Self-Report and Behavioural Measures of Impulsivity as Predictors of ‘Real World’ Impulsive Behaviour and Psychopathology in Male Prisoners

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Running Head: Impulsivity and substance misuse in prisoners
Abstract

Impulsivity is an important factor in adverse outcomes such as substance use, problem gambling and psychopathology. Extensive research has shown these negative outcomes are associated with both self-report and behavioural measures of impulsivity but these two measurement domains are not themselves associated. There has been limited research in prison samples. This is surprising given the high variability in impulsive behaviours that should make them ideal for investigating the convergence of impulsivity measures. Using a cross sectional design we investigated the associations of impulsivity – measured by self-report and two behavioural indices - with substance misuse and psychopathology in a sample of 72 male prisoners. We found higher self-reported impulsivity was associated with crack/cocaine use, problem gambling and a positive screen for personality disorder. Behavioural measures of impulsivity showed fewer associations with problematic behaviours; they were also not independent predictors of impulsive behaviour in multivariate analyses. These data suggest that self-reported impulsivity is a more consistent predictor of problematic behaviours than behavioural measures in a sample of people with significant levels of substance use and psychopathology. This difference could reflect relevance of self-reported measures to emotionally charged decision-making in daily life compared to more neutral behavioural measures.

Key words: addiction; impulsivity; substance use; prison; offender; problem gambling
Introduction

Higher levels of impulsivity have been linked to a range of behaviours that impact on daily functioning (Sharma, Markon & Clark, 2014). For example, there is evidence for higher levels of impulsivity in those with greater substances use problems (Verdejo-Garcia et al., 2006), psychopathology (Chamorro et al. 2012) or offending behaviour (Leverso, Bielby and Hoelter, 2015). Such findings have informed theories of addiction (West and Brown, 2013) and criminality (Lynam and Miller, 2004). However, recent commentary suggests further understanding of the construct of impulsivity is required before additional theoretical progress can be made in making sense of its role in adverse behavioural outcomes (Sharma et al. 2014). One obstacle concerns the measurement of impulsivity, which falls into two broad categories. One domain is self-report measures that are assumed to capture what participants do across time and situations. In contrast, behavioural measures are intended to capture the manifestations of underlying traits assessing what people do in specific situations. A recent comprehensive review of the literature found robust support for associations between problematic daily life behaviours and impulsivity across both the two main domains of measurement (Sharma et al. 2014). However, the associations between self-reported and laboratory behavioural measures of impulsivity were consistently low. They authors concluded that each domain is tapping unique variance in daily life behaviour, if true this would question of the validity of impulsivity as a single construct, and support the notion of ‘varieties of impulsity’ (Evenden, 1999).

The majority of research on the two domains of impulsivity has been conducted within specific clinical groups or general population samples. These studies have shown weak associations between the two domains of impulsivity
measures though associations are somewhat stronger between measures of impulsivity and impulsive behaviour (Sharma et al. 2014). A possible means to increase sensitivity is to identify samples where there is likely to be high variability on measures of impulsivity, psychopathology and types of problematic behaviours.

Prison populations are ideal in this regard as high levels of both substance use (Fazel et al. 2006; Cooper et al. 2016) and mental health difficulties (Fazel & Danesh, 2002) produce greater variability on these dimensions compared to the general population.

