Title: Measuring Adaptive Behavior in Autism: Minimal Clinically-Important Differences on the Vineland-II

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Lay Abstract Title: Establishing Meaningful Change in Adaptive Behavior

The Vineland Adaptive Behavior Scales (2nd edition; Vineland-II) is the most widely-used scale for assessing day-to-day “adaptive” skills. Scores on the Vineland-II are linked to educational attainment, need for support and independent living. Researchers use the Vineland-II to study the effects of new treatments for Autism Spectrum Disorder (ASD). However, it is not clear how much a person’s score on the Vineland-II has to increase after treatment for the change to be regarded as a truly clinically-significant improvement. To answer this question, we launched a collaboration across academia, patient advocacy groups, and industry. By pooling data from over 9000 individuals with ASD, we estimate that changes of 2 to 3.75 points on the Vineland-II Composite score represent the “minimal clinically-important difference”. Smaller changes are probably not meaningful. For the more specific Vineland-II domains measuring Socialization, Communication, and Daily Living Skills, score changes need to be numerically larger to be clinically-meaningful. These estimates will help in the evaluation of treatments for ASD.
Scientific Abstract (max 250 words):

Autism Spectrum Disorder (ASD) is associated with persistent impairments in adaptive abilities across multiple domains. These social, personal, and communicative impairments become increasingly pronounced with development, and are present regardless of IQ. The Vineland Adaptive Behavior Scales, Second Edition (Vineland-II) is the most commonly used instrument for quantifying these impairments, but minimal clinically-important differences (MCIDs) on Vineland-II scores have not been rigorously established in ASD. We pooled data from several consortia registries (EU-AIMS LEAP study, ABIDE-I, ABIDE-II, INFOR, Simons Simplex Collection and Autism Treatment Network [ATN]) and clinical investigations and trials (Stanford, Yale, Roche) resulting in a dataset of over 9,000 individuals with ASD. Two approaches were used to estimate MCIDs: distribution-based methods and anchor-based methods. Distribution-based MCID [d-MCID] estimates included the standard error of the measurement, as well as one-fifth and one-half of the covariate-adjusted standard deviation (both cross-sectionally and longitudinally). Anchor-based MCID [a-MCID] estimates include the slope of linear regression of clinician ratings of severity on the Vineland-II score, the slope of linear regression of clinician ratings of longitudinal improvement category on Vineland-II change, the Vineland-II change score maximally differentiating clinical impressions of minimal vs. no improvement, and equipercentile equating. Across strata, the Vineland-II Adaptive Behavior Composite standardized score MCID estimates range from 2.01-3.2 for distribution-based methods, and from 2.42 to 3.75 for sample-size-weighted anchor-based methods. Lower Vineland-II standardized score MCID estimates were observed for younger and more cognitively-impaired populations. These MCID estimates enable users of Vineland-II to assess both the statistical and clinical significance of any observed change.

Key Words: VABS, intelligence, assessment, treatment, efficacy, intellectual disability, ASD
Introduction

The diagnostic criteria for Autism Spectrum Disorder (ASD) include clinically-significant, functional impairments in social communication and social reciprocity, and the presence of restricted or repetitive behavior and sensory anomalies (DSM-5; APA, 2013), but ASD is also associated with broader impairments in adaptive behaviors that support everyday functioning (Kanne et al, 2011). These adaptive behavior problems span multiple domains including socialization, communication, and daily living skills, and are not fully accounted for by differences in cognitive ability (Charman et al, 2011; Klin et al, 2007). These impairments predict real-world outcomes in ASD, including educational attainment and the likelihood of independent living (Farley et al, 2009; de Bildt, Sytema, Kraijer, Sparrow & Minderaa, 2005). Adaptive behavior impairments in ASD are also associated with both the number of support services received, and the service needs that will go unmet (Taylor & Henninger, 2015). Thus, adaptive behavior is a key target for interventions directed at individual or societal outcomes in ASD (e.g., Veenstra-VanderWeele et al, 2016).

It is not currently possible to fully evaluate the efficacy of interventions directed at adaptive behavior. This is because efficacy claims require a demonstration of not only statistical significance, but also clinical meaningfulness (Jacobson & Truax, 1991; FDA PRO Guidance 2006; 2009; Coon & Cappelleri, 2016). While statistical significance is commonly evaluated, clinical meaningfulness requires that treatment effects exceed a minimally clinically-important difference (MCID). The term MCID was first described in 1989 (Jaeschke, Singer, & Guyatt, 1989) and is now a well-accepted concept to clinicians, regulators and payors to support evaluation of treatments (e.g., Guyatt et al, 2002; Food & Drug Administration, 2009; Medicare Evidence Development & Coverage Advisory Committee, 2006). However, there is currently no consensus on the MCID for adaptive behavior among individuals with ASD.

To remedy this gap, we calculated estimates of the MCID from a large sample of individuals with ASD who were assessed with the Vineland Adaptive Behavior Scales, Second Edition (Vineland-II; Sparrow et al, 2005) Survey Interview form. The Vineland-II is the most commonly used instrument for quantifying impairments in adaptive behaviors necessary for socialization, communication, and daily functioning (Pugliese et al, 2015; see also Figure 1A). Adaptive behaviors ranging from early developmental milestones to sophisticated demands on language understanding and attention are assessed. The scale can be used from birth to 99 years of age; since the expression of adaptive behavior changes across the lifespan, the score is age-normalized (much like IQ). This instrument
characterizes adaptive behaviors at multiple levels of granularity, including an overall composite score as well as domain- and subdomain-level constructs.

