Understanding Older Adults’ Labour Market Trajectories:
A Comparative Gendered Life Course Perspective

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ABSTRACT
The recent push to keep older adults in the labour force glosses over who is likely to follow what kind of employment trajectory and why. In this paper, we broaden understandings of later-life labour market involvement by applying a comparative gendered life course perspective. Our data come from the Survey of Health, Ageing and Retirement in Europe and the Health and Retirement Study, two representative panel studies of individuals aged 50-plus. Using a unique modeling strategy, we examine employment biographies for older women and men from four nations with diverse policy regimes (Germany, Italy, Sweden, and the US), along with their links to family experiences and earlier attachment to the labour force. We find that, in every nation, women prevail in groups representing a weak(er) attachment to the labour market and men in groups signifying a strong(er) attachment. However, this pattern is much stronger for Germany and Italy than for Sweden and the US. Similarly, both family experiences and prior employment matter more for later-life labour market involvement in Germany and Italy. Our findings demonstrate that older adults’ employment trajectories are gendered; moreover, there is evidence that they are influenced by policies related not only to paid work but also to caregiving, and by those affecting not only current decisions but also those made earlier in the life course.
INTRODUCTION
In the wake of population ageing and the baby boom cohort’s impending retirement, analysts have turned their attention to older adults’ labour force participation (e.g., OECD, 2006). Yet, the dominant conceptual emphases—state-funded incomes that ‘pull,’ and workplace barriers that ‘push,’ older adults out of the labour force—often gloss over life course considerations that illuminate who is likely to follow what kind of later-life labour market trajectory and why (Cooke, 2006).

Studies framed by a gendered life course perspective offer a corrective. They focus on long-term outcomes, the continuity between experiences at younger and older ages, the role of family ties, and the influence of institutional contexts (Moen & Flood, 2013). These factors especially shape the employment patterns of women and men (Gornick & Meyers, 2006). Yet, studies of older adults’ labour market trajectories and the social and institutional processes that shape them are surprisingly rare (Börsch-Supan et al., 2009). Most longitudinal research on later-life employment focuses on transitions—either withdrawal (e.g., Radl, 2013) or re-entry (e.g., Kail & Warner, 2013). While transitional events are important, they are but one element in a succession of labour market experiences that characterize the later years (Bowlby, 2007). This suggests that older adults’ employment patterns are best modeled as extended biographical sequences, using methods that take the entire series as the unit of analysis (Billari & Piccarreta, 2005).

In this study, we draw on a gendered life course perspective, and begin to flesh out the what, who, and why of long-term patterns of later-life labour market involvement. Employing a unique modeling strategy that captures biographical sequences over ages 50-69, we examine their gendered correlates across four contrasting institutional contexts: Germany, Italy, Sweden, and the United States (US). We ask:

1. How distributions across later-life labour market biographies differ for women and men in each country;
2. Whether male-female differences in the distribution of these trajectories are the (gendered) outcome of family and prior work experiences; and
3. How both these aspects vary across the four institutional contexts.

We begin with an outline of the gendered life course perspective, existing evidence, and our hypotheses. Following a description of our methods, the results section charts male-female differences in later-life labour market biographies for the four nations, then examines how family experiences and earlier attachment to the labour force contribute to these gender disparities. We conclude by discussing our findings and their policy implications.

BACKGROUND

The gendered life course

Researchers increasingly argue that older adults’ labour market involvement is best understood by drawing on a life course perspective (e.g., Alley, Putney, Rice, & Bengtson, 2010). This entails a focus on individuals’ long-term circumstances in a given domain, such as employment (Mayer, 2004). It also encompasses the notions of biographical continuity and linked lives—the ideas that older adults’ labour market trajectories follow from their paid work experiences at younger ages (Han & Moen, 1999a), and that those experiences are shaped by ties to family (Moen, 1996). Importantly, the division of labour by sex means that roles will often be gendered, at least earlier in the life course, with men assuming primary responsibility for breadwinning and women providing the bulk of care (Blossfeld & Drobnič, 2001).

The life course perspective also emphasizes that biographies are shaped by institutional contexts—most obviously, welfare regimes (Mayer, 2004). These can be distinguished by the relative roles of the state, markets, and families in providing life’s necessities (Esping-Andersen, 1990), and the model of family life implied by that balance (Korpi, Ferrarini, & Englund, 2013). The distinctions identified apply to the years when our sample members were in their prime working and childrearing years (Katz, 2010; Persky, 2011; Sundström, 2003). In corporatist regimes (typified by Germany), families are the primary providers of social welfare. The state supports a breadwinner-caregiver model of family life: full male employment with fairly
generous benefits for periods of non-employment, and primary caregiving for women. Wives and mothers are assumed to work for pay only secondarily, if at all, and child care spaces are limited in number and hours covered. In *southern European* nations (typified by Italy), breadwinner-caregiver families also play a central role in the provision of social welfare, leading to marked gender differences in employment (Siaroff, 1994). Social spending is generally low, and, at least historically, families have often maintained themselves as multi-generational households in which women care for one or more male breadwinners (Gal, 2010). In *liberal* regimes (typified by the US), social spending is also low, but the distribution of resources is left to the market. The state passively backs a market-oriented dual-earner model, in which the majority of individuals (male and female) work for pay throughout their lives despite a dearth of public supports for caregiving. *Social democratic* nations (typified by Sweden) are distinguished by the state’s prominent role in the distribution of resources and care provision. They actively support a dual-earner model—offering comprehensive publicly-funded caregiving services, fairly generous benefits for periods of non-employment, flexible work arrangements, and better-quality part-time jobs to sustain caregivers’ labour force attachment (Halldén, Gallie, & Zhou, 2012). Table 1 (row 1) summarizes these distinctions.

Welfare states are also distinguished by the extent to which their labour market policies aim to remove or retain older workers (Blossfeld, Buchholz, & Hofäcker, 2006; Blossfeld, Buchholz, & Kurz, 2011). (Table 1, row 2.) Again, the broad strokes of these approaches have remained quite stable during the years that affect our cohort (Blossfeld et al., 2006). Corporatist and southern regimes favour *employment exit*. High levels of wage replacement through public pensions and early retirement schemes—including, in Italy, relatively lenient access to illness/disability pensions for older workers (Beckstette, Lucchini, & Schizzerotto, 2006)—make early withdrawal attractive. Because these regimes tend not to integrate women into the labour force earlier in the life course, the implications of exit strategies for women’s later-life labour market involvement, or for gender differences in that regard, are unclear.