Existing research has established consistent associations between self reported impulsivity and substance use in prison samples (Cuomo et al., 2008, Devieux et al., 2002, Ireland and Mooney et al., 2008, Bernstein et al., 2015), with only isolated exceptions (Fishbein and Reuland, 1994). In contrast, there have been very few studies using laboratory measures in samples of prisoners. An exception is discounting, which is the tendency to perceive and attribute reduced value to delayed rewards, even if these are preferable to more immediate gratification (Bickel and Marsch, 2001). There is evidence of higher rates of discounting in prisoners compared to the general population (Arantes, Berg, Lawlor & Grace, 2013; Wilson and Daly, 2006) and evidence that discounting is associated with criminal thinking styles (Varghese, Charlton, Wood & Trower, 2014). However, only one of these studies examined the association between discounting and substance use in prisoners (Arantes et al. 2014) and this found this was not significant. In the same study there was no evidence of convergence between the domains of impulsivity. Instead they were, surprisingly, negative correlated (Arantes et al. 2013). The latter single finding is an indication of the very limited research on the degree of convergence between the two domains of impulsivity measurement in prison samples.
As already noted, mental health problems are elevated amongst prisoner populations but potentially the most important problem for current purposes is the heightened rates of personality disorder (Fazel & Danesh, 2002). There is substantial evidence of comorbidity between substance use and personality disorders (Nace et al, 1991; Bowden-Jones et al, 2004; Compton et al, 2007). It may be that this reflects bidirectional or mutual causation, whereby substance use is a response to extreme emotional states in those with personality disorder, which subsequently exacerbates affective disturbance. Similar bi-directionality also has been found in relation to criminality and substance use (Xue et al. 2009). Another perspective is that impulsivity could be a common factor underlying comorbid personality disorder and substance use (Trull et al, 2000). Indeed, the presence of a comorbid personality disorder in those that abuse substances is associated with markedly high impulsive behavior on various tasks (Petry, 2002; Dom et al, 2006; Rubio et al, 2007). Another revealing study by Dom et al (2006) found that while abnormalities in response inhibition distinguished problem drinkers with personality disorder from those without it, this discrepancy was not found on a delay-discounting task. This highlights the importance of studying impulsivity with multiple measures as this holds potential for disentangling the overlap between substances use and personality disorder in offending populations. Alternatively it may be that impulsivity is an underlying shared mechanism across these distinct adverse outcomes.

The current study aimed to determine both the relationship between the measurement domains of impulsivity and their individual associations with daily-life impulsive behaviour in a prison sample. For the behavioural domain of measurement we included the discounting task previously employed in studies of prisoners. This behavioural measure was complemented with matching familiar figures test (MFFT)
The two tasks are thought tap different properties of impulsivity with discounting defined as ‘choice impulsivity’ and the MFFT measuring ‘reflection impulsivity’ (Sharma et al. 2014). Reflection impulsivity is the tendency for individuals to engage in behaviour without appropriate reflection or deliberation. Greater reflection impulsivity is associated with problematic use of various substances (Morgan 1998; Clark, Robbins, Ersche & Shakian, 2006). To our knowledge, no studies have been conducted on the association between reflection impulsivity and substance use in samples of prisoners. There has also been no research on the association between distinct behavioural measures of impulsivity in this population.

Our first aim was to establish the relationships between self-report and behavioural measures of impulsivity in a prison sample. Our second aim was to determine the associations of the three impulsivity measures with the extent of substance misuse, problem gambling and psychopathology. Evidence of psychopathology was determined using a screening tool for personality disorder (Moran et al. 2003).

Method

Participants

Seventy-two participants were recruited from a Category C adult male prison for prisoners aged 21 and older in London, United Kingdom (UK). Category C prisons are the third highest level of security in the UK justice system. They provide
closed conditions so that prisoners’ movement is restricted so that they must spend much of their time confined in cells. Recruitment took place through a prison mental health service. The service screened prisoners less than 35 years old upon reception into prison for early detection of at risk mental states for psychosis (Jarrett et al., 2012). For the purpose of this study all prisoners screened were asked to participate independent of the outcome of their screening. Exclusion criteria included prisoners not screened by the mental health service (i.e. above 35 years or those refusing screening); those who could not speak English; and those identified as experiencing a current psychotic and/or severe depressive episode and/or those reporting a history of head injury, given potential interference of such difficulties during neuropsychological assessment (Heerey et al., 2007, Lempert and Pizzagalli, 2010, Slaughter et al., 2003). These participants’ eligibility against these criteria was determined by interview.

