability of family; religion, prohibition, censorship and suppression of immorality and unseemly behaviour. Unfortunately it is not possible to disaggregate this category further so we should treat the results with some caution. However, in recent elections mentions of family have been much more common than mentions of the other categories, which have fallen out of use – so if anything the over-time trends should understate the extent to which family issues were particularly salient in 2010.

iii All individual level data analysed in this paper is taken from the British Election Study which is a publicly funded large national dataset. The British Election Study adheres to stringent ethical guidelines as set out by the British Economic and Social Research Council.

iv Data on candidate sex was collated by Pippa Norris and made available through the British Parliamentary database. [www.pippanorris.com](http://www.pippanorris.com)

v The British Election Study 2010 was conducted by Harold Clarke, David Sanders, Marianne Stewart and Paul Whiteley. [http://bes2009-10.org/](http://bes2009-10.org/)

vi There is no evidence that considering only candidates standing in British Election Study constituencies produces a biased sample when compared to the full population of parliamentary candidates. For example, in the full candidate data 48% of the seats were won by the Conservative party in 2010 and 41% by the Labour party. In the BES data (weighted using the post-election weight) 50% of respondents were drawn from constituencies where the Conservatives won in 2010 and 40% were drawn from constituencies that the Labour party won in 2010. In the full candidate data women made up 24% of Conservative
Candidates and 30% of Labour candidates. In the BES file respondents were drawn from constituencies where 24% of Conservative candidates and 28% of Labour candidates were women.

Since these variables are case specific the main effect on vote choice is specified as a party interaction. Thus the coefficient for the variable Descriptive representation*Conservative refers to the main effect of descriptive representation on voting Conservative relative to the baseline category Labour; and the coefficient for the variable Descriptive representation*Liberal Democrat refers to the main effect of descriptive representation on voting Liberal Democrat relative to Labour. These coefficients can thus be interpreted in much the same way as would be the case with a regular multinomial logit/probit model.

P-values from a between groups comparison of means, ANOVA.