Social democratic and liberal regimes are oriented toward *employment maintenance*. This takes two forms. The *public-induced* strategies of social democratic regimes combine active labour
market policies (e.g., retraining and job protection) with generous publicly-funded pensions, to foster labour force attachment until—but not necessarily beyond—state pension age. Nevertheless, relatively easy access to disability benefits has provided some older Swedish workers with an early exit route (Lindquist 2006). The market-induced approaches of liberal regimes blend minimal labour market regulation with meagre public pensions that encourage some older adults to remain employed past state pension age and/or if their work histories are not strong. Moreover, for Americans in this cohort health care coverage has been tied to employment—or the receipt of social assistance/disability benefits. This may motivate many older workers to remain in the labour force until they qualify for Medicare at age 65. Both types of maintenance strategies also foster part-time work in later life. Public-induced approaches can include the option to ‘phase out’ while receiving a generous partial pension; in market-induced regimes, part-time work may reflect the combined influence of minimally regulated markets and financial need and/or the greater flexibility afforded those with more personal resources.

Table 1: Institutional contexts for breadwinning/caregiving, and later-life labour market involvement

<table>
<thead>
<tr>
<th></th>
<th>Germany</th>
<th>Italy</th>
<th>Sweden</th>
<th>US</th>
</tr>
</thead>
<tbody>
<tr>
<td>Breadwinning &amp; caregiving earlier in the life coursea</td>
<td>• Families are primary providers of social welfare; state supports are strong</td>
<td>• Families are primary providers of social welfare; state supports are weak</td>
<td>• State plays a role in social welfare provision &amp; employment supports</td>
<td>• State’s role in both social welfare &amp; employment supports is minimized</td>
</tr>
<tr>
<td></td>
<td>• Breadwinner-caregiver model of family life</td>
<td>• Breadwinner-caregiver model of family life</td>
<td>• Publicly-oriented dual-earner model of family lifeb</td>
<td>• Market-oriented dual-earner model of family life</td>
</tr>
<tr>
<td></td>
<td>• Support for full male employment; fairly generous benefits for periods of non-employment</td>
<td>• Fairly ungenerous benefits for periods of (male) non-employment</td>
<td>• Active commitment to full adult employment</td>
<td>• Passive promotion of full adult employment</td>
</tr>
<tr>
<td></td>
<td>• Little/no direct support for mothers’ (full-time) employment</td>
<td>• Little/no support for mothers’ employment</td>
<td>• Fairly generous benefits for periods of non-employment</td>
<td>• Ungenerous social safety net</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td>• Publicly-funded child care; better-quality part-time jobs</td>
<td>• Little/no public support for family responsibilities</td>
</tr>
</tbody>
</table>
Later-life labour market involvement

<table>
<thead>
<tr>
<th>Employment exit: high levels of wage replacement through public pensions and attractive early retirement schemes</th>
<th>Employment exit: high levels of wage replacement through public pensions and attractive early retirement schemes</th>
<th>Employment maintenance (public-induced): generous public pensions; only moderate incentives for early retirement</th>
<th>Employment maintenance (market-induced): low public pensions; few/moderate incentives for early retirement</th>
</tr>
</thead>
</table>
| Official retirement age: 65
d | Official retirement age: 60 for men, 55 for women
d | Official retirement age: 65
d | Official retirement age: 65
d |
| Earliest retirement age: 63 for men, 60 for women
d | Earliest retirement age: 57 (for men)
d | Earliest retirement age: 61
d | Earliest retirement age: 62
d |

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a Adapted from Korpi (2000) and Trifiletti (1999: 54, 56).  
b In a more recent article, Korpi, Ferrarini, and Englund (2013) label Sweden’s model of family life “earner-carer,” due to the introduction of policies designed to encourage male caregiving. However, these policies would not have applied during the years when our respondents were in their prime child-rearing years.  
c Adapted from Buccholz, Hofacker, and Blossfeld (2006).  
d Data sources: Duval (2003: 35) and OECD (2005, 2006, 2009). There will be variation in some of these ages in some countries (e.g., Italy) based on individual work histories. State pension ages listed are those that apply to our birth cohort.

Although broad approaches to older workers differ across welfare regimes, state pension ages for our cohort are generally 65, with uptake possible in the early 60s (Table 1, row 2). The exception is Italy, where both normal and early pension ages are lower—in the former case, especially for women.

The framework of Blossfeld and colleagues (2006, 2011) suggests that later-life labour market trajectories will vary cross-nationally despite similar state pension ages—though the extent and nature of gendering is unclear. Esping-Andersen-based typologies imply that institutional contexts may gender paid and family work earlier in the life course; but how this plays out at older ages is less apparent.

Existing research

Cross-sectional and time-series data show variations in older adults’ labour force participation in line with welfare regime theories. Rates are lowest in Italy, followed by Germany, then the US, and highest in Sweden (OECD, 2011). In all four countries, participation is higher for older men than for their female counterparts; but the gap is largest in Italy, followed closely by Germany, intermediate in the US, and narrowest in Sweden (Kahn, 2010: 30). Among those in the labour force, women exit earlier than men in all four nations, although the gap is somewhat
smaller in the US and Sweden than in Germany and Italy (OECD, 2011). Later working (ages 65–69) is least common in Germany and Italy, more frequent in Sweden, and relatively widespread in the US (OECD, 2011). Part-time work plays a significant role in the lives of older women (but not men) in Germany and Sweden (Morris & Mallier, 2003), is rare for older Italians of both sexes (Morris & Mallier, 2003), and is relatively common among older Americans—especially, but not exclusively, women (US Bureau of Labor Statistics, 2000).

These aggregate-level data cannot, however, reveal long-term biographical patterns or uncover their correlates and antecedents. Research using panel data provides clues, but tends to examine single transitions, such as retirement timing (e.g., Radl, 2013; Warner & Hofmeister, 2006) or re-entry (e.g., Kail & Warner, 2013; Pleau, 2010). This emphasis is problematic from a gender perspective, since it excludes anyone without a reasonably steady work history—that is, many women. By contrast, an approach based on extended labour market sequences—regardless of status at a given age—can include all older adults in all countries, and fully capture cross-national differences in the gendering of later-life labour market involvement.

