Psychosocial Factors Associated with Persistent Pain in People with HIV: A Systematic Review with Meta-Analysis

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Abstract

Chronic pain remains a prevalent and disabling problem for people living with the human immunodeficiency virus (HIV) in the current antiretroviral (ART) treatment era. Psychosocial treatments may have promise for managing the impact of this pain. However, research is needed to identify psychosocial processes to target through such treatments. The current systematic review and meta-analysis examined the evidence for psychosocial factors associated with pain, disability, and quality of life in people living with HIV and persistent pain. Observational and experimental studies reporting on the association between one or more psychosocial factor and one or more pain-related variable in an adult sample of people living with HIV and pain were eligible. Two reviewers independently conducted eligibility screening, data extraction, and quality assessment. Forty-six studies were included in the review and 37 of these provided data for meta-analyses (12493 participants). ‘Some’ or ‘moderate’ evidence supported an association between pain outcomes in people with HIV and the following psychosocial factors: depression, psychological distress, post-traumatic stress, drug abuse, sleep disturbance, reduced ART adherence, healthcare use, missed HIV clinic visits, unemployment, and protective psychological factors. Surprisingly few studies examined protective psychological factors or social processes, such as stigma. There were few high quality studies. These findings can inform future research and psychosocial treatment development in this area. Greater theoretical and empirical focus is
needed to examine the role of protective factors and social processes on pain outcomes in this context. The review protocol was registered with PROSPERO (CRD42016036329).

**Key Words:** HIV, pain, systematic review, psychosocial factors

**Introduction**

The human immunodeficiency virus (HIV) remains a significant global health concern with 36.7 million people living with HIV worldwide [130]. The availability of combined antiretroviral therapies (cART) has drastically improved life expectancy [9,93,120]. In well-resourced countries, and increasingly in less well-resourced regions, the shift in HIV from a terminal illness to a chronic condition has led to a focus on disease and symptom management [59].

Chronic pain is a common symptom in people with HIV. Data from one systematic review indicate that 54 to 83% of people with HIV may experience clinically-meaningful persistent pain, and these estimates appear stable from the pre- to current cART era [80].

Neuropathic pain is a frequent complication of HIV and/or antiretroviral therapy. Approximately 42 to 66% of people with HIV have peripheral sensory neuropathy (HIV-SN), and around 54-78% of these experience neuropathic pain [84,88,129]. Importantly, pain in people with HIV is associated with increased disability and reduced quality of life [27].

There are few pharmacological options for managing chronic HIV-related pain. A systematic review of 19 randomized controlled trials (RCTs) of pharmacotherapy for painful HIV-SN found efficacy only for topical capsaicin, smoked cannabis, and subcutaneous nerve growth factor [85]. However, nerve growth factor is not clinically available, capsaicin is not feasible in lower-resourced settings, and a subsequent review of cannabis showed no effect.
on neuropathic pain and concerns about long-term side-effects [32]. Additional negative RCTs of pregabalin, capsaicin, and amitriptyline have been published [17,24,103].

In the wider literature, psychological approaches are common in chronic pain management [34]. Psychological treatments, including cognitive-behavioural therapy (CBT), are associated with improved functioning and mood for chronic pain that is primarily musculoskeletal [124]. However, research on psychological treatments for pain in HIV is less well developed. Only two RCTs have examined CBT for people with HIV and chronic pain, but interpretation of these trials is hampered by small samples [118] and high drop-out rates [29]. An observational study of CBT for HIV-related pain showed similarly poor treatment completion [21,113]. There is a clear need for improving psychological treatments for people with HIV-related pain.

Improving psychological approaches for chronic pain in HIV will require consideration of the psychosocial complexities associated with HIV. For example, stigma, mental health problems, and substance abuse may influence pain and treatment engagement in people with HIV [36,69,71,111,123,132]. However, research has not systematically examined psychosocial factors associated with pain in this context. The systematic review by Parker et al. (2014) which estimated the prevalence of pain in HIV described five studies reporting psychosocial factors. However, that review did not specifically include assessment of psychosocial factors in the eligibility criteria. Furthermore, 33 potentially eligible studies were excluded due to low quality ratings [80], which limits our understanding of the range of psychosocial factors examined in this context. Therefore, we conducted a systematic review and meta-analyses to examine the associations between psychosocial factors and persistent pain in HIV. As the aims of the review were exploratory, we did not formulate specific hypotheses about the associations between these variables.
Methods

The review protocol was registered with PROSPERO (http://www.crd.york.ac.uk/PROSPERO/display_record.php?ID=CRD42016036329).

Inclusion and Exclusion Criteria

Inclusion:

1) People with HIV ≥ 18 years old.

2) The original protocol specified the study must have a ‘(sub)sample with average pain duration of ≥3 months’. After piloting this criterion, a large number of studies did not define or report pain duration. We contacted authors to enquire about pain chronicity; however, these data were generally not available. Given the potentially high prevalence of chronic pain in HIV [80], we decided to include studies with (sub)samples of ambiguous pain duration, provided that “pain” versus “no pain” subgroup analyses were reported in studies for which chronic pain was not an eligibility criterion or which did not report pain duration.

3) Data on presence of pain, pain intensity, functioning, and/or quality of life.

4) Data on one or more psychosocial variable, representing any potentially modifiable cognitive, affective, behavioural, or interpersonal process. Adherence to antiretroviral therapy and healthcare use variables represent modifiable behaviour patterns. Therefore, we considered these as psychosocial variables eligible for this review.