**Figure 1.** Structure of the Vineland-II and of the analysis strategy for estimating the MCID. A. The Vineland-II assesses 11 subdomains of adaptive behavior grouped into four domains; sums of these scores are standardized into Communication, Daily Living Skills, Socialization and Motor Skills domain standard scores. Sums of the domain standard scores are then standardized into the Adaptive Behavior Composite score. B. Individual-level data was pooled across multiple consortia, trials and registries, and subjected to several quality control steps. Next, both anchor-based and distribution-based techniques for estimating the MCID were performed on the entire pooled dataset, as well as stratified subsets of the pooled dataset. *Sample-size weighting is used, such that the number of
individuals contributing to each anchor- or distribution-based estimate is directly proportional to the influence of that estimate on the weighted-average; †Both cross-sectional and longitudinal assessments were performed.

The Vineland-II has been used widely in clinical, educational, and research settings, and with populations as diverse as ASD (e.g., Pugliese et al, 2015; Veenstra-VanderWeele et al, 2016; Kanne et al, 2015; Szatmari et al, 2015), Fragile X syndrome (e.g., Fisch et al, 1996), Williams syndrome (Greer et al, 1997), ADHD (Stein et al, 1995), low birth weight (Fjørtoft et al, 2015), Down syndrome (Dyckens, Hodapp & Evans, 2006), and other neurodevelopmental disorders. A recent Autism Speaks sponsored working group panel recommended the Vineland-II as suitable (having adequate reliability, validity and responsiveness) to quantify social communication deficits in clinical trials of ASD (Anagnostou et al., 2015; see also McConachie et al, 2015). Despite its widespread use as a comprehensive measure of adaptive behavior in developmental disorders, its MCID is as yet unclear. Previous estimates of clinically-meaningful change in the Vineland-II varied between 2.4 and 6 points (Frye et al, 2016), although this variability could reflect the relatively small, age-restricted sample (n=48 3-14 year-olds) or the use of only a single method of MCID estimation.

Very small changes on Vineland-II scores might be associated with clinically meaningful benefits, given that Vineland-II assesses adaptive behavior in real-world scenarios. Indeed, MCID estimates can (in principle) fall below the measurement error of the instrument itself (e.g., King 2011). If this were the case for the Vineland-II, then any statistically-significant change could also be considered clinically meaningful. In addition, the MCID could be dependent on other clinical factors like age or IQ. For example, a given change in Vineland-II scores might be associated with greater clinical benefit early in life, where adaptive behaviors may bootstrap across development. Conversely, MCID estimates might be larger or smaller among those with cognitive impairments: larger if substantive improvements in adaptive behavior are required before real-world functioning undergoes a meaningful change; or perhaps smaller, given the already significant special support provided to individuals with ASD and comorbid IQ deficits (e.g., Aljunied & Frederickson, 2011).

**Methods**

**Overview.** To estimate both the MCID magnitude and its potential dependence on these other clinical factors, we implemented a pipeline involving data pooling, Vineland-II standard score rederivation according to US norms,
quality control, and the calculation of both distribution-based and anchor-based MCID estimates (Figure 1B). These MCID estimation methods are consistent with recommendations from scientific literature and federal guidance alike (Revicki et al, 2008; King et al, 2011; Food & Drug Administration, 2009; Coon & Cappelleri, 2016; Crosby, Kolotkin & Williams, 2003). At a high-level, distribution-based methods compute the MCID as the change required to exceed some proportion of the intrinsic variability within the affected population; these estimates approximate those achieved through other methods (Norman et al, 2003). By contrast, anchor-based approaches calibrate scores across assessments – mapping from the target assessment to one or more “anchor” assessments where clinical meaningfulness has already been established. For anchor-based estimates, we utilize explicit assessments of clinically-relevant improvement, as quantified by the Clinical Global Impression scale (including its Severity [CGI-S] and Improvement [CGI-I] subforms; Guy, 1976) or, for patients where CGI scores were unavailable, the Ohio Autism Clinical Impressions Scale (OACIS; OSU Research Unit on Pediatric Psychopharmacology, 2005, including its General Level of Autism Severity [OACIS-S] and General Level of Autism Improvement [OACIS-I] subforms). These assessments were carried out only by clinicians, and not by caregivers or patients themselves. Alternative anchors were not evaluated due to lack of standardization across datasets (see also Discussion).

Data Pooling. Individual-level data was pooled from several consortia/registries (EU-AIMS LEAP, ABIDE-I, ABIDE-II, INFOR, Simons Simplex Collection, Autism Speaks Autism Treatment Network), observational studies (Stanford), and clinical trials (Stanford, Yale, Roche). Our pooled data is therefore a large sample of Western European and American individuals with ASD. Table 1 describes the studies from which individual-level data was drawn (see also Supplementary Table S1 for further details on each study). All studies used the Survey Interview Form, where adaptive behaviors are assessed in the context of a semi-structured interview with open-ended questions administered by a trained clinician. This form of the test is endorsed as the most reliable and accurate method for obtaining information on adaptive behavior (e.g., Sparrow et al, 2005).
### Table 1. Descriptives for pooled individual-level data.