Procedure

Participants were seen for assessment in accordance with local prison policies governing the times during which prisoners are allowed out of their cells, usually for approximately two to three hours during the morning and for a similar period in the afternoon. The study was approved by both the local Research Ethics Committee and National Offender Management Service (NOMS). After informed consent was obtained they completed the measures in the order presented below.

Measures

**Barratt impulsiveness scale (BIS).** The BIS (Version 11; (Patton et al., 1995)) is a 30-item measure widely used to assess impulsive personality traits, comprising a total score and subscale scores for trait domains of (i) attentional, (ii)
motor and (iii) non-planning impulsiveness. The current analysis used the total BIS score as a measure of trait impulsivity, with scores treated as continuous.

**Monetary choice questionnaire (MCQ).** The MCQ (Kirby and Marakovic, 1996, Kirby et al., 1999) was used to measure delayed reward discounting. The MCQ is a 27-item assessing how quickly individuals tend to discount delayed rewards in favour of immediate rewards; the discounting rate, \( k \). Trials differ both in terms of temporal delay to receipt of larger reward and in size of delayed reward. Guidance from Kirby (Kirby, 2000) was used to infer \( k \) for each reward magnitude; the geometric mean of these was taken as an overall measure of discounting, as used previously (Kirby et al., 1999). Higher \( k \) is considered indicative of elevated discounting. Discount rates were treated as continuous for analysis.

**Matching familiar figures test (MFFT).** The MFFT is a 20-item behavioural measure of reflection impulsivity (Cairns & Cammock, 1978). The MFFT-20 has been used in the studies of reflection impulsivity in adult substance users (e.g. Morgan et al, 1998). The format for administration of the MFFT-20 involves the presentation of a familiar figure, such as a leaf or a house, alongside six similar figures where only one of these six matches the familiar figure exactly. Participants are asked to choose which of the six options matches the presented figure exactly. Individual performance is determined by calculating a participant’s mean latency to first response and their total number of errors, each of which is computed into a standardised Z-score. An index of impulsivity is then created (i-score), by subtracting Z-latency from Z-error (Salkind & Wright, 1977; Messer & Brodzinsky, 1981).

**Substance misuse.** A semi-structured interview was employed that had been previously developed for use in this setting (Cooper et al 2016). This utilised
questions from Cannabis Experience Questionnaire (Barkus et al., 2006) that were extended to obtain detailed information on lifetime use of alcohol, tobacco, cannabis, cocaine (including crack cocaine), amphetamines and opiates (including substitute medication). The psychometric properties of this instrument have been established (Barkus, Stirling et al. 2006). Prison is an environment where current access to substances is restricted so lifetime substance use prior to incarceration was used as the main indicator of the severity of this aspect of impulsive behaviour. When the lifetime use of a substance was reported a further question on the intensity participants’ peak lifetime episode of substance use was rated on five-point scale from “only once or twice a year” to “everyday”. Prisoners were assured that their responses were confidential and would not be disclosed to other services.

**Problem Gambling.** The Problem Gambling Severity Index (PGSI) is a 9-item self-report questionnaire designed as a screening measure of problem gambling severity within the general population (Ferris & Wynne, 2001). Each item on the PGSI asks informants a question relating to some aspect of any gambling behaviour over the previous 12 months. A score of 8 or more is defined as a cut-off for severe problem gambling.

**Personality Disorder Screening.** The Standardised Assessment of Personality Abbreviated Scale (SAPAS) is a brief 8-item structured interview developed for use as a clinical screen for personality disorder (Moran et al, 2003). Existing studies support its validity as a brief screening tool and these include when used with offender populations (Pluck et al, 2012).

**Statistical Analyses.** Correlation analyses were initially used to establish the pattern of bivariate associations between 1) each of the measures of impulsivity and
2) all three measures of impulsivity with each substance use domain. Pearson correlations were used for normally distributed interval level data, Spearman Rank correlations for ordinal data and Point Biserial correlations for dichotomous variables. The latter analyses were required for the lifetime substance misuse measure, problem gambling and personality disorder screen. Multiple logistic regressions were next used to assess whether independent variables emerged as independent predictors of frequent use of a particular substance or personality disorder screening status. The MCQ-K variable was positively skewed so a root transform was performed for its use as a predictor in the multiple regression analyses.