5) Observational (cross-sectional, case-control, or prospective) or experimental studies (RCTs) reporting between- or within-groups associations between pain and psychosocial variables in a (sub)sample with pain.

7) Any language, from any region, from 1981 onwards (the date that HIV was identified in the literature).
8) Studies (published and unpublished) with an available full-text. Where only abstracts or trial registration summaries were available, the authors provided unpublished data for the review. Unpublished studies are commonly included in systematic reviews given recognition of overestimation of effects in published research [62]. Additionally, studies conducted in lower-resourced countries where HIV is particularly prevalent may not always proceed through to publication. Therefore, the inclusion of unpublished studies and dissertations may help overcome this disparity and allows us to consider potential contextual differences.

Exclusion:

1) Studies only measuring associations between unchangeable demographic factors (e.g., age, ethnicity) and pain. While piloting the eligibility criteria, we found a number of studies reporting history of injecting drug use (IDU) as participants’ HIV risk factor. Given lack of further information about substance abuse history or current abuse, we excluded studies for which IDU history was the only psychosocial factor. Likewise, we excluded studies reporting only average units of alcohol consumed, rather than alcohol abuse.

2) Qualitative studies.

Search Strategy:

We searched the following databases during March 2016: Medline, EMBASE, CINAHL, PsycINFO, Cochrane, and Web of Science. We also searched ISRCTN, clinicaltrials.gov, and EU Clinical Trials Register. Reference lists of eligible studies were searched and key authors were contacted. We re-ran the search in August 2017. The search included terms for the target population (HIV or AIDS), outcome (chronic pain), and exposure measurement (psychosocial factors). Relevant search terms were identified from previous reviews on pain in HIV [80,124], psychosocial factors in HIV [99], and psychosocial factors in chronic pain [40] (Appendix A).
Data extraction

Two reviewers (WS and CA) independently screened titles/abstracts and full-texts for eligibility. The following data were extracted from eligible studies: year; design; country; sample size; demographics (i.e., age, sex, and race/ethnicity); clinical factors (i.e., HIV duration, use of ART, CD4+ count and viral load, pain duration and type); assessment of pain and psychosocial variables; and, statistical analyses. In cases where both cross-sectional and prospective data reported the same (or an overlapping) cohort and variables, the prospective analyses were extracted. Data were extracted from all studies by WS, and independently by CA and KK who each extracted data from approximately half of the studies. Disagreements regarding eligibility and data extraction were discussed to reach consensus and, where discrepancies remained, WS discussed these with the wider team. The reviewers were not blinded to the authorship of the studies reviewed.

Quality Assessment

We assessed the methodological quality of studies using an adapted version of quality assessment tools used in previous systematic reviews of observational studies relevant to pain and HIV [3,40,80]. The quality assessment tool contained items assessing: study purpose, recruitment, response rates, sample description, assessment measurements, data analysis, and confounding/matching (Appendix B). Additional items assessed features specific to prospective designs. Thus, quality scores differed for cross-sectional and prospective studies. In some cases, the overall study design did not correspond to the nature of the data extracted for the purpose of this review. Additionally, in some cohorts a cross-sectional design was used to examine one psychosocial variable, while a prospective design was used to examine another variable using the same sample. In all cases, the quality assessment was applied to the design used for the nature of the data extracted for a given
psychosocial variable. Quality assessment items were rated as ‘positive’ (1), ‘negative’ (0), or ‘unclear’ (?), and total scores were computed and classified as low (<50%), medium (50–80%), and high (>80%) [3,40]. WS completed quality assessment ratings for all studies, while CA and KK each independently completed the quality assessment for approximately half of the studies. The strength of evidence was assessed according to the levels outlined by Ariens et al. in a systematic review of observational studies of psychosocial risk factors for neck pain [3]: 1) Strong: Consistent results in multiple high quality prospective and/or case-control studies; 2) Moderate: Consistent results in multiple prospective and/or case-control studies; 3) Some Evidence: Findings in one prospective or case-control study, or consistent findings in multiple cross-sectional studies with at least one high quality study; and 4) Inconclusive: Inconsistent findings in multiple studies or consistent findings in multiple low quality cross-sectional studies [3].

Data Synthesis

Meta-analyses were conducted using Stata 15.0 where there were at least two studies [43] with the same design and effect estimate of the association between the same pain (e.g., intensity) and psychosocial variables (e.g., depression). We took a broad approach to the meta-analyses [37,45], and grouped psychosocial variables on the basis that they reflected conceptually similar underlying constructs with overlapping measurement content. All analyses were conducted using random effects given likely heterogeneity [43]. Between-study heterogeneity ($I^2$ statistic) was interpreted as low (<25%), medium (25-50%), and high (>50%) [124]. For between-groups comparisons of continuous data, means, standard deviations, and sample sizes were extracted to compute the pooled standardized mean difference (SMD). For between-groups comparisons of dichotomous data, events data and sample sizes were extracted. Where events data were not reported, odds or hazard
ratios, 95 percent confidence intervals (95% CIs), and sample sizes were extracted. To aggregate studies reporting a mixture of odds ratios and events data, odds ratios were first computed from studies reporting events data and then pooled with odds ratios reported in other studies. Odds ratios and hazard ratios were analysed separately. Where applicable, correlation coefficients (Pearson’s r) were extracted with sample sizes. We transformed r to Fisher’s z and computed 95% CIs of z to compute the pooled estimate [18,100].