<table>
<thead>
<tr>
<th>Dataset</th>
<th>N w/ Vineland</th>
<th>% male (n)</th>
<th>Data Type*</th>
<th>% w/ follow-up</th>
<th>mean Age (range)</th>
<th>mean IQ (range)</th>
<th>% w/ IQ (n)</th>
<th>% w/ follow-up &amp; IQ (n)</th>
</tr>
</thead>
<tbody>
<tr>
<td>AS ATN</td>
<td>5312</td>
<td>83% (4384)</td>
<td>L &amp; C</td>
<td>7% (368)</td>
<td>6.18 (1.16 - 17.76)</td>
<td>84.49 (30 - 150)</td>
<td>16% (856)</td>
<td>2% (107)</td>
</tr>
<tr>
<td>SSC</td>
<td>2747</td>
<td>86% (2373)</td>
<td>C</td>
<td>n/a</td>
<td>9.03 (4 - 18)</td>
<td>81.17 (7 - 167)</td>
<td>100% (2743)</td>
<td>n/a</td>
</tr>
<tr>
<td>BP28420</td>
<td>207</td>
<td>100% (207)</td>
<td>L &amp; C</td>
<td>90% (186)</td>
<td>25.46 (18.19 - 45.57)</td>
<td>98.27 (70 - 143)</td>
<td>98% (203)</td>
<td>90% (186)</td>
</tr>
<tr>
<td>LEAP</td>
<td>206</td>
<td>70% (145)</td>
<td>C</td>
<td>n/a</td>
<td>17.36 (6.24 - 30.15)</td>
<td>98.16 (40 - 148)</td>
<td>99% (204)</td>
<td>n/a</td>
</tr>
<tr>
<td>Fragxis</td>
<td>104</td>
<td>80% (83)</td>
<td>L &amp; C</td>
<td>99% (103)</td>
<td>23.68 (14.08 - 49.73)</td>
<td>51.83 (38.71 - 81.69)</td>
<td>91% (95)</td>
<td>90% (94)</td>
</tr>
<tr>
<td>INFOR</td>
<td>81</td>
<td>86% (70)</td>
<td>L &amp; C</td>
<td>86% (70)</td>
<td>17.28 (6.07 - 55.74)</td>
<td>93.02 (40 - 143)</td>
<td>47% (38)</td>
<td>36% (29)</td>
</tr>
<tr>
<td>ABIDEII</td>
<td>74</td>
<td>89% (66)</td>
<td>C</td>
<td>n/a</td>
<td>8.88 (5.13 - 34.76)</td>
<td>103.76 (67 - 143)</td>
<td>99% (73)</td>
<td>n/a</td>
</tr>
<tr>
<td>ABIDEI</td>
<td>68</td>
<td>85% (58)</td>
<td>C</td>
<td>n/a</td>
<td>14.91 (7.13 - 39.1)</td>
<td>108.35 (78 - 148)</td>
<td>100% (68)</td>
<td>n/a</td>
</tr>
<tr>
<td>Ox Simons/Stanford</td>
<td>62</td>
<td>76% (47)</td>
<td>C</td>
<td>n/a</td>
<td>8.14 (3.08 - 13)</td>
<td>84.93 (40 - 140)</td>
<td>92% (57)</td>
<td>n/a</td>
</tr>
<tr>
<td>Stanford PRT-G</td>
<td>48</td>
<td>79% (38)</td>
<td>L &amp; C</td>
<td>79% (38)</td>
<td>4.19 (2.3 - 6.9)</td>
<td>52.68 (19.37 - 83.12)</td>
<td>90% (43)</td>
<td>75% (36)</td>
</tr>
<tr>
<td>Stanford PRT-P</td>
<td>42</td>
<td>88% (37)</td>
<td>L &amp; C</td>
<td>93% (39)</td>
<td>3.87 (2.17 - 5.92)</td>
<td>46.73 (20.45 - 99.11)</td>
<td>88% (37)</td>
<td>81% (34)</td>
</tr>
<tr>
<td>YALE-ACE</td>
<td>42</td>
<td>100% (42)</td>
<td>C</td>
<td>n/a</td>
<td>13.19 (8 - 18)</td>
<td>101.58 (68 - 167)</td>
<td>100% (42)</td>
<td>n/a</td>
</tr>
<tr>
<td>YALE_PRT</td>
<td>28</td>
<td>57% (16)</td>
<td>L &amp; C</td>
<td>100% (28)</td>
<td>6.14 (4.5 - 9)</td>
<td>97.5 (44 - 127)</td>
<td>100% (28)</td>
<td>100% (28)</td>
</tr>
<tr>
<td>BP27801</td>
<td>19</td>
<td>100% (19)</td>
<td>C</td>
<td>n/a</td>
<td>24.02 (18.64 - 40.86)</td>
<td>99.95 (78 - 136)</td>
<td>100% (19)</td>
<td>n/a</td>
</tr>
<tr>
<td>BP28421</td>
<td>19</td>
<td>100% (19)</td>
<td>C</td>
<td>n/a</td>
<td>25.41 (19.39 - 39.76)</td>
<td>102.26 (71 - 130)</td>
<td>100% (19)</td>
<td>n/a</td>
</tr>
<tr>
<td>PNO_Stanford</td>
<td>8</td>
<td>88% (7)</td>
<td>L &amp; C</td>
<td>25% (2)</td>
<td>21.81 (18.08 - 33.33)</td>
<td>46 (40 - 52)</td>
<td>12% (1)</td>
<td>0% (0)</td>
</tr>
<tr>
<td><strong>Total</strong></td>
<td>9067</td>
<td>84% (7611)</td>
<td>L &amp; C</td>
<td>9% (834)</td>
<td>8.22 (1.16 - 55.74)</td>
<td>83.56 (7 - 167)</td>
<td>50% (4526)</td>
<td>6% (514)</td>
</tr>
</tbody>
</table>