Few studies consider the extended sequences of older adults; and we know of only one that models employment per se for any of the countries in our analysis. Han and Moen (1999b) assess work histories from age 30 to retirement, and identify five distinct biographies: delayed-entry, orderly, fast-track, steady part-time, and intermittent. Their study is, however, based on a small non-representative American sample, making their results difficult to generalize. In two loosely related studies, Fasang (2010, 2012) uses nationally representative data to analyze income source sequences of older Germans and Britons. She concludes that, in both countries, men tend to follow institutionalized pathways to retirement, while women predominate in part-time work and non-employment trajectories. Although Fasang speculates that gendered family roles shape women’s pathways, she does not test this hypothesis or investigate whether the influence of family experiences differs in the two countries.

Many studies explore the effects of family circumstances on older adults’ employment or exit probabilities—typically focusing on current marital status. For Americans, being married
decreases older women’s attachment to the labour force (e.g., Moen & Flood, 2013), but increases that of older men (e.g., Williamson & McNamara, 2003). Being in a couple has no impact on the odds of employment for German men, but reduces them for German women; makes no difference for Swedish women, but increases employment for Swedish men; and reduces employment for Italian men and, particularly, Italian women (Ogg & Renaut, 2007). However, the life course perspective’s emphasis on long-term circumstances suggests that the critical factor may be marital (in)stability, rather than current marital status. Certainly, the loss of a partner ends or compromises the pooling of resources (Wilmoth & Koso, 2002) and the gendering of breadwinning and caregiving, and might reasonably be expected to promote employment, especially for women.

There is some evidence on the influence of parenthood. Being a mother delays labour market withdrawal for German and American women, as does later age at the time of the first birth (Hank, 2004; Moen & Flood, 2013; Pienta, 1999). Conclusions are mixed regarding the links between labour force participation and having dependent children at older ages. The association may be positive for men and negative for women (Denting & Clarkberg, 2002; Pienta & Hayward, 2002); negative for women only (Moen & Flood, 2013; Szinovacz, DeViney, & Davey, 2001); or nonexistent (Denaeghel, Mortelmans, & Borghgraef, 2011). Importantly, though, results are for single countries (or do not test country interactions), and thus cannot reveal cross-national distinctions in effects. In one cross-national study, Hank and Korbmacher (2013) find that the positive association between being a mother and leaving employment in later life is specific to corporatist regimes for our cohort—although their sample was restricted to those with at least some paid work from age 50.

Regardless, the notion of biographical continuity suggests that earlier work histories—themselves linked to family experiences—may be decisive. Among employed married Americans aged 50-75, men (but not women) with longer work histories are more likely to leave the labour market (Szinovacz & DeViney, 2000). Older American women (married or not) are more likely to be in the labour force if they had fewer employment interruptions earlier in the life course, at least for the cohort born in the 1920s (Pienta, Burr, & Mutchler, 1994).
These studies pertain to the US only. Contrasting institutional contexts mean that prior employment has varied cross-nationally, especially among women—implying that its impact on the gendering of later-life labour market involvement may also vary. In Germany, long maternity leaves, a lack of childcare facilities, and a tax system favouring breadwinner-caregiver families have encouraged women to withdraw or work part-time after marrying or having children (Drobnič, Blossfeld, & Rohwer, 1999; Hofäcker, Stoilova, & Riebling, 2013; Sundstrom, 2003). For Italians, the combination of relatively low education, a lack of part-time work, an underdeveloped service sector, limited child care, and strong norms around the sexual division of labour often persuaded women with family responsibilities to leave the labour force (Rydell, 2002; Sundstrom, 2003). In Sweden, where daycare spending has been high, work-family reconciliation policies relatively comprehensive, equal pay legislation long-standing, and the service sector large (Rovny, 2011: 341), women tended to remain employed during the prime childrearing years (Sundstrom, 2003). How this plays out at older ages is uncertain: Hank and Korbmacher (2013) conclude, based on data from thirteen European countries, that length of prior employment is positively associated with the odds of leaving the labour force, but they do not consider gender differences, either within or between nations.

In sum, a number of gaps are evident in existing knowledge of older adults’ labour market biographies. The emphasis on transitions underrepresents or excludes women; and the few studies that consider long-term trajectories do not focus on employment per se, are not generalizable, lack a comparative dimension, and/or do not take gender into account. Beyond this, comparative evidence on the roles of family and paid work experiences earlier in the life course is sparse, mixed, and not always available for both sexes. In what follows, we begin to flesh out these aspects of older adults’ labour market biographies by testing four hypotheses:

Hypothesis 1: Distributions for later-life labour market biographies will be gendered—with men concentrated in more ‘attached’ (mostly full-time; shorter stretches of non-employment), and women in less ‘attached’ (stretches of part-time; longer periods of non-employment), trajectories.
Hypothesis 2: Male-female distinctions in later-life labour market biographies will be stronger in Germany and Italy than in Sweden and the US. This may be especially true for Italy, given the gender gap in pension eligibility.

Hypothesis 3a: Gender differences in older adults’ employment trajectories will be partly attributable to gender-specific influences of family experiences over the life course. Hypothesis 3b: The impact of family experiences will be stronger in Germany and Italy than in Sweden and the US.

Hypothesis 4a: The gendering of older adults’ labour market trajectories will be at least partly explained by gendered earlier attachment to the labour force. Hypothesis 4b: The explanatory power of work histories will be stronger in Germany and Italy than in Sweden and the US.

METHODS

Data

Our data come from two nationally representative panel studies: the Survey of Health, Ageing and Retirement in Europe (SHARE); and the Health and Retirement Study (HRS) (Börsch-Supan & Jürges, 2005; Heeringa & Connor, 1995). Both surveys sample individuals aged 50+, and include detailed employment/activity histories, as well as key socio-demographic variables. SHARE covers more than 85,000 individuals from 21 countries, over a maximum, to date, of four waves: 2004/5, 2006/7, 2008/9, and 2011/12. We restrict analyses to those born in the 1930s or early 1940s (controlling for historical period and matching the HRS sample) and living in one of three countries of interest—(former West) Germany, Italy, and Sweden. HRS has interviewed more than 26,000 Americans aged 50+, biennially since 1992 and most recently in 2012. We use the original cohort, and only those who were age-eligible according to HRS criteria (born 1931-1941). From both studies, we select those with information on at least ten of the 20 possible labour market state variables over ages 50-69 (N = 475 for Germany, 935 for Italy, 654 for Sweden, and 8,832 for the US). And in each case, we utilize current and retrospective information on employment/activity start and end dates for all years in which
respondents were aged 50-69 years, along with retrospective information on work, marital, and fertility histories up to age 50, and data on gender and other relevant socio-demographics.

**Measures**

*Later-life labour market biographies.* Our study takes a novel approach to modelling later-life labour market involvement—maintaining the integrity of individuals’ *long-term* sequences and classifying them using dynamic Hamming distances (Lesnard, 2010) to make analysis feasible.