Data extracted were from bivariate analyses. Multivariate data (e.g., adjusted odds ratios) were only extracted where bivariate data were not available. We focused on bivariate analyses, as many studies did not report multivariate analyses. Moreover, studies that reported multivariate models varied substantially with respect to control variables included, and inconsistently used psychosocial variables as independent or dependent variables. Taken together, these differences limit meaningful interpretation of multivariate analyses across studies.

Several studies presented data on more than two pain/no pain groups, often with idiosyncratic group definitions, which limited our ability to compare studies. Where studies reported three or more groups, we collapsed these into two to represent groups with and without pain (e.g., frequent/moderate/severe versus infrequent/mild/none), and computed effects between these. This approach facilitated more direct comparison across studies and thus enabled us to include a larger number of studies in the analyses. For studies comparing participants on the presence of neuropathy, we prioritised extracting data from these comparisons in the following order depending on the data reported: 1) painful versus non-painful neuropathy; 2) painful neuropathy versus no neuropathy; and 3) neuropathy versus no neuropathy. Where there was more than one measure of the same variable, we extracted data for the measure with the widest usage or the longer measure to increase
reliability [124]. Our protocol specified that funnel plots would be inspected to assess for publication bias. However, due to the relatively small number of studies in each meta-analysis and the likelihood of high heterogeneity, inspection of funnels plots was not appropriate [109] and, therefore, was not undertaken.

We conducted sensitivity analyses to examine the influence of the following study and patient characteristics on the findings: certainty of pain chronicity, pain type, immune functioning and viral suppression, ART treatment era, and healthcare system. With the exception of the pain chronicity analysis, these sensitivity analyses were pre-specified. Given the large number of potential analyses, we restricted sensitivity analyses to the between-groups SMDs for depression, as this was the analysis with the largest number of studies.

**Results**

Forty-six studies were included in the review (13480 participants) and 37 of these provided data for meta-analyses (12493 participants). Most (83%) were conducted in the United States (US), with four studies from South Africa [79,88,121,122], and one each from the United Kingdom (UK) [84], Thailand [89], Uganda [95], and Russia [114]. Participants were primarily recruited from HIV clinics, or using multifaceted strategies that also included recruitment from substance abuse clinics and community outreach. One study recruited exclusively from a methadone clinic [8], while two others recruited in high poverty areas [39,110]. The samples were comprised predominantly of men in 41 studies, with the proportion of men in these studies ranging from 51% [50] to 100% [28,104]. Five studies (four from South Africa, one from the US) recruited women exclusively [79,92], or predominantly (proportion of women ranging from 72-88%) [88,121,122]. The mean age ranged from 30.1 (SD=5.2) [114,115] to 51.0 (SD=9.3) years [119]. HIV duration was not
consistently reported; however, of the studies providing data, duration ranged from 2.09 (SD=1.22) [102] to 16.95 years (SD=8.70) [119]. Eighteen studies (39%) reported on mixed HIV/AIDS samples (reported proportion with AIDS ranged from 10-74%). Four studies included only participants with AIDS, one study excluded patients with AIDS, while 23 studies did not clearly report the proportion (if any) with AIDS. Supplemental Table 1 shows further demographic characteristics of the study samples (available at http://links.lww.com/PAIN/A643).

Table 1 provides a summary of study designs, quality, and evidence level for each psychosocial factor. The studies showed substantial variability in the measurement of pain and psychosocial variables. Most studies (63%) were medium quality. Fifteen studies were low quality, and only two were high quality [79,84]. The most common limitations included unclear reporting of response rates, no a priori sample size justification, and poor reporting of HIV and pain characteristics. There is no single agreed upon strategy to best address low quality studies within meta-analyses, an issue which is compounded by the arbitrary nature of study quality scoring and cut-off points [43]. This can be dealt with by only including high quality studies, performing sensitivity analyses, or including all studies irrespective of quality and discussing risk of bias [43]. Given that only 2 of 46 studies were rated as high quality, a meta-analysis of these cannot be regarded as reflecting most of the studies. Sensitivity analyses would likewise not be meaningful. Including all studies is thus the most justifiable approach for the current data. Although we have chosen to focus on data from bivariate analyses for reasons outlined in the Methods, studies that reported a multivariate model of the association between psychosocial and pain variables are shown in bold in Table 1 for ease of reference.
Depression

Depression was the most frequently assessed psychological variable, investigated in 29 studies. Two prospective studies reported hazard ratios for baseline depression predicting time to onset of symptomatic neuropathy. The pooled hazard ratio was significant and indicated that baseline depression was more severe in participants who developed symptomatic neuropathy at follow-up than those who did not: HR=1.04 (95% CI 1.02-1.07), z=3.23, \( p = 0.001 \) (Supplemental Figure 1, available at http://links.lww.com/PAIN/A643). Heterogeneity was 0.0%. Two further prospective studies reported odds ratios. The pooled odds ratio was significant and indicated that higher baseline depression symptoms were associated with greater likelihood of follow-up pain: OR=2.26 (95% CI 1.47-3.47), z=3.72, \( p < 0.001 \) (Supplemental Figure 2, available at http://links.lww.com/PAIN/A643). Heterogeneity was medium (40.1%). Nine cross-sectional studies provided events data or odds ratios. The pooled odds ratio was significant such that depression was more likely in participants with versus without pain: OR=2.65 (95% CI 1.62-4.34), z=3.90, \( p < 0.001 \) (Figure 2). Heterogeneity was high (83.0%). Twelve cross-sectional studies provided data to compute SMDs (Figure 3). The overall effect was significant and showed moderately greater depression in participants with versus without pain: SMD=0.68 (95%CI 0.42-0.93), \( z = 5.22, p < 0.001 \). Heterogeneity was high (\( I^2 = 89.2\% \)). Another cross-sectional study which reported the median and interquartile range found no difference in depression between groups with (n=125) and without pain (n=72) [88].