**Results**

**Sample characteristics.** Table 1 reports the age, substance use and summary statistics for the PGSI and SAPAS scales. The proportion of participants screening positive for problem gambling was (14%, N=10) and a positive screen for personality disorder was (51%, N=37). We chose to focus our analyses of substances use on lifetime usage of combined crack/cocaine and opiates. Participants’ very frequently reported lifetime use of the remaining the substances (alcohol or cannabis in > 85% of participants) so no further analyses of associations with impulsivity were conducted using these variables.

== Table 1 about of here ==

**Associations between impulsivity measures (Table 2).** We found that higher scores on the BIS were weakly, but significantly, associated with greater reflection on the MFFT-I score and choice impulsivity indexed by the MCQ-K score. The two behavioural measures of impulsivity, however, not associated with one another.
Associations of impulsivity measures with substance use, problem gambling and personality disorder screening status. Higher BIS scores were associated with lifetime cocaine/crack use, problem gambling and a positive SAPAS screen. There were only isolated findings for the behavioural measures with a higher MFFT-I score associated with being a lifetime crack/cocaine use and higher MCQ-K associated with lifetime opiate use (see Table 2). We next examined associations of impulsivity measures with peak intensity of use in the subgroups of lifetime crack/cocaine or opiates user and found no significant correlations. A large majority of lifetime users of these substances reported intense past peak episodes of use (> 80% of participants reported taking the substance at least twice weekly) so there was limited variance on this measure.

Impulsivity measures as independent predictors of crack/cocaine use, problem gambling and SAPAS screening status. To determine if MFFT-I, MCQ-K and BIS had an independent relationship with substance use indicators we carried out a logistic regression analysis for cocaine/crack use as the dependent variable, with impulsivity variables (MFFT-I, MCQ-K and BIS) and SAPAS screening status as the predictor variables. We found the model was significant (Cox $R^2 = .26$, $\chi^2(1, N = 72) = 21.6$, $p < 0.001$) with BIS total explaining independent variance in lifetime crack/cocaine use but the other variables MFFT-I, SAPAS and MCQ-K did not reach significance as independent predictors (see Table 3). A logistic regression was used for prediction of those falling into the PGSI category of problem gambling but the overall model was not significant ($\chi^2(1, N = 72) = 4.6$, $p = .20$). The same was the case.
for the logistic regression model where impulsivity variables were entered as predictors of opiate use ($\chi(1, N = 72) = 4.3, p = .29$). Finally, we examined three impulsivity variables as predictors of screening status on the SAPAS. We found this model was significant ($\text{Cox R}^2 = .15, \chi(1, N = 72) = 11.4, p = .01$). The BIS total score was again the only independent predictor in the model with the MCQ-K and MFFT-I not significant (see Table 3).

**Discussion**

The first aim of the study was to establish the relationship between self-report and behavioural measures of impulsivity in a prison sample. We found limited overlap between the three impulsivity measures, which is consistent with the literature across both general population studies and those of substance using populations (Sharma et al. 2014). Specifically, we found that higher scores on behavioural measures of impulsivity were weakly associated with higher self-reported impulsivity. This is in the expected direction and contrasts with the previous study by Arantes et al. (2013). The weak correlations between self-report and behavioural measures are thought to reflect the lack shared method variance and demonstrate how both domains are necessary for sufficiently broad assessment of impulsivity (Sharma et al. 2014). In the current sample, where highly impulsive individuals are likely to have been over sampled, the pattern of associations between impulsivity measures was consistent with general population studies and no larger. Furthermore, we found the two behavioural measures of impulsivity were not associated with each other which supports the notion that the two measure tap onto distinct forms of impulsivity (Evenden, 1999).
Our second aim was to determine the association of these impulsivity measures with the extent of substance misuse, problem gambling and psychopathology in the prison sample. Increases in self-reported impulsivity were also related to an increased frequency of crack/cocaine use, problem gambling and a positive screen for personality disorder. These findings are again consistent with the large body of evidence supporting the association between self-reported impulsivity and substance use in both prisoners and general population samples that was reported in the introduction. However, while there was some evidence for associations between the behavioural measures of impulsivity and substance use these were less consistent than those with self reported impulsivity. Furthermore, behavioural measures of impulsivity not independent predictors of addictive behaviours or a personality disorder screen when regression models included both domains. Taken together these findings do not support the idea the behavioural measures capture any additional variance in impulsive behaviour in prisoners that is not already explained by a self-report measure.