The *Analysis* section provides details on the construction of these employment biographies. The resulting variable comprises seven trajectory ‘types’ summarizing patterns over ages 50-69: 1 = Full-time throughout; 2 = Full-time, exit around 65 (reference); 3 = Full-time, exit around 60; 4 = Full-time, exit around 55; 5 = Non-employed throughout; 6 = Part-time dominant; and 7 = Other (Table 2). Note that, because we match each individual to their *closest* model biography, groups are not internally homogeneous; hence, the table and variable labels describe categories as comprising those who were *mostly* employed full-time, part-time or not at all, and/or who exited *around* age 65, age 60, etc.

*Gender* distinguishes women (yes = 1; no = 0) from men.

*Mid-life and prior family experiences* comprise two parental measures—*dependent children at age 50* (yes = 1; no = 0) and *age became a parent* (in years)—and two marital measures—*married at age 50* (yes = 1; no = 0) and *marital ending by age 49* (yes = 1; no = 0). In each case, one measure represents mid-life (age 50) circumstances and one captures prior history. Both marital measures are constructed from the starts and ends of current and prior marriages. Dependent children are those under age 19. Age at the birth of the first child is conditionally-coded (Ross & Mirowsky, 1992: 223-224) to permit the retention of non-parents. As such, the measure is centered on its country-specific sample mean and entered into models in conjunction with a dummy for *parent* (yes = 1; no = 0). Thus, individuals without children form the reference group, and those with children differ from them by an average amount (the coefficient for *parent*) plus an amount based on the age at which they entered parenthood (the coefficient for *age became a parent*).
Table 2. Description of the seven later-life labour market trajectories

<table>
<thead>
<tr>
<th>Later-life labour market trajectory</th>
<th>Employment pattern age 50-69</th>
</tr>
</thead>
<tbody>
<tr>
<td>1. Full-time throughout</td>
<td>Mostly employed full-time ages 50-69</td>
</tr>
<tr>
<td>2. Full-time, exit 65</td>
<td>Mostly employed full-time to approximately age 64, exit approximately age 65</td>
</tr>
<tr>
<td>3. Full-time, exit 60</td>
<td>Mostly employed full-time to approximately age 59, exit approximately age 60</td>
</tr>
<tr>
<td>4. Full-time, exit 55</td>
<td>Mostly employed full-time to approximately age 54, exit approximately age 55</td>
</tr>
<tr>
<td>5. Non-employed throughout</td>
<td>Mostly non-employed ages 50-69</td>
</tr>
<tr>
<td>6. Part-time dominant</td>
<td>Mostly employed part-time with exits at various ages, or ‘downshifting’ from full- to part-time</td>
</tr>
<tr>
<td>7. Other (residual)</td>
<td>Erratic or unclassified pattern</td>
</tr>
</tbody>
</table>

Prior employment uses starts and ends for employment periods to code total years employed over ages 15-49. Distributions are equivalized by cutting at the values nearest each within-country quintile threshold, to produce a five-category ordinal variable.

Controls consist of three variables likely to be associated with gender, family experiences, and labour market involvement: education (low = 1, medium = 2, high (reference) = 3);¹ self-rated health at age 50 (1 = excellent, 2 = very good; 3 = good; 4 = fair; 5 = poor), treated as ordinal in the models;² and, for the US, ‘minority’ status (black/Hispanic = 1; non-Hispanic white = 0).

Table 3 gives weighted means and percentages, by gender and country, for family experiences and prior employment. Of note are the comparatively large proportion of Italians (especially men) with dependent children and/or no marital ending at/by midlife, and the large share of Americans with a marital ending by midlife. These observations are broadly consistent with OECD data for the decades when our samples were under 50, which show low/non-existent divorce rates in Italy, comparatively high rates in the US, and relatively late mean ages at childbirth among Italian women (presumably even later for Italian men, given the gender age gap at first birth) (OECD, 2015a). In addition, while women in all four nations worked less than men in the years before age 50, the size of the gap ranged from less than 6 years for Sweden, through nearly 10 for the US and a little more than 11 for Germany, to 17 for Italy. The relative
magnitude of these gaps is in line with population-level data on men’s and women’s prime-age labour force participation during the years relevant to our cohort (OECD, 2015b).

Table 3. Weighted means and percentages for family and employment history measures, by gender and country

<table>
<thead>
<tr>
<th></th>
<th>Germany</th>
<th>Italy</th>
<th>Sweden</th>
<th>United States</th>
</tr>
</thead>
<tbody>
<tr>
<td>Married at age 50 (%)</td>
<td>87.1</td>
<td>88.4</td>
<td>81.0</td>
<td>78.2</td>
</tr>
<tr>
<td>Women</td>
<td>87.6</td>
<td>87.3</td>
<td>78.5</td>
<td>75.5</td>
</tr>
<tr>
<td>Men</td>
<td>86.6</td>
<td>89.8</td>
<td>83.9</td>
<td>81.1</td>
</tr>
<tr>
<td>Dependent child(ren) at age 50 (%)</td>
<td>40.2</td>
<td>53.1</td>
<td>43.2</td>
<td>39.0</td>
</tr>
<tr>
<td>Women</td>
<td>32.8</td>
<td>45.4</td>
<td>33.9</td>
<td>33.2</td>
</tr>
<tr>
<td>Men</td>
<td>48.8</td>
<td>62.4</td>
<td>53.5</td>
<td>45.2</td>
</tr>
<tr>
<td>Ever had a child (%)</td>
<td>91.4</td>
<td>89.0</td>
<td>87.7</td>
<td>82.1</td>
</tr>
<tr>
<td>Women</td>
<td>93.6</td>
<td>90.0</td>
<td>87.9</td>
<td>85.9</td>
</tr>
<tr>
<td>Men</td>
<td>88.9</td>
<td>87.7</td>
<td>87.5</td>
<td>77.9</td>
</tr>
<tr>
<td>Age in years at first child (if parent, (\bar{x}))</td>
<td>26.4</td>
<td>27.2</td>
<td>25.9</td>
<td>24.5</td>
</tr>
<tr>
<td>Women</td>
<td>24.7</td>
<td>25.6</td>
<td>24.7</td>
<td>22.8</td>
</tr>
<tr>
<td>Men</td>
<td>28.6</td>
<td>29.1</td>
<td>27.2</td>
<td>26.4</td>
</tr>
<tr>
<td>Ever had a marriage end by age 49 (%)</td>
<td>12.8</td>
<td>7.0</td>
<td>20.8</td>
<td>34.5</td>
</tr>
<tr>
<td>Women</td>
<td>13.8</td>
<td>9.0</td>
<td>21.7</td>
<td>36.5</td>
</tr>
<tr>
<td>Men</td>
<td>11.6</td>
<td>4.5</td>
<td>19.8</td>
<td>32.4</td>
</tr>
<tr>
<td>Years employed ages 15-49 ((\bar{x}))</td>
<td>23.6</td>
<td>21.6</td>
<td>27.4</td>
<td>24.0</td>
</tr>
<tr>
<td>Women</td>
<td>18.2</td>
<td>13.9</td>
<td>24.8</td>
<td>19.3</td>
</tr>
<tr>
<td>Men</td>
<td>29.9</td>
<td>30.8</td>
<td>30.3</td>
<td>29.0</td>
</tr>
</tbody>
</table>