*L=longitudinal, C=Cross-sectional*
Quality Control. This combined data underwent additional quality control steps, in addition to those already performed on each study independently. First, domain-level scores on the Vineland-II were recomputed based on Vineland-II raw scores (or v-scale, where raw was unavailable) as a function of the age at assessment. Second, to exclude potentially duplicated visits, we excluded any individuals who had precisely identical standardized scores across all three domains in any pair of follow-up visits. Third, the ranges of all variables were checked (both overall and by dataset) to exclude cases with out-of-range variables (e.g., Vineland-II standard scores outside the range of 20-160, or invalid ages). Fourth, within each dataset, any individuals with recomputed Vineland-II scores falling more than three times the interquartile range above or below the 25th or 75th percentiles (respectively) were excluded as extreme outliers. Overall, 2% of individuals were excluded from the pooled data using these criteria.

Stratification and Covariates. We stratified our combined sample by age (children [<13 years of age], adolescents [13 to <18 years of age] and adults [18 years and above]) and cognitive impairment (as defined by full-scale IQ scores < 70). Within each stratum and in an unstratified analysis (Figure 1B) we controlled for the fixed effects of age, sex, full-scale IQ, and the random effect of dataset where possible. We further adjusted for the effect of age bin, as used in the Vineland-II standard scoring procedure (Sparrow et al, 2005), and its interaction with age. For analysis of longitudinal distribution-based estimates, baseline Vineland-II scores were also used as a covariate. Finally, where regression-derived anchor-based estimates were calculated, potential differences in the administration of the anchor between studies (e.g., OACIS/CGI-I and OACIS-S/CGI-S) were captured by allowing the slope between Vineland-II and the anchor to differ as a random effect of dataset. The random effect of dataset is intended to account for differences related to the specific populations enrolled, the differing IQ tests and IQ scoring methods used, and other factors that may increase variability due to pooling individual-level data from many sources. Therefore, the MCID estimates generated were adjusted for the influence of sex, IQ, multiple influences of age and dataset, and (where relevant) baseline Vineland-II scores within each stratum. See Table 1 for differences between datasets. Each dataset contained a minimum of 10 observations for regression-based analyses.

Distribution-based MCID estimates. For each stratum with more than 30 distinct individuals, three types of distribution-based estimates were computed after adjusting for covariates described above (using R package “lme4”):
(1) the standard error of the measurement (SEM; Wyrich, Nienaber & Tierney 1999; Wyrwich, Tierney & Wolinsky, 1999; Wyrwich, Tierney & Wolinsky, 2002; Copay et al, 2007; Farivar, Liu, & Hays, 2004). This measure, distinct from the standard error of the mean, is a lower-limit estimate of scale measurement error for a given individual. It is computed as SD * [SQRT(1-r)], where r is the corrected test-retest reliability estimate reported by the Vineland-II manual.

(2) one half the standard deviation (Norman et al, 2003; Le, Doctor, Zoeliner & Feeny, 2013); and


Anchor-based MCID estimates. For each stratum, anchor-based estimates were computed using the following methods wherever more than 30 individuals could contribute to a given estimate:

(1) the slope of a linear regression in which OACIS-/CGI-S category predicted Vineland-II score and the slope of a linear regression in which OACIS-/CGI-I category predicted Vineland-II change score, computed where the Spearman rank correlation of Vineland-II and anchor (OACIS-/CGI-S or OACIS-/CGI-I) exceeded .3 (Revicki, Hays, Cella & Sloan, 2008; see also Fayers & Hays, 2014; Hays, Farivar & Liu, 2005).

(2) the Vineland-II change score which maximally (i.e., maximal [sensitivity + specificity]) differentiated OACIS-/CGI-I = 4 (“no change”) from OACIS-/CGI-I = 3 (“minimally improved”) patients, as determined by the receiver operating characteristics (ROC) curve (Froud & Abel, 2014)

(3) equipercentile equating (Leucht et al, 2006; Samara et al, 2014; Kolen & Brennan, 2014), whereby the percentile rank of each subject’s Vineland-II change score was used to find corresponding percentile rank on the OACIS-/CGI-I, thereby identifying the minimal Vineland-II score change that is associated with minimal improvement on the OACIS/CGI (using R-package “equate”).