Six cross-sectional studies reported correlation coefficients between depression and pain severity. The pooled correlation was small, but significant: Fisher’s \( z=0.26 \) (95% CI 0.18-0.33), \( z=6.77, p < 0.001 \). Heterogeneity was 0.0%. One additional study found a non-significant correlation, although the coefficient was not reported [27]. Four cross-sectional
studies reported correlations between depression and pain interference/disability. The pooled correlation was moderate: Fisher’s $z=0.48$ (95% CI 0.41-0.56), $z=12.48$, $p<0.001$. Heterogeneity was 0.0%. Three cross-sectional studies reported correlations between depression and quality of life. The pooled correlation was large and significant: Fisher’s $z=-0.52$ (95% CI -0.75 to -0.30), $z=4.51$, $p<0.001$. Heterogeneity was high at 73.3% (all correlation analyses, Supplemental Figure 8, available at http://links.lww.com/PAIN/A643). One final cross-sectional study (n=120) reported a moderate correlation between pain presence and depression [55].

**Depression Sensitivity Analyses**

We conducted sensitivity analyses on the SMDs for depression between pain and no pain groups (Supplemental Figures 3-7, available at http://links.lww.com/PAIN/A643). We excluded data from the Kirkland study here, as the SMD from this study was substantial and appeared to be driving heterogeneity in the primary meta-analysis. Excluding the Kirkland data reduced heterogeneity from $I^2=89.2\%$ to 52.2%. Thus, eleven studies were included in sensitivity analyses.

The pooled SMD for depression was medium for studies with certain (0.61, 95% CI 0.11-1.12, $z=2.39$, $p=0.02$; $I^2=75.1\%$) and uncertain pain chronicity (0.53, 95%CI 0.41-0.64, $z=9.04$, $p<0.001$; $I^2=34.5\%$). In the analysis by pain type, the pooled SMD for depression was moderate in studies with mixed pain types (0.75, 95% CI 0.58-0.92, $z=8.51$, $p<0.001$; $I^2=31.7\%$) and for which pain type was not reported (0.52, 95%CI 0.39-0.65; $z=7.91$, $p<0.001$; $I^2=0.0\%$). In contrast, studies with neuropathic pain (0.31, 95%CI 0.11-0.52; $z=2.96$, $p=0.003$; $I^2=0.0\%$) or headache (0.38, 95%CI 0.06-0.70, $z=2.34$, $p=0.02$) showed small but significant differences between groups on depression.
The pooled SMD for depression was moderate for studies in which participants had less than adequate immune functioning and viral suppression (0.56, 95%CI 0.42-0.70, z=7.77, \( p<0.001; \, \chi^2=0.0\% \)) and for studies in which these indicators were uncertain (0.53, 95%CI 0.33-0.73, \( z=5.28, \, p<0.001; \, \chi^2=69.1\% \)). There were no studies with ‘adequate’ functioning on these indices in this analysis. Studies from the pre-cART (0.49, 95%CI 0.32-0.66, \( z=5.75, \, p<0.001; \, \chi^2=0.0\% \)), cART (0.73, 95%CI 0.37-1.10, \( z=3.91, \, p<0.001 \)), and current cART era (0.55, 95%CI 0.38-0.72, \( z=6.21, \, p<0.001; \, \chi^2=62.5\% \)) all had moderate or near moderate pooled SMDs. Lastly, pooled SMDs for depression were similar in studies from the US which has a mixed healthcare system (0.57, 95%CI 0.43-0.72, \( z=7.58, \, p<0.001; \, \chi^2=55.7\% \)) and one study from the UK, which has universal healthcare (0.68, 95%CI 0.09-1.28; \( z=2.25, \, p=0.03 \)). The pooled SMD of two studies conducted in lower- and middle-income countries was smaller, but statistically significant (0.43, 95%CI 0.14-0.72, \( z=2.90, \, p=0.004; \, \chi^2=54.2\% \)).

**Psychological Distress**

Eighteen studies examined variables representing psychological distress, including anxiety-related constructs and the presence of ‘mental illness’, which generally described a combination of anxiety and depression. Five cross-sectional studies provided events data (Figure 4). The pooled odds ratio was significant and indicated that participants with pain were more likely to have psychological distress than those without pain: \( OR=2.56 \, (95\% \, CI \, 1.67-3.90), \, z=4.34, \, p<0.001 \). Heterogeneity was high (\( \chi^2=68.3\% \)). One prospective study (n=127) found that baseline mental illness did not predict presence of pain over follow-up [54]. Seven cross-sectional studies provided means and standard deviations (Figure 5). The pooled SMD showed a large and statistically significant difference between groups such that distress was worse in participants with versus without pain (SMD=0.85, 95% CI 0.35-
1.35); $z = 3.33$, $p = 0.001$). Heterogeneity was very high ($I^2 = 95.4\%$). One further study which reported the median and interquartile range found no difference between groups [88].