*Weighted using the population/design/attrition weights supplied with the surveys.

Analysis

Creating and grouping labour market sequences. To construct the late-life labour market biographies, we first coded, for each of 20 time points (ages 50-69), a set of categorical variables representing whether the respondent was primarily employed full-time, part-time, or not at all. These 20 age-specific variables constitute the later-life labour market sequences. To make analysis feasible while also retaining the integrity of these biographies, we then grouped them using dynamic Hamming distances. Dynamic Hamming distances quantify how alike or distinct individual biographies are, one from another (Abbott and Tsay 2000), by measuring the ‘cost’ of converting one person’s sequence to another’s (MacIndoe and Abbott 2004). Sequences can then be grouped on that basis. We initially calculated distances relative to a set
of 12 model biographies, constructed using existing knowledge of work and retirement among older adults. Because numbers were too small to support further analysis for some groups in the SHARE samples, we subsequently combined the seven biographies involving stretches of part-time work (leaving the remaining five as they were), to create a 6-category measure for all countries. We then generated a seventh group by pulling, from the original 12, those cases deemed a less than ideal match to any reference sequence, based on their distance from the one to which they were closest. Initial model sequences were developed separately by all authors, and later inspected for overlap (found to be substantial) and unique patterns of theoretical interest.

The validity of the 12-group classification was checked by examining between-group heterogeneity and within-group homogeneity—both considerable—using information on individuals’ ‘own-group’ distance measures (available on request). In addition, group-specific sample means for employment states at each age (Appendix 1) demonstrate that the collapsed typology represents identifiable longer-term patterns in later-life labour market involvement. Finally, further dynamic Hamming analyses using model reference sequences derived from detailed inspection of the 13th group (not shown) verified that no additional viable later-life employment biography groups could be extracted. Notwithstanding all of the above precautions, it is wise to keep in mind that the later-life labour market categories are not internally homogeneous, but rather, are clusters of individuals with similar employment sequences.

Statistical analyses. After assessing the gendering of older adults’ labour market trajectories at baseline, we use nested multinomial regression models to investigate the roles of family experiences and prior employment in shaping observed patterns. The models first include controls (education, health at age 50, and, for the US, minority status), then add family circumstances at age 50 (whether married and whether dependent children), followed by family history (whether a parent, age at first child, and whether any marital ending by age 49), and finally include prior employment (years employed ages 15-49). All analyses are stratified by gender and country to allow for expected gender- and country-specific effects of at least the
family variables. All models are run in Stata, with procedures designed to handle multiple imputations. Estimates are weighted using the population/design/attrition weights supplied with the surveys.

As multinomial coefficients are difficult to interpret in their raw form—and because we run separate models for women and men—we evaluate the roles of family experiences and prior employment by converting the logits to adjusted probabilities (with confidence intervals) for each trajectory group, by gender and country.\(^5\) Probabilities are assessed at the country-specific sample means for all covariates, and confidence intervals are calculated using Rubin’s rules (Rubin, 1987) to accommodate the uncertainty associated with multiple imputations. We then quantify the gendering of later-life labour market biographies using ‘gender gradients’ or relative risks (adjusted probability among women / adjusted probability among men) for each trajectory type, by country. Where male and female confidence intervals do not overlap, we compare the magnitudes of these gradients within and between countries, and across models.

**Imputation.** Our method requires complete data on the sequence variables.\(^6\) When data were missing, we imputed values to retain as many cases as possible and minimize bias.\(^7\) Imputation used a two-fold fully conditional multiple imputation specification (Van Buuren, Brand, Groothuis-Oudshoorn, & Rubin, 2006) implemented in Stata -ice-. Two-fold ice was designed specifically for panel data, to impute using all cross-sectional data plus prior and subsequent values for variables with missingness. In brief, the method computes, for each missing value, its posterior distribution conditional on other variables in an imputation model. A value is then sampled from this distribution under the assumption that missingness is random given the values of the other variables in the model. The method uses a Markov chain Monte Carlo algorithm. After double iteration of the algorithm, a complete dataset is created, consisting of a mix of imputed and known values. Enough complete datasets are generated—20, in our case—to ensure the accuracy of substantive model estimates (Graham, Olchowski, & Gilreath, 2007). All analyses are based on the simultaneous investigation of these 20 data sets, averaging over them and deriving standard errors according to Rubin’s rules (Rubin 1987). Following optimal
matching, missing covariate values were imputed using chained equations (White, Royston, & Wood, 2011) in Stata -mi impute-.

RESULTS

Baseline distributions and gradients

Table 4. Weighted male and female percentages (CIs) in each later-life labour market trajectory type, along with gender gradients, by country