Calculation of Weighted-average MCID estimates. For each stratum individually as well as for the non-stratified analysis, an average distribution-based MCID estimate was computed, weighted by the sample size for each MCID estimate; likewise, an average anchor-based MCID estimate was computed, again weighted by the sample size for each such MCID estimate.
**Results**

*The association of Vineland-II Adaptive Behavior to covariates*

The pooled and cleaned final dataset included 9,067 individuals with ASD and a total of 10,326 observations, ranging from 1.16 to 55.7 years of age; 84% of the subjects were males, and 72% of the subjects had IQ \( \geq 70 \).

Sample sizes stratified by age and the presence of IQ and OACIS/CGI are presented in Table 2.

**Table 2.** Sample size for anchor- and distribution-based analyses, by age and presence of IQ scores.

<table>
<thead>
<tr>
<th></th>
<th>Anchor based</th>
<th>Distribution based</th>
</tr>
</thead>
<tbody>
<tr>
<td># individuals (all ages)</td>
<td>421 (OACIS-/CGI-I), 627 (OACIS-/CGI-S)</td>
<td>9067</td>
</tr>
<tr>
<td># individuals (all ages) w/ IQ</td>
<td>365 (OACIS-/CGI-I), 527 (OACIS-/CGI-S)</td>
<td>4526</td>
</tr>
<tr>
<td># children (&lt;13 years) w/ IQ</td>
<td>91 (OACIS-/CGI-I), 160 (OACIS-/CGI-S)</td>
<td>3398</td>
</tr>
<tr>
<td># adolescents (13-18 years) w/ IQ</td>
<td>38 (OACIS-/CGI-I), 49 (OACIS-/CGI-S)</td>
<td>695</td>
</tr>
<tr>
<td># adults (18+ years) w/ IQ</td>
<td>236 (OACIS-/CGI-I), 318 (OACIS-/CGI-S)</td>
<td>433</td>
</tr>
</tbody>
</table>

As expected, individuals with ASD showed substantial adaptive behavior impairments across all domains and at all ages (Figure 2). The most impaired adaptive behavior domain was Socialization, followed by Communication, Daily Living Skills, and Motor Skills, each with statistically distinct degrees of impairment (Table 3; paired t-test with unequal variances; all \( p \)'s<4.3e-14). In addition to falling well below the age-normed value of 100 (one-sample t-test; all \( p \)'s<2.2e-16), all Vineland-II standardized scores also showed a pattern of increased impairment with age (Figure 2). A moderate floor effect was apparent among adults in the Communication and Socialization domains, driven largely by adults with IQ < 70 (see also Figure 3, below). In addition, an apparent inflection point was noted between the ages of 8 and 10 in Communication standardized scores, such that these scores might be seen to improve during early childhood before showing a subsequent decline after age 10. However, this pattern is no longer apparent once adjustments are made for the age bins used in the final stage of Vineland-II scoring (see Supplementary Figure S1 and Supplementary Text S1 for a fuller discussion of this issue), so we refrain from further interpretation.
Figure 2 depicts the Adaptive Behavior Composite and domain-level standard scores for Communication, Daily Living Skills and Socialization as a function of age in the combined sample. Points represent individual patients; vertical dotted lines demarcate the age bins of the normalization tables for Vineland-II standard scores (per Vineland-II user manual; Sparrow et al., 2005); red line represents a smoothing spline.

Full-scale IQ scores were highly variable among those with ASD (with a standard deviation 1.8 times larger than the expected value of 15; Table 3). However, variability in adaptive behavior was not as markedly elevated, remaining relatively close to the expected standard deviation of 15 (Table 3).

Table 3. Descriptive statistics of Vineland-II standardized scores and full-scale IQ in ASD.

<table>
<thead>
<tr>
<th>Score Type</th>
<th>Mean (95% CI)</th>
<th>SD (95% CI)</th>
<th>Skew (95% CI)</th>
<th>Excess Kurtosis (95% CI)*</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Area</td>
<td>Mean (95% CI)</td>
<td>SD (95% CI)</td>
<td>Correlation (95% CI)</td>
<td>Full-scale IQ (95% CI)</td>
</tr>
<tr>
<td>-------------------------------</td>
<td>---------------</td>
<td>-------------</td>
<td>----------------------</td>
<td>-----------------------</td>
</tr>
<tr>
<td>Adaptive Behavior Composite</td>
<td>71.58 (71.32-71.84)</td>
<td>12.63 (12.45-12.82)</td>
<td>-0.12 (-0.19--0.05)</td>
<td>0.83 (0.67-0.99)</td>
</tr>
<tr>
<td>Socialization</td>
<td>70.2 (69.94-70.46)</td>
<td>12.76 (12.58-12.95)</td>
<td>-0.07 (-0.15-0)</td>
<td>0.62 (0.46-0.79)</td>
</tr>
<tr>
<td>Daily Living Skills</td>
<td>75.05 (74.76-75.35)</td>
<td>14.39 (14.18-14.6)</td>
<td>0.08 (0.02-0.14)</td>
<td>0.37 (0.23-0.5)</td>
</tr>
<tr>
<td>Communication</td>
<td>74.19 (73.86-74.53)</td>
<td>16.32 (16.09-16.56)</td>
<td>-0.15 (-0.21--0.09)</td>
<td>0.4 (0.3-0.5)</td>
</tr>
<tr>
<td>Motor Skills</td>
<td>80.65 (80.29-81.01)</td>
<td>13.6 (13.35-13.86)</td>
<td>0.35 (0.28-0.43)</td>
<td>0.34 (0.15-0.51)</td>
</tr>
<tr>
<td>Full-scale IQ</td>
<td>83.56 (82.77-84.34)</td>
<td>27.01 (26.47-27.58)</td>
<td>-0.34 (-0.39--0.29)</td>
<td>-0.32 (-0.42--0.23)</td>
</tr>
</tbody>
</table>

* 95% CIs obtained by bootstrapping.