Four cross-sectional studies reported correlations between distress and pain severity (Supplemental Figure 9, available at http://links.lww.com/PAIN/A643). The pooled correlation was moderate: Fisher’s $z=0.35$ (95% CI 0.09-0.60), $z=2.68$, $p=0.007$. Heterogeneity was high (82.2%). Three cross-sectional studies reported correlations between distress and pain interference/disability. The pooled correlation was moderate: Fisher’s $z=0.59$ (95% CI 0.24-0.93), $z=3.34$, $p=0.001$. Heterogeneity was high (81.2%). One prospective study (n=45-62) found a non-significant correlation between change in distress and pain severity following cognitive-behavioural therapy, and a significant, moderate correlation between change in distress and pain interference [44,86]. Lastly, one cross-sectional study reported a non-significant correlation between pain intensity and distress ($r$ not reported), and a small negative correlation between distress and quality of life [88].

**Post-Traumatic Stress**

Three studies investigated post-traumatic stress. These studies are reported separate from studies measuring psychological distress, given the specificity of post-traumatic stress as a variable. Different study designs and analyses precluded meta-analysis.

One prospective study (n=143) found that post-traumatic stress symptoms (PTSS) were associated with significantly higher pain severity and interference over time in a sample with HIV and persistent pain [107]. One high quality cross-sectional study found that participants with pain (n=170) had significantly higher PTSS than those without pain (n=59) [79]. Within the pain group in this study, there was a non-significant correlation between PTSSs and pain severity, and small but significant correlations between PTSSs and pain...
interference (positive correlation) and quality of life (negative correlation) [79]. Post-traumatic stress disorder did not differ between groups with (n=150) and without (n=128) neuropathy in another cross-sectional study [31].

**Drug Abuse**

Fourteen studies examined drug abuse. We prioritised extracting opioid abuse data when multiple drug abuse categories were reported given the relevance of opioid use in chronic pain. Two prospective studies reported odds ratios for pain predicting heroin use at the time of follow-up. The pooled odds ratio indicated that participants with pain at baseline were more likely at follow-up to be using heroin: OR=1.70 (95%CI 1.22-2.38), z=3.13, \( p=0.002 \) (Supplemental Figure 10, available at http://links.lww.com/PAIN/A643). Heterogeneity was low (\( I^2 = 14.0\% \)). Conversely, another prospective study (n=493) reported that baseline opioid use disorder history predicted new onset of neuropathic pain, OR=2.87 (1.31-6.28), \( p<0.01 \) [60]. One low quality prospective study (n=127) found that baseline drug abuse history did not predict the presence of pain at follow-up, 0.55 (0.25-1.21) [54]. These two studies could not be combined due to different coding of the dependent variable.

Eight cross-sectional studies reported events data (Figure 6). The pooled odds ratio was significant such that participants with pain were more likely to have co-morbid drug abuse than those without pain: OR=1.59 (95%CI 1.12-2.26), z=2.58, \( p=0.01 \). Heterogeneity was high (\( I^2 = 69.8\% \)), mainly attributable to one study which found the opposite effect, such that participants with symptomatic distal sensory polyneuropathy (DSP) were less likely to have opioid use disorder than those with asymptomatic DSP [76]. One low quality cross-sectional study (n=503) found that participants with ‘untreated’ pain had greater dependence symptoms than those with ‘treated’ pain or without pain [110]. Another low
quality cross-sectional study (n=73) found a small positive correlation between ‘aberrant drug behaviours’ and pain interference, but not quality of life [81].

Alcohol Abuse

Eleven studies investigated alcohol abuse. Two prospective studies reported odds ratios for baseline pain predicting subsequent alcohol abuse. The pooled odds ratio was not significant: OR=0.94 (95% CI 0.39-2.26), z=0.13, p=0.90 (Supplemental Figure 11, available at http://links.lww.com/PAIN/A643). Heterogeneity was high (84.1%). Two additional prospective studies examined baseline alcohol abuse as a predictor of developing pain/neuropathy but could not be combined due to different analyses. Both studies reported a non-significant association between these variables [60,77]. Seven cross-sectional studies provided events data or odds ratios. The pooled odds ratio was not significant: OR=1.22 (95%CI 0.92 to 1.62), z=1.36, p=0.17 (Supplemental Figure 12, available at http://links.lww.com/PAIN/A643). Heterogeneity was medium (I^2=39.0%).

Sleep Disturbance

Three studies investigated sleep disturbance. Two cross-sectional studies reported means and standard deviations. The pooled effect was significant and showed moderately greater sleep problems in participants with versus without pain: SMD=0.66 (95% CI 0.45-0.87), z=6.12, p<0.001. Heterogeneity was 0.0% (Supplemental Figure 13, available at http://links.lww.com/PAIN/A643). Another cross-sectional (n=45) study reported a significant correlation between pain severity and sleep disturbance, and a non-significant correlation between sleep and functioning [94].
ART Non-Adherence

Seven studies investigated the association between pain and suboptimal ART adherence. Data were analysed separately according to whether the adherence variable was coded in the direction of non-adherence or adherence. One prospective study (n=258) reported that severe pain at baseline predicted higher odds (OR=1.37, 95%CI 1.02 to 1.85) of follow-up ART non-adherence [46]. One cross-sectional study provided events data while another provided an odds ratio. The pooled odds ratio was significant and indicated that participants with pain were more likely to report non-adherence: OR=1.40 (95% CI 1.07 to 1.82), z=2.50, p=0.01 (Supplemental Figure 14, available at http://links.lww.com/PAIN/A643). Heterogeneity was 0.00%. One cross-sectional study (n=42) found significant positive correlations between pain severity and adherence forgetfulness and fears [57].