<table>
<thead>
<tr>
<th>Later-life labour market trajectory</th>
<th>Germany</th>
<th>Italy</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Men</td>
<td>Women</td>
</tr>
<tr>
<td>Total</td>
<td>100.0</td>
<td>100.0</td>
</tr>
<tr>
<td></td>
<td>(sample N=252)</td>
<td>(sample N=223)</td>
</tr>
<tr>
<td>1 Full-Time Throughout</td>
<td>5.2 (2.3, 8.2)</td>
<td>1.5 (-0.6, 3.7)</td>
</tr>
<tr>
<td>2 Full-Time, Exit 65</td>
<td>39.6 (33.1, 46.0)</td>
<td>7.5 (3.1, 11.9)</td>
</tr>
<tr>
<td>3 Full-Time, Exit 60</td>
<td>34.5 (28.2, 40.8)</td>
<td>12.5 (7.2, 17.8)</td>
</tr>
<tr>
<td>4 Full-Time, Exit 55</td>
<td>12.3 (7.7, 17.0)</td>
<td>4.0 (1.4, 6.6)</td>
</tr>
<tr>
<td>5 Non-Employed Throughout</td>
<td>2.5 (-0.1, 5.0)</td>
<td>47.8 (40.5, 55.0)</td>
</tr>
<tr>
<td>6 Part-time Dominant</td>
<td>2.9 (0.3, 5.4)</td>
<td>21.2 (15.5, 26.8)</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Later-life labour market trajectory</th>
<th>Sweden</th>
<th>United States</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Men</td>
<td>Women</td>
</tr>
<tr>
<td>Total</td>
<td>100.0</td>
<td>100.0</td>
</tr>
<tr>
<td></td>
<td>(sample N=316)</td>
<td>(sample N=338)</td>
</tr>
<tr>
<td>1 Full-Time Throughout</td>
<td>14.2 (10.0, 18.4)</td>
<td>5.2 (2.7, 7.7)</td>
</tr>
<tr>
<td>2 Full-Time, Exit 65</td>
<td>46.2 (40.3, 50.2)</td>
<td>29.5 (24.3, 34.8)</td>
</tr>
<tr>
<td>3 Full-Time, Exit 60</td>
<td>24.3 (19.2, 29.4)</td>
<td>18.3 (13.7, 22.9)</td>
</tr>
<tr>
<td>4 Full-Time, Exit 55</td>
<td>5.1 (2.3, 7.8)</td>
<td>5.8 (2.9, 8.8)</td>
</tr>
<tr>
<td>5 Non-Employed Throughout</td>
<td>1.2 (-0.3, 2.6)</td>
<td>6.6 (3.9, 9.4)</td>
</tr>
<tr>
<td>6 Part-time Dominant</td>
<td>5.7 (2.9, 8.5)</td>
<td>26.4 (21.3, 31.5)</td>
</tr>
</tbody>
</table>

* Gradients are the ratio of the female proportion to the male proportion. As such, they measure the extent to which one gender predominates in a given trajectory group in a given country. Light grey shading identifies larger gradients (> 2.0 for female-dominated or < 0.5 for male-dominated trajectories), for which male and female confidence intervals do not overlap; dark grey shading identifies very large gradients (> 5.0 or < 0.2).

† Confidence intervals do not overlap for male and female estimates, but gradients are smaller than two times. Confidence intervals tend not to overlap for the American sample, simply because of the much larger N; hence, we focus on the magnitudes of (significant) gaps.

Table 4 displays weighted baseline percentages for each trajectory group, by gender and country, along with the resulting gender gradients (female percentage / male percentage). Light grey shading identifies larger gradients (> 2.0 for female-dominated trajectories or < 0.5 for male-dominated) — and dark grey pinpoints very large gradients (> 5.0 or < 0.2) — in which gender-specific confidence intervals do not overlap. Note that we focus on the magnitudes of (significant) gaps as confidence intervals tend not to overlap for the American sample even
when gradients are fairly small, simply because of the much larger N. We highlight two features: the amount and type of gendering; and how this varies between countries. We leave aside discussion of the residual category, as its meaning is not consistent across genders or countries—although all models do include this group.

Most trajectory types are gendered, with gradients generally supporting Hypothesis 1: Men predominate in the ‘attached’ biographies (1-3) and women in the less ‘attached’ (5 and 6). Interestingly, though, gradients are typically largest for female-dominated trajectories. For example, continual non-employment in Germany is almost entirely female; and the part-time dominant path in Germany, as well as continual non-employment in Italy and Sweden, all have gradients larger than five. Only one male-dominated biography has a gradient of more than five times (i.e., less than 0.20): full-time to exit around 65 in Germany.

The shading in Table 4 provides clear support for Hypothesis 2: The gendering of later-life employment biographies is strongest in Germany and Italy, where only the very small groups display non-significant male-female differences. Unexpectedly, however, gradients are often larger for Germany than for Italy, despite the gender gap in state pension ages in the latter country. Proportions underlying the strongest gradients show that a far smaller share of German than Italian men followed the female-dominated continual non-employment path, and a far greater proportion of German men are in the male-dominated full-time to exit around 65 group.

Beyond this, the results add nuance to the notion that male-female differences will be weakest in Sweden and the US (Hypothesis 2). Instead, gendering is intermediate for Sweden, where one male-dominated and two female-dominated trajectories have steep gradients, and weakest for the US, where only continual non-employment is clearly gendered. Of note is the part-time dominant biography among Swedes, which is both strongly feminized and large (26% of women). This implies that the Swedish dual-earner strategy of incorporating mothers into the labour force partly via better-quality short-hour jobs carries through to older ages. Subsequent investigation (not shown) reveals that Swedish women in the part-time dominant group had
spent, on average, nearly a third of their younger years working ‘short’ hours, versus about a tenth of those years for all other Swedish women combined. Despite this evidence of gendering, we note that the (always male-dominated) full-time to exit around 65 path is actually the most common one taken by Swedish women; and this cannot be said of any other country.

**Family and prior work**

Table 5. Female/male gender gradients\(^a\) for later-life labour market trajectories, by country