The behavioural measures of impulsivity that were employed in the current study featured neutral materials that lacked direct relevance to value events or goals of the participants. Self-report measures, however, encourage the recall of daily life decisions where choices could have significant short-term benefits, potential risks and a more enriched emotional context. Cross, Copping, and Campbell (2011) suggest that there is a distinction between impulsivity that occurs in this type of ‘hot’ emotional context versus ‘cool’ situations. The behavioural measures use in the current study may reflect the ‘cool’ variety of impulsivity and therefore, be less likely to predict substance use behaviour as would presumably entail an emotionally charged decision making processes. Thus, our findings support the use of self-report
or interview based assessments that enable the elaboration of crucial details of
decision-making in a personally relevant context.

There were a number of limitations to the design of the study. The cross-
sectional design of the study prevents any causal interpretations of the any
relationship between impulsivity and addiction in prisoners. Causal relationships are
likely to be complex as there is evidence supporting the notion that impulsivity can
both act as a determinant of substances use but also a consequence of their use (de
Wit, 2008). Sampling for the current study within one prison setting also restricts how
these findings can be generalised to other prisoner groups. For instance recruitment
was undertaken from a specific cohort of prisoners, limiting the relevance of findings
to those not represented in the sample (e.g. those detained in non-Category C prisons,
female prisons or young offender institutions). It is important to note that this study,
in common with others on discounting in offending populations, could have been
subject to a desirability bias to present themselves in a positive light (Mills & Kroner,
2006). Paulhus (2002) suggests that those displaying a high in desirability bias should
demonstrate dissociation between self-report and behavioural measures of a construct.
The finding that these measures were associated in the current sample, albeit weakly,
suggests reporting biases were not so prominent to undermine the validity of the self-
report. Another consideration is for the use of a more conservative alpha as a control
for multiple comparisons. A number of correlations in between behavioural measures
of impulsivity were also small in magnitude. However, the effect sizes of these
predictors in the regression models were negligible.

In conclusion, we found that self reported impulsivity was associated with
impulsive behaviour and behavioural measures of impulsivity. The latter did not
explain independent variance in substance use or psychopathology. This is consistent
with findings of previous research in general population and clinical samples. The findings of this study need to be interpreted with greater caution than those with general population samples because of the limitations highlighted above in relation to potential impression management. However, we did not find clear evidence against the validity of self-report impulsivity, as there were associations with laboratory measures, albeit weak. This allows the tentative conclusion that previous research in prison samples, which has been dominated by studies using self-report measures, would not have been substantially improved by including potentially time-consuming behavioural tasks.
References


Center on Addiction and Substance Use. 2010. Behind Bars II: Substance Abuse and America's prison Population. USA.


Table 1: Sample Characteristics: age, Lifetime substance use, Problem Gambling Severity Index (PGSI) and Standardised Assessment of Personality Abbreviated Scale (SAPAS)