Vineland-II and full-scale IQ show a curvilinear association (Figure 3), such that the positive slope relating full-scale IQ and Vineland-II standardized scores becomes shallower at higher IQs (the same pattern is observed regardless of whether verbal, non-verbal, or full-scale measures of IQ are used; Supplementary Figure S2). This attenuation appears least pronounced in the Communication domain, where uniformity among individuals with ASD was found to be lowest (see above).
Figure 3 A-D, depicting Vineland-II standard scores as a function of IQ (x-axis), demonstrating the importance of stratification by IQ; red line represents a smoothing spline.

Sex-related differences in Vineland-II standardized scores were relatively minor. While sex differences in full-scale IQ were 5.6 points on average (95% CI= -7.868--3.33, males: 84.3 [range 7-167]; females: 78.7 [range 18-154] independent unequal variance Welch t(878)= -4.8428 p=1.5e-06), sex differences in Vineland-II scores were always less than half that magnitude, and were statistically significant only for the Daily Living Skills domain (mean difference of 1.8 (95% CI= -3--0.576, males: 75.9 [range 21-123]; females: 74.1 [range 25-136] independent unequal variance Welch t(876.6)= -2.8948 p=0.003; all other p’s > .16). Figure 4 depicts these sex-related differences alongside the larger IQ-related differences in adaptive behavior. (See also Supplementary Text S1, Figure S1, and Table S2 for further details on the association of sex and other covariates to the Vineland-II.)
Figure 4 depicts the distributions of Vineland-II standard scores as a function of sex and IQ; solid horizontal lines reflect sex-specific averages; the dotted line reflects the typically-developing mean of 100 in each Vineland-II standard score. Sex has minimal influence on the observed distributions, whereas the presence of cognitive impairment has a strong effect on the mean score.

MCID estimates.

Table 4 reports stratified, sample size-weighted estimates of both anchor-based and distribution-based MCID estimates for Vineland-II scores. For the Adaptive Behavior Composite score, these weighted anchor-based estimates ranged from 2.44 to 3.76; weighted distribution-based estimates were in good agreement with these
weighted anchor-based methods within each stratum, all falling within a half point of the weighted anchor-based estimates.

Good agreement was also observed between the weighted anchor- and distribution-based methods across the Vineland-II domain-level scores. The mean difference between weighted MCID estimates from these two methods was 0.19 points. There was a positive correlation between weighted distribution-based MCID estimates and the corresponding weighted anchor-based MCID estimates across the various strata and Vineland-II standardized scores for which weighted anchor-based MCID estimates could be calculated (n=21, as shown in Table 4; R=.47, p =.029; ICC(2,1) = .42; see also Supplementary Figure S7).

We assessed the robustness of these weighted MCID estimates using sensitivity analysis, including precision-weighting rather than sample size-weighting (Supplementary Table S3), excluding individuals with known syndromic forms of ASD (Fragile X syndrome; Supplementary Table S4), excluding data collected in countries where official non-USA norms are available (Supplementary Table S5), and evaluating alternative age groupings occasionally requested by regulatory authorities (e.g., defining adolescence as starting at 12 years of age and adulthood as starting at age 16 [FDA, 1996; EMA, 2001]; Supplementary Table S6) or utilized in early-life interventional studies (e.g., defining early childhood as age 0 to 5, and middle-to-late childhood as age 5 to age 12; Supplementary Table S7). Across these analyses, the weighted MCID estimates reported in Table 4 were stable, if slightly conservative: the sensitivity analyses revealed slight reductions in weighted MCID estimates on average in the affected strata, although the mean change was small (mean difference: -1.15, SD=.59).

These weighted MCID estimates were generally larger for the domain-level standard scores than for the Adaptive Behavior Composite. Weighted MCID estimates for the Socialization and Daily Living Skills domains were larger than the Composite weighted MCID (Wilcoxon Signed Rank Test p=.001 and p=.019, respectively), although this same effect was detected at only a trend level for the Communication domain (Wilcoxon Signed Rank Test p=.08).

The weighted MCID estimates tended to increase with IQ: they were lower in the IQ<70 strata in all 21 of the comparisons enabled by Table 4 (Wilcoxon Signed Rank Test p< 0.0001). Weighted MCID estimates also generally increased with age: they tended to be lower for children than for adolescents (Wilcoxon Signed Rank Test p=.02), and lower for adolescents than adults (Wilcoxon Signed Rank Test p=.04).
These fully-stratified analyses were more consistent than partially-stratified or non-stratified analyses. In non-fully stratified analyses, discrepancies between weighted distribution- and anchor-based MCID estimates could exceed two points. Indeed, anchor- and distribution-weighted averages failed to reliably correlate in analyses that were not fully stratified ($R = .11$; Supplementary Figure S7).