<table>
<thead>
<tr>
<th></th>
<th>Model 1 (controls only)</th>
<th>Model 2 (+ family age 50)</th>
<th>Model 3 (+ family history)</th>
<th>Model 4 (+ work history)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Germany</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1. Full-time throughout(^b)</td>
<td>0.31</td>
<td>0.33</td>
<td>0.07</td>
<td>0.07</td>
</tr>
<tr>
<td>2. Full-time, exit 65/65+</td>
<td>0.60</td>
<td>0.59</td>
<td>0.59</td>
<td>0.69</td>
</tr>
<tr>
<td>3. Full-time, exit 60</td>
<td>0.73</td>
<td>0.67</td>
<td>0.74</td>
<td>0.78</td>
</tr>
<tr>
<td>4. Full-time, exit 55</td>
<td>0.95</td>
<td>0.98</td>
<td>0.98</td>
<td>1.07</td>
</tr>
<tr>
<td>5. Non-employed throughout</td>
<td>†</td>
<td>†</td>
<td>†</td>
<td>†</td>
</tr>
<tr>
<td>6. Part-time dominant</td>
<td>5.10</td>
<td>8.13</td>
<td>8.56</td>
<td>7.70</td>
</tr>
<tr>
<td>Italy</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1. Full-time throughout</td>
<td>0.19</td>
<td>0.19</td>
<td>0.21</td>
<td>0.35</td>
</tr>
<tr>
<td>2. Full-time, exit 65</td>
<td>0.31</td>
<td>0.29</td>
<td>0.30</td>
<td>0.47</td>
</tr>
<tr>
<td>3. Full-time, exit 60</td>
<td>0.41</td>
<td>0.38</td>
<td>0.15</td>
<td>0.31</td>
</tr>
<tr>
<td>5. Non-employed throughout</td>
<td>1.37</td>
<td>1.55</td>
<td>1.69</td>
<td>4.95</td>
</tr>
<tr>
<td>Sweden</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>1. Full-time throughout</td>
<td>0.49</td>
<td>0.49</td>
<td>0.50</td>
<td>0.57</td>
</tr>
<tr>
<td>2. Full-time, exit 65</td>
<td>0.64</td>
<td>0.63</td>
<td>0.63</td>
<td>0.68</td>
</tr>
<tr>
<td>3. Full-time, exit 60</td>
<td>0.72</td>
<td>0.71</td>
<td>0.73</td>
<td>0.80</td>
</tr>
<tr>
<td>4. Full-time, exit 55</td>
<td>1.26</td>
<td>1.23</td>
<td>1.24</td>
<td>1.23</td>
</tr>
<tr>
<td>5. Non-employed throughout</td>
<td>2.94</td>
<td>3.03</td>
<td>3.15</td>
<td>1.91</td>
</tr>
<tr>
<td>6. Part-time dominant</td>
<td>1.27</td>
<td>1.29</td>
<td>1.27</td>
<td>1.40</td>
</tr>
</tbody>
</table>
In the calculation of the underlying adjusted probabilities, all covariates are set to their within-country sample means. Gradients are highlighted where confidence intervals for male and female adjusted proportions do not overlap. Light grey = male-dominated, dark grey = female-dominated.

Trajectories 1 and 2 were combined for Germany only.

The very large gradients (relative to Table 4) occur because low education dramatically decreases men’s, but not women’s, odds of having this trajectory. This is largely a function of small cells: There are few German men in the part-time dominant group and, as a result, the cell for low education is actually empty.

†Virtually all cases are female.

We next investigate the contributions of family experiences and prior employment to the above patterns (Table 5). Again, we quantify male-female distinctions using gender gradients. But here the underlying probabilities are calculated from regression coefficients that adjust for the influence of family and earlier work. (See Appendix 2 for the adjusted probabilities and confidence intervals, and Appendix 3 for average marginal covariate effects.) In this table, light-grey shading indicates male-dominated, and dark-grey, female-dominated trajectories. Again, we do not discuss the residual category—although all models do include this group. In addition, for German men we combine full-time to exit around 65 and full-time throughout, to correspond to women’s trajectories in that country, where collapsing was necessary because so few were in ongoing full-time employment. For similar reasons, we fold the small number of German and Swedish men in continual non-employment into the residual category. Model 1, which adjusts for controls only, is included as the comparator for Model 2. Models 2 and 3 test Hypothesis 3a—that gender differences in later-life labour market biographies are rooted partly in the differential impact of family experiences over the life course. Model 4 tests the hypothesis that the gendering of later-life labour market involvement is at least partly due to differential attachment to the labour force at younger ages (Hypothesis 4a). We highlight two features of Table 5: how gradients shift with the addition of explanatory variables; and cross-national differences in those shifts.

Relative to the model with controls, some Model 2 gender gaps increase—most obviously for the part-time dominant path among Germans and Swedes and ongoing full-time employment among Italians. (For gradients < 1, a larger gap is represented by a decrease from one model to the next, meaning a greater preponderance of men.) In addition, women’s greater probability of continual non-employment increases for Germany and Sweden, although no gradient is
shown because no male comparison group exists. Model 3 (Table 5) displays the additional influence of family experiences earlier in the life course. Gender differentials increase for the part-time dominant path among Germans, ongoing full-time employment among Swedes, and the full-time to exit around 55 biography among Italians.

Models 2 and 3 provide support for the notion that the gendering of later-life labour market involvement arises, in part, from the gender-specific influences of family experiences over the life course (Hypothesis 3a). They only partially back Hypothesis 3b—that effects will be stronger for Germany and Italy than for Sweden and the US. Instead, family experiences appear more important in all three European countries (two breadwinner-caregiver and one dual-earner) than in the US (dual-earner), where they have little or no impact.

Model 4 (Table 5) tests Hypothesis 4a—that the gendering of older adults’ labour market trajectories arises at least partly from gendered earlier attachment to the labour force. Indeed, prior employment partially explains many of the family-adjusted gender gaps (trajectories 2 and 5 in Germany; trajectory 6 and, to some extent, trajectory 2 in Sweden; trajectories 3-5 in Italy; and trajectory 5 in the US), or reduces them to non-significance (trajectory 2 in Italy; trajectory 3 in Germany and the US). The results are also consistent with Hypothesis 4b: Male-female differences in earlier employment account for a greater share of the gendering of later-life labour market involvement in Germany and Italy, where most trajectory types are affected, than in Sweden and the US, where effects primarily involve one male-dominated and one female-dominated biography.

DISCUSSION
This study adopted a comparative gendered life course approach, to extend knowledge of older adults’ labour market involvement. While the majority of previous work has focused on retirement and thereby underrepresented women, our analysis drew on extended biographical sequences and fully incorporated both genders in four nations with diverse approaches to family and paid work over the life course. This strategy—to our knowledge, unique among
studies of older adults’ labour market involvement—generated results that both confirmed and expanded existing knowledge.

Consistent with expectations, in every nation women prevailed in groups representing a weak(er) attachment to the labour market in later life and men in those signifying a strong(er) attachment. Additionally, this pattern was especially pronounced for Germany and Italy, where policies have generally supported men’s, but not women’s, paid work earlier in the life course. These findings are in line with cross-sectional estimates of older adults’ labour force participation (OECD, 2011). However, they also offer novel insights into how gender and institutional contexts play out with respect to long-term patterns of labour market involvement. For example, across all nations, gradients tended to be stronger for female- than male-dominated biographies, implying that women were more likely to follow ‘male’ paths than the reverse. Future research should investigate the institutional and cultural factors behind this finding (e.g., the extent of support for men’s family work), along with its relevance to upcoming cohorts where family roles may be less gendered.