Two cross-sectional studies reported data for the association between pain and adherence (events data or odds ratio). The pooled odds ratio was less than one, indicating the likelihood of adherence was lower in participants with pain, but this was not statistically significant: OR=0.32 (95% CI 0.08 to 1.32), z=1.57, p=0.12 (Supplemental Figure 15, available at http://links.lww.com/PAIN/A643). Heterogeneity was high (74.3%). Lastly, one low-quality cross-sectional study (n=377) found that pain presence was not associated with adherence in a structural equation model [72].

Healthcare Use

Six studies examined healthcare use. One prospective study (n=1521) found that baseline pain predicted significantly higher odds (OR=1.6, 95%CI 1.2-2.0) of urgent care visits [68]. Two cross-sectional studies reported events data. The pooled odds ratio was not
significant: OR=0.98 (95% CI 0.58-1.66, z=0.07, \( p=0.94 \)) (Supplemental Figure 16, available at http://links.lww.com/PAIN/A643). Heterogeneity was 0.0%. Two further cross-sectional studies reported means and standard deviations. The pooled effect was small but significant, such that participants with pain had greater healthcare use than those without pain: SMD=0.36 (95% CI 0.21-0.51), \( z=4.66, \ p<0.001 \) (Supplemental Figure 17, available at http://links.lww.com/PAIN/A643). One further cross-sectional study (n=1120) found that participants with pain and daily opioid use had more clinic visits than those with pain without daily opioid use and those without pain (standard deviation not reported) [53].

**Missed HIV Clinic Visits**

Two prospective studies reported odds ratios for baseline presence of pain predicting missed HIV clinic visits over one year follow-up. The pooled odds ratio was significant, such that those with pain at baseline had higher odds of a missed HIV clinic visit: OR=1.42 (95% CI 1.13-1.79), \( z=2.98, \ p=0.003 \) (Supplemental Figure 18, available at http://links.lww.com/PAIN/A643). Heterogeneity was 0.0%.

**Unemployment**

Seven cross-sectional studies provided events data or odds ratios for the association between unemployment and pain. The pooled odds ratio was significant, and indicated that participants with pain had higher odds of being unemployed than those without pain: OR=2.09 (95% CI 1.59-2.76), \( z=5.25, \ p<0.001 \) (Supplemental Figure 19, available at http://links.lww.com/PAIN/A643). Heterogeneity was moderate (48.6%). One further cross-sectional study (n=229) that did not have data available for meta-analysis likewise found
that participants with pain were significantly more likely to be unemployed than those without pain [79].

**Protective Factors**

Five studies examined protective psychological factors. One prospective study (n=62) found significant small and medium correlations between change in self-reported pain acceptance during CBT and post-treatment pain severity and interference, respectively [86]. One cross-sectional case-control study observed lower resilience in participants with (n=99) versus without pain (n=98; medium effect); however, this study found non-significant correlations between resilience and pain severity and interference in the pain group [121].

One high quality cross-sectional study found that participants with pain (n=170) reported lower disease management self-efficacy than did those without pain (n=59) (small effect) [79]. Within the pain group in this study, there were non-significant correlations between self-efficacy and pain severity and interference, and a small positive correlation between self-efficacy and quality of life [79]. One low quality cross-sectional study found that those with greater adherence self-efficacy were less likely to report pain (n=70) [8]. Lastly, one low quality cross-sectional study found lower mean self-reported optimism in participants with (n=50) versus without pain (n=46) (small effect) [101].

**Social Factors**

Four studies investigated social factors. The BEACON study (n=377) explored social processes across three papers, two of which describe prospective data (medium quality) while the third reported cross-sectional data (low quality). Baseline chronic pain predicted “negative social support” (i.e., overly intrusive or insensitive responses from others and a
lack of support) at 12 months, controlling for baseline social support [73]. Another prospective analysis showed that no chronic pain at baseline predicted greater support reciprocity at follow-up [74]. Chronic pain was associated with significantly poorer ratings of patient-provider engagement in cross-sectional analyses [72].

Two studies examined self-reported stigma, but could not be combined. One medium quality cross-sectional study (n=50) found a moderate positive correlation between stigma and pain severity [122]. One low quality cross-sectional study (n=201) found that participants with ‘Pain Disorder’ reported higher stigma scores than those without ‘Pain Disorder’ [98]. One medium quality cross-sectional study found no difference in mean number or quality of self-reported social supports between participants with (n=274) and without pain (n=164) [91].

Discussion

This review including over 13000 participants found ‘some’ or ‘moderate’ evidence supporting an association between pain outcomes and depression, psychological distress, post-traumatic stress, drug abuse, sleep disturbance, healthcare use, missed HIV clinic visits, ART adherence, unemployment, and protective psychological factors in people with HIV. Surprisingly few studies have examined protective psychological factors or social processes. There is a lack of high quality research on psychosocial factors related to chronic pain in people with HIV. These findings can inform future research and treatment development in this area.

The association between depression and poorer self-reported pain outcomes in HIV is consistent with the wider pain literature [4,64]. Data from prospective studies suggest depression is a risk factor for pain. However, caution is warranted in this interpretation
given the observational nature of studies. There is likely a bi-directional relationship, with shared neurobiological pathways, cognitive appraisals, and behavioural disengagement underpinning this association [4,7,14]. Evidence supporting the association between pain and sleep disturbance is consistent with the wider literature that reports reciprocal associations between pain, sleep, and depression [106].