Table 4. Sample size-weighted MCID estimates for the Adaptive Behavior Composite score and each domain standard score, stratified by age and IQ.

<table>
<thead>
<tr>
<th>Adaptive Behavior Composite</th>
<th>Sample Size-Weighted Anchor-based MCID estimates</th>
<th>Sample Size-Weighted Distribution-Based MCID estimates</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>$IQ &lt; 70$</td>
<td>$IQ \geq 70$</td>
</tr>
<tr>
<td>Children (0-13)</td>
<td>2.5</td>
<td>2.9</td>
</tr>
<tr>
<td>Adolescents (13-18)</td>
<td>2.4</td>
<td>n/a*</td>
</tr>
<tr>
<td>Adults (18+)</td>
<td>2.8</td>
<td>3.8</td>
</tr>
<tr>
<td>Communication Domain</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Children (0-13)</td>
<td>2.0</td>
<td>3.1</td>
</tr>
<tr>
<td>Adolescents (13-18)</td>
<td>3.2</td>
<td>n/a*</td>
</tr>
<tr>
<td>Adults (18+)</td>
<td>3.1</td>
<td>3.3</td>
</tr>
<tr>
<td>Socialization Domain</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Children (0-13)</td>
<td>2.6</td>
<td>3.1</td>
</tr>
<tr>
<td>Adolescents (13-18)</td>
<td>2.9</td>
<td>5.5</td>
</tr>
<tr>
<td>Adults (18+)</td>
<td>3.0</td>
<td>5.5</td>
</tr>
<tr>
<td>Daily Living Skills Domain</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Children (0-13)</td>
<td>2.6</td>
<td>3.7</td>
</tr>
<tr>
<td>Adolescents (13-18)</td>
<td>1.7</td>
<td>n/a*</td>
</tr>
<tr>
<td>Adults (18+)</td>
<td>3.0</td>
<td>5.1</td>
</tr>
</tbody>
</table>

* - Among studies including adolescents with $IQ \geq 70$, anchors met the inclusion criterion of a correlation above .3 only with Socialization standard scores (see Methods for further details on the correlation criterion). Consequently, anchor-based MCIDs for Daily Living Skills, Communication, and the Adaptive Behavior composite are not available (n/a) for this stratum.

The individual, unweighted estimates contributing to the above anchor- and distribution-based weighted-average MCID estimates are presented in Supplementary Figures S3-6 for the Adaptive Behavior Composite, Communication, Socialization, and Daily Living Skills domains (respectively), and are presented in tabular form in Supplementary Table S8.

Discussion
Individuals with ASD increasingly fall behind their peers in adaptive behavior with age. Using the largest sample of adaptive behavior among individuals with ASD to date, we identified the most impaired domain of adaptive behavior as Socialization, followed by Communication, Daily Living Skills, and Motor Skills. We quantify the partial dependency of these impairments on factors like sex, age, and IQ, enabling more precise and powerful assessments of adaptive behavior in ASD (e.g., Lerer et al., 2008).

Pronounced age- and IQ-related shifts are observed in all domains of adaptive behavior among individuals with ASD, with comparatively minor effects of sex (consistent with some prior results; see Howe, Yatchmink, Viscidi & Morrow, 2014). Despite these effects, we find that adaptive behavior impairments are also much more uniform than those observed in full-scale IQ (see Table 3). ASD is considered a spectrum disorder because it is highly heterogeneous (e.g., Beversdorf & MAS Consortium, 2016; Szatmari et al, 2015), yet we find a greater uniformity of adaptive behavior in ASD with persistent impairment across demographic variables like IQ and sex (see Results). This striking uniformity reflects the fact that, even when intellectual ability is intact, adaptive behavior is still impaired – particularly in terms of Socialization and Daily Living Skills (Charman et al, 2011; Klin et al, 2007; Duncan & Bishop, 2013). The curvilinear association between full-scale IQ and adaptive behavior suggests that IQ does not fully explain adaptive behavior impairments among individuals with ASD, particularly at high IQs where the association between IQ and adaptive behavior is attenuated. In summary, given that adaptive behaviors are assessed in real-world, everyday scenarios and represent an area of specific, disproportionate impairment (relative to IQ), the Vineland-II is expected to be an important measure of interventional efficacy in ASD.

To support evaluations of efficacy using the Vineland-II interview form, we used both distribution-based and anchor-based methods to establish estimates of clinically meaningful improvement (as recommended by Coon & Cappelleri, 2016; Crosby, Kolotkin & Williams, 2003; and FDA, 2006, 2009). When stratifying by age and IQ, the sample size-weighted MCIDs produced good agreement both with one another (a half-point or less of discrepancy in the Adaptive Behavior Composite score) and with prior work (Frye et al, 2016’s sample of 48 children with ASD using a single method of estimation, the standard error of the measurement). Anchor- and distribution-based weighted MCID estimates were also positively correlated, suggesting both agreement between these methods, as well as systematic variation in the weighted MCID estimates across strata and Vineland-II standardized scores.
One source of systematic variation is the type of standardized score that is used. Weighted MCID estimates were generally larger for any single domain of adaptive behavior as compared to the Adaptive Behavior Composite. This pattern does not mean it is somehow “easier” to achieve a clinically-meaningful effect on the Adaptive Behavior Composite: the criterion for clinical meaningfulness is merely numerically smaller. In other words, a given numerical change on the Adaptive Behavior Composite is more likely to be clinically meaningful than a numerically-identical change that is limited to any given domain (e.g., Communication, Socialization, or Daily Living Skills), and not that numerically-identical changes are equally easy to achieve across domains. These differing MCID estimates highlight the psychometric independence of the Adaptive Behavior Composite from the domain-level standardized scores that contribute to it (e.g., Sparrow et al 2005). Interventions which do not exceed the MCID on each domain individually could still theoretically exert a clinically-meaningful effect on the Adaptive Behavior Composite, in aggregate.