We also found support for the idea that male-female differences in later-life labour market involvement arise, in part, from the gendered influence of family experiences over the life course. More importantly, we offered new evidence that family ties matter more in the corporatist and southern welfare regimes, where a breadwinner-caregiver division of labour structures the adult life course, than they do in the liberal nation, where a dual-earner model is (passively) promoted. This highlights the need for approaches to older adults’ labour market involvement that consider not only state benefits that ‘pull,’ and workplace barriers that ‘push,’ older adults out of the labour force, but also policies related to family circumstances throughout the life course. Future studies might extend these findings by drawing on more detailed marital and parental histories and identifying the specific social, cultural, and policy mechanisms by which they operate in various institutional contexts.

The real significance of family policies undoubtedly lies in their impact on the gendering of earlier attachment to the labour force, which may carry through to older ages. Indeed, prior
employment accounted for a significant share of male-female distinctions in later-life labour market involvement. Again, this was most apparent for Germany and Italy, where, historically, women have been encouraged to leave the labour force once married and/or raising children. Conversely, work histories explained less of the gendering at older ages in Sweden, where the state has, for several decades, actively supported women’s employment throughout the life course, and in the US, where the state has passively promoted dual-earning. These findings reiterate the need for approaches to older adults’ employment that consider the entire life course.

Three key results were not anticipated. First, in some ways the divide was between the three European countries and the US, rather than between those adhering to a breadwinner-caregiver and a dual-earner family model. Family experiences made almost no difference to the gendering of older Americans’ labour market biographies. But the part-time dominant biography among Swedes was both highly gendered (female) and common; and the effects of family experiences were particularly strong for that group. These findings, along with the disproportionate levels of earlier part-time work found for Swedish women in the part-time dominant group, suggest that that nation’s strategy of incorporating mothers into the labour force partly by means of better-quality short-hour jobs carries through to older ages.

Second, we turned up particularly strong evidence of gendering for Germany—stronger than that found for the other breadwinner-caregiver nation (Italy) and despite the gender gap in the latter nation’s pension ages. The much smaller share of German than Italian men in the female-dominated continual non-employment group may follow from Italy’s relatively lenient access to illness/disability pensions for older workers; and the far greater proportion of German men with the male-dominated full-time to exit around 65 biography may reflect Germany’s pension system, which is closely tied to (male) earnings over the life course (Buchholz 2006).

And finally, the results for work histories were somewhat surprising. We found roughly similar effects for the US and Sweden, as expected; but this was despite the larger gender gap in prior employment in the US, which, all else equal, should have meant that work histories accounted
for more gendering of later-life labour market involvement in that country. That they did not imply that some American women with a weak(er) prior attachment to the labour force ended up behaving more like men with respect to their labour market involvement in older adulthood. This may reflect contextual factors—a heavy reliance on accumulated personal resources to fund retirement—requiring some without steady employment records to increase their work effort as they approach (or pass) state pension age. Pending the availability of more detailed data, future research should further investigate cross-national differences in the role of work histories (including full-time/part-time status and interactions with family histories) in shaping older women’s employment biographies.

Our contributions should be viewed in light of several limitations. One is the necessary sacrifice of some precision in the construction of the later-life labour market groups. We clustered individuals whose trajectories are similar but not identical, and we cannot say whether this ‘muddied’ associations with the key explanatory variables. This constraint is, however, inseparable from one of the study’s strengths: the ability to include women, even those without steady (or any) employment in their later years, and thus to assess gender differences in older adults’ employment trajectories across a range of welfare states. A second consideration is that imputations introduced some uncertainty into the labour market sequences. While this is of concern, we minimized its impact—using a two-fold fully conditional specification, running a large number of imputations, and appropriately adjusting standard errors. Moreover, the vast majority of sequences were complete for all countries except the US, and our sensitivity analysis using a reduced US sample did not alter our major conclusions. A third limitation is the possibility that our measures of family experiences did not capture all relevant aspects, while our prior employment variable could not incorporate, for example, part-time work. These shortcomings reflect a combination of data and methodological constraints—inconsistent information on prior work periods in the HRS, and relatively small sample sizes for the SHARE countries, which restricted the number and type of covariates possible.

Our discussion highlights at least three policy implications. First is the need to go beyond a focus on workplace and economic factors that shape employment decisions in the years
surrounding public pension eligibility. We found evidence that older adults’ labour market trajectories are influenced by policies related not only to paid work but also to caregiving, and by those affecting not only immediate decisions but also choices made earlier in the life course. Second are possible vulnerabilities associated with the drive to extend working life. The feminization of non-employment in Germany and Italy, and of part-time work in Germany and Sweden, may leave women out in the cold if policies fail to consider the gendering of paid work and caregiving earlier in the life course. Third, the American case hints at concerns related to weakly regulated labour markets and scanty public provisions. That women with weak(er) work histories may end up behaving more like men with respect to later-life labour market involvement, raises questions about the quality of at least some older Americans’ work—questions whose urgency will increase if adults are required to work longer. These considerations should be a central focus of ongoing cross-national investigations of older adults’ labour market trajectories.

1 These correspond to ISCED levels 0-2 (less than secondary), level 3 (completion of secondary school), and levels 4-plus (beyond secondary) (UNESCO, 2006). Because very few Italians were educated beyond secondary school, high and medium education are collapsed for that country.
2 Because most respondents were not interviewed at age 50, we used all available health information, along with a range of relevant predictors, to estimate country-specific growth models with random age-50 intercepts. These random intercepts serve as imputed values for self-rated health at age 50 (Biesanz, Deeb-Sossa, Papadakis, Bollen, & Curran, 2004; Ware & Wu, 1981).
3 Distributions for labour market trajectory type and the imputed covariates may vary across the imputation datasets. If, for any of the 20 imputed datasets, there is a zero cell in the cross-classification of employment trajectory group with any of the categorical independent variables, standard errors cannot be estimated.
4 We also ran the US analysis without controlling for minority status. Results (available on request) are essentially identical to those presented here.
5 We stratify by gender because collapsing means we do not always have the same groups for women and men in each country (see Table 5).
6 The alternative is to treat missingness as a category in its own right. However, this strategy introduces a ‘state’ whose meaning is difficult to interpret.
7 Most cases in the analytic sample actually had complete information: 95, 97, 93, and 66%, respectively, for Germany, Italy, Sweden, and the US. We performed a sensitivity analysis—available on request—using an American sample comprising all cases with complete information plus a 10% random sample of cases with sequence missingness. This yielded 91.4% complete sequences, a rate comparable to that of the SHARE samples. The analysis prompted no change in our major conclusions.
8 The gradient for full-time-to-exit around 55 drops to non-significance among Americans, but confidence intervals were already adjacent in Model 1.
ACKNOWLEDGEMENTS

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