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Mean (SD)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age Mean (SD)</td>
<td>21.0 (2.1)</td>
</tr>
<tr>
<td>Lifetime substance use; % (n)</td>
<td></td>
</tr>
<tr>
<td>Binge Drinking</td>
<td>94 (68)</td>
</tr>
<tr>
<td>Cannabis</td>
<td>85 (61)</td>
</tr>
<tr>
<td>Lifetime Cocaine / Crack</td>
<td>40 (29)</td>
</tr>
<tr>
<td>Opiates</td>
<td>17 (12)</td>
</tr>
<tr>
<td>Problem Gambling; Mean (SD)</td>
<td></td>
</tr>
<tr>
<td>PGSI</td>
<td>2.1 (4.7)</td>
</tr>
<tr>
<td>Personality Disorder Screen; Mean (SD)</td>
<td></td>
</tr>
<tr>
<td>SAPAS</td>
<td>2.7 (1.4)</td>
</tr>
<tr>
<td>Impulsivity; Mean (SD)</td>
<td></td>
</tr>
<tr>
<td>Barrett Impulsivity Scale</td>
<td>64.7 (13.7)</td>
</tr>
</tbody>
</table>
Table 2: Correlation coefficients of associations between the Barrett Impulsivity Scale (BIS), Monetary Choice Questionnaire score K (MCQ-K), Matching Familiar Figures Test I score (MFFT-I), Lifetime Substance Use, Problem Gambling Severity Index (PGSI) and Standardised Assessment of Personality Abbreviated Scale (SAPAS)

<table>
<thead>
<tr>
<th></th>
<th>BIS</th>
<th>MCQ – K</th>
<th>MFFT – I</th>
</tr>
</thead>
<tbody>
<tr>
<td>Impulsivity measures</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>MCQ –K</td>
<td>.29*</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>MFFT</td>
<td>.26*</td>
<td>.02</td>
<td>-</td>
</tr>
<tr>
<td>Lifetime Substance use</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Cocaine / crack&lt;sup&gt;PB&lt;/sup&gt;</td>
<td>.43**</td>
<td>.16</td>
<td>.27*</td>
</tr>
<tr>
<td>Opiates&lt;sup&gt;PB&lt;/sup&gt;</td>
<td>.14</td>
<td>.24*</td>
<td>-.06</td>
</tr>
<tr>
<td>Problem Gambling</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>PGSI&lt;sup&gt;PB&lt;/sup&gt;</td>
<td>.24*</td>
<td>.07</td>
<td>.03</td>
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<tr>
<td>Personality Disorder Screen</td>
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<td></td>
<td></td>
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<tr>
<td>SAPAS&lt;sup&gt;PB&lt;/sup&gt;</td>
<td>.36**</td>
<td>.2</td>
<td>.01</td>
</tr>
</tbody>
</table>

Note: *p < .05, **p < .01.
PB Point biserial correlation
Table 3: Individual predictors from logistic regression models of crack/cocaine use and SAPAS screening status

<table>
<thead>
<tr>
<th>Predictors</th>
<th>Wald</th>
<th>Exp(B)</th>
<th>95% Confidence Interval</th>
<th>Significance</th>
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<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td>Lower</td>
<td>Upper</td>
</tr>
<tr>
<td>Lifetime Crack/Cocaine</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>BIS Total</td>
<td>5.6</td>
<td>1.1</td>
<td>1.0</td>
<td>1.2</td>
</tr>
<tr>
<td>MCQ-K</td>
<td>2.7</td>
<td>4.6</td>
<td>0.0</td>
<td>951</td>
</tr>
<tr>
<td>MFFT-I</td>
<td>.2</td>
<td>1.4</td>
<td>1.0</td>
<td>1.9</td>
</tr>
<tr>
<td>SAPAS</td>
<td>1.6</td>
<td>3.1</td>
<td>1.0</td>
<td>9.8</td>
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<tr>
<td>SAPAS Screen</td>
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<td></td>
<td></td>
</tr>
<tr>
<td>BIS Total</td>
<td>7.4</td>
<td>1.1</td>
<td>1.0</td>
<td>1.1</td>
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<tr>
<td>MCQ-K</td>
<td>.722</td>
<td>7.8</td>
<td>.1</td>
<td>876</td>
</tr>
<tr>
<td>MFFT-I</td>
<td>.46</td>
<td>.9</td>
<td>.67</td>
<td>1.2</td>
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</tbody>
</table>