A second source of systematic variance in Vineland-II weighted MCID estimates relates to the effects of age and IQ, with demonstrably lower MCID estimates for the young and those with cognitive disability. These systematic differences in the MCID estimates do not imply that any given intervention would be more clinically beneficial early in development or among those with cognitive disability, nor do they obviate the need for empirical testing of a given intervention in each population of interest. The numerical criterion for clinical meaningfulness in adaptive behavior is smaller in these groups, but this does not obviate empirical testing of a given intervention in each population of interest. Indeed, the dependence of MCID estimates on these stratifications underscores the methodological importance of stratification for the design and analysis of interventions influencing adaptive behavior.

Large-scale analysis of pooled individual-level data can attenuate the sampling and ascertainment biases affecting any given study (e.g., Howe, Yatchmink, Viscidi & Morrow, 2014), but pooling can also inflate heterogeneity. Here this concern is mitigated by the reduced variance in Vineland-II scores we observe, relative to normalization samples (Sparrow et al, 2015); the robustness of our estimates across various pooling strategies (see Supplementary Tables S2-S6); the statistical modeling of random effects of dataset (see Methods: Stratification and Covariates); and the orderly associations observed between age and Vineland-II scores. Relatedly, our anchors (CGI and OACIS) did not uniformly show robust associations with Vineland-II scores, though again the observed agreement between
stratified, weighted distribution-based and anchor-based MCID estimates may lessen this concern. When making these global assessments, clinicians are likely to incorporate features beyond those assessed in the Vineland-II (e.g., eye contact, hand flapping, etc); future work with the Vineland-II should confirm that effects exceeding our weighted MCID estimates are viewed as important by individuals with autism themselves. Future work may also investigate whether the current weighted MCID estimates apply to alternative forms of the Vineland-II (including the Vineland-II Parent/Caregiver Rating Form, Expanded Interview Form, the Teacher Rating Form, and the recently published Vineland-3; Sparrow, Cichetti & Saulnier, 2016), or even to the clinical meaningfulness of biomarkers associated with adaptive behavior (e.g., Lerer et al., 2008).

Interventions exceeding the MCID are often said to be those which would mandate a change in a patient’s treatment, (in the absence of other countervailing factors like treatment cost, potential side effects, etc). However, interventions falling short of the MCID at the group-level may nonetheless still drive meaningful change in a subset of individuals. The anchor-based MCID estimates we report follow guidelines for defining clinically-meaningful differences at the individual-level (as emphasized in the FDA PRO guidance, 2009). Analogously, the MCID estimates reported here may be used to identify subgroups of patients with meaningfully distinct degrees of clinical severity, or to identify biomarkers that index clinically meaningful variation (e.g., Marquand, Wolfers, Mennes, Buitelaar & Beckmann, 2016; Lerner et al., 2008). Vineland MCID estimates may therefore play a broader role, enabling more personalized and biologically-informed treatments for ASD (e.g., Loth, Murphy & Spooren, 2016).

These MCID estimates for adaptive behavior in ASD may be useful for prioritizing the development of promising new interventions. MCID estimates are also acknowledged to play a role in the decision making processes of regulators and payors relating to the potential value of new interventions (see Rennard, 2009 for a specific example in pulmonology, or Hammad & Neyarapally, 2016 for a broader discussion). In those contexts, MCID estimates are not a replacement for directly establishing the meaningfulness (or lack thereof) in any given individual. For example, a sizable proportion of individuals can still experience meaningful improvement even when the MCID is not exceeded on average.

Conclusions
Our results provide robust, stratified MCID estimates for Vineland-II scores obtained from a large sample of individuals with ASD, and underscore the importance of defining population-specific MCID estimates using multiple methods. Agreement between weighted distribution- and anchor-based methods was highest in the fully stratified analysis, supporting the use of age- and IQ-appropriate MCID estimates in evaluating new interventions for efficacy. MCID estimates for the Vineland-II Adaptive Behavior Composite standardized score range from 2-3.8 points for the Adaptive Behavior Composite depending on stratum. Clinical efficacy as measured by the Vineland-II is generally more likely to be seen with early intervention, and among those with impaired intellectual abilities, for a given treatment-related change on the Vineland-II. The developmentally-persistent and relatively uniform (as compared to IQ) profile of adaptive behavior in ASD suggests that adaptive behaviors are a powerful and important dimension of clinically-meaningful variation in this condition. These estimates of the MCID represent important benchmarks by which new treatments for ASD can be evaluated.
Acknowledgments and Conflicts of Interest

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