There was substantial variability in the assessment of ‘psychological distress’, which may have contributed to the statistical heterogeneity observed. Although different measures were used to assess variables such as pain catastrophizing, pain-related fear, stress, and general anxiety, these measures overlap conceptually and in item content. The consistency of results within the psychological distress category suggests the findings are robust across different assessment methods. Several studies assessed ‘mental illness’ based on a range of diagnoses in participants’ medical file without clear diagnostic criteria. Studies exploring mental health diagnoses should use valid and reliable criteria and, ideally, semi-structured clinician-administered interviews as the gold standard (e.g., [128]). In light of high rates of PTSD in HIV [99], further research is particularly needed to understand the role of PTSD in pain in this context. Alternately, rather than focusing on specific mental health diagnoses, research investigating psychosocial processes that explain the impact of a range of psychological difficulties may prove useful moving forward [42].

Few studies investigated fear-avoidance model (FAM) variables, such as pain catastrophizing and pain-related fear, which have dominated the musculoskeletal pain literature. Fear of movement is strongly associated with musculoskeletal pain disability [20]. However, neuropathic pain is often spontaneous and not clearly provoked by movement, although it may inhibit movement. Therefore, research is needed to determine the relevance of FAM constructs in neuropathic pain which is common in HIV [19].
A bi-directional association between drug abuse and pain in HIV is suggested by prospective data showing opioid abuse as both predictor [60] and outcome [114] of pain. In a population where there are concerns about analgesic prescribing [58], poorly managed pain may contribute to increased abuse of non-prescribed opioids, which may be exacerbated by depression [114]. Alternately, prolonged opioid abuse may disrupt descending pain inhibition, exacerbating pain [49]. Differing definitions of drug and alcohol abuse across studies may account for variability in effects. Future research on substance abuse in this context should use validated assessments, either screening tools or diagnostic interviews that capture key features of abuse, such as continued use despite harm [1,41,128].

Adherence to ART and retention in care are psychosocial factors unique to the HIV context, and are of vital importance given their associations with mortality, morbidity, and drug resistance [112]. The finding that pain was associated with reduced ART adherence and missed HIV clinic visits highlights the necessity of adequate pain management in HIV. Understanding the links between pain, ART adherence, and retention in care likely requires consideration of other psychosocial factors, such as substance abuse and depression, which may mediate or moderate this association [68,92].

Findings that pain was associated with greater unemployment and healthcare use highlight the individual, societal, and economic costs of pain in HIV. This is consistent with the broader literature, although healthcare use is typically under-assessed in trials of psychotherapy for chronic pain [87]. Studies assessing healthcare use were restricted to the US, while the unemployment-pain link was consistent in studies from the US, Russia, and South Africa. The association between pain and healthcare use differed across studies on the basis of the type of healthcare assessed. Assessment of the most frequently accessed
services (e.g., GP visits), rather than relatively infrequent events (e.g., hospitalisations) may increase the interpretability of future healthcare data.

Surprisingly, only five studies assessed protective psychological factors. The lack of studies on protective factors mirrors historical trends in the general pain literature, although there has been greater focus on protective factors more recently. The focus on ‘maladaptive’ responses to pain is problematic, as such responses can be understood as a function of their short-term utility [126]. Moreover, abnormal conceptualizations often fail to specify pathways through which recovery and successful functioning occur when pain is present. The psychological flexibility model, within which pain acceptance has been conceptualised, might prove useful for future research [63].

A recent proposal for updating the definition of pain highlights the central role of social factors [127]. However, our review identified only four studies exploring interpersonal variables. The lack of research on stigma in relation to pain is particularly surprising as managing stigma is key to the success of the HIV/AIDS response [108]. Stigma has recently been highlighted as important for the well-being of patients with chronic pain in general [23,125]. Future research is needed to determine the function of stigma in chronic pain in people with and without HIV.

The study samples included in our review varied widely in terms of the proportion of men and women, participant age, ethnicity, and duration and severity of HIV. Our sensitivity analyses support the potential applicability of findings across pain types, ART treatment eras, and healthcare systems. Due to poor reporting of viral loads and CD4+ counts, our analysis stratifying by these indicators is difficult to interpret. Caution is also warranted regarding the cross-cultural applicability of the findings, as most studies were from the US. One South African study with a predominantly female sample found patients with and
without pain did not differ on depression or anxiety, likely due to high scores across the sample [88]. Socioeconomic factors, such as poverty and gender, may thus alter the relationships between pain, functioning, and mental health [88,121]. Care is needed in applying Western psychological concepts in non-Western cultures [51,82,83]. Research must also determine unique cultural features that influence the experience and expression of pain in HIV.

A guiding theoretical model is needed to integrate psychosocial processes relevant to pain and HIV. Such a model should make specific predictions about the relative contributions of cognitive, affective, behavioural, and sociocultural processes in relation to specific pain outcomes. This review identified a number of closely related psychosocial constructs. Therefore, a theoretical model may benefit from identifying a key set of non-overlapping variables [63]. This may draw on prominent models within the field of pain, such as the fear-avoidance [19] and psychological flexibility models [63], and those within the HIV literature that focus heavily on sociocultural perspectives to understand the impact of processes, such as stigma, on well-being [78].

The current findings suggest the relevance of psychosocial treatments to manage persistent pain in HIV. To our knowledge, only three small RCTs have evaluated cognitive-behavioural therapy (CBT) and mindfulness-based treatment [29,35,118]. Non-randomized trials of CBT [113] and hypnosis [25] have also been conducted. Further evaluation of psychosocial treatments for HIV and chronic pain is thus needed. The development of treatments that specifically target psychosocial factors identified in this review with ‘some’ or ‘moderate’ evidence may prove fruitful.

Several limitations warrant consideration. We used a comprehensive search strategy that included efforts to identify grey literature to limit publication bias; however, relevant
studies may have been missed given the broad nature of the search. We used an adapted quality assessment tool. Although we based this on previously validated tools, the adaptations may have limited the reliability and validity of our quality assessment. Assessment of pain was inadequate in many studies. Future research should assess information regarding pain duration, intensity, location, and type. Studies investigating chronic pain should specify eligibility criteria in line with recognized definitions: the presence of daily, clinically meaningful pain intensity and functional interference for at least three months [10,26,131]. Given the relevance of neuropathic pain in this population, the use of well-validated screening tools of neuropathy signs and symptoms is important [16,33,129].

This review identified a large number of psychosocial factors. As evidence on specific psychosocial factors develops in this area, it may be useful for a future review to use a more targeted approach to synthesize data on a smaller number of pre-specified variables. We focused on bivariate analyses and dichotomized multiple between-groups analyses to facilitate comparison across studies and minimise pairwise comparisons. However, this may have limited an in-depth understanding of psychosocial factors from multivariate models and more subtle subgroup analyses. Future research examining the association between psychosocial factors and pain outcomes in HIV should consider controlling for such variables as age, sex, race/ethnicity, socioeconomic status, HIV duration, current and nadir CD4+ count and current and peak viral load, current and past ART regimens, and other medical comorbidities (e.g., hepatitis C, diabetes, tuberculosis). Where multiple psychosocial variables are included, sufficient rationale for each variable should be provided and care should be taken to minimize overlap in assessment content between variables.
Despite these limitations, this is the first systematic review to explore psychosocial variables associated with persistent pain in HIV. From this review it is recommended that researchers (a) focus greater attention on protective psychological factors and social processes, such as stigma and processes to undermine stigma; (b) use higher quality assessment tools; and (c) develop and test treatments to target key psychosocial factors to improve pain outcomes in HIV. Improving quality of life is a priority as people with HIV live longer. Adequate, whole-person pain management is vital to achieve this goal.

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Supplemental video content

Video content associated with this article can be found at http://links.lww.com/PAIN/A645.

Figure Legend

Figure 1. Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA [75]) flow diagram.

Figure 2. Forest plot of cross-sectional odds ratios (OR) for depression. Depression was more likely in participants with versus without pain, as reflected in the pooled odds ratio (OR) of >1. Note: CI, Confidence Interval. Grey boxes show weighting of individual studies; the red dotted line indicates the pooled effect around which effects from individual studies vary; the blue diamond shows the 95% CI around the pooled effect.
Figure 3. Forest plot of cross-sectional standardized mean differences (SMD) for depression. Depression symptoms were more severe in participants with versus without pain, as indicated by a positive pooled SMD. Note: CI, Confidence Interval; SD, standard deviation.

Figure 4. Forest plot of cross-sectional events data for psychological distress. Distress was more likely in participants with versus without pain, as reflected in the pooled odds ratio (OR) of >1. Note: CI, Confidence Interval.

Figure 5. Forest plot of cross-sectional standardized mean differences (SMD) for psychological distress. Distress was more severe in participants with versus without pain, as reflected by a positive pooled SMD. Note: CI, Confidence Interval; SD, standard deviation.

Figure 6. Forest plot of cross-sectional events data for drug abuse. Drug abuse was more likely in participants with versus without pain, as reflected in the pooled odds ratio (OR) of >1. Note: CI, Confidence Interval.

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Table 1. Summary of the evidence for psychosocial variables.

<table>
<thead>
<tr>
<th>Study Quality</th>
<th>Psychosocial Assessment</th>
<th>Study Design/Analyses</th>
<th>Main Summary</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Depression</strong></td>
<td></td>
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<tr>
<td>High: (n=2)</td>
<td>BDI* (n=11); BSI (n=2)</td>
<td>-Prospective (HR) (n=2; Fig S1)</td>
<td>Depression consistently positively associated with pain presence, intensity, and interference, and negatively with quality of life in prospective and cross-sectional studies. Evidence level: Moderate</td>
</tr>
<tr>
<td>[79,84]</td>
<td>CDC (n=1); CES-D (n=6)</td>
<td>-Prospective (OR) (n=2; Fig S2)</td>
<td></td>
</tr>
<tr>
<td>Medium: (n=19)</td>
<td>CIDI* (n=1); DAPOS (n=1)</td>
<td>-Cross-sectional (OR) (n=9, includes 1 cross-sectional case-control; Fig 2)</td>
<td></td>
</tr>
<tr>
<td>[2279,31,50,6</td>
<td>DASS (n=1); GAIN (n=1)</td>
<td>-Cross-sectional (SMD) (n=13; Fig 3/text)</td>
<td></td>
</tr>
<tr>
<td>061,67,88,89,91,92,96,97,102,104,114,116,119]</td>
<td>HADS (n=1); HSC (n=1)</td>
<td>-Cross-sectional (correlation) (n=8; Fig S8/text)</td>
<td></td>
</tr>
<tr>
<td>Low: (n=8)</td>
<td>PHQ (n=2); PRISM (n=1)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>[8,28,39,55,57,95,101,110]</td>
<td>Non-validated (n=1)</td>
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<tr>
<td><strong>Psychological Distress</strong></td>
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<tr>
<td>High: (n=1)</td>
<td>BSI (n=2); CSQ (n=1)</td>
<td>-Prospective (OR) (n=1; in text)</td>
<td>Distress consistently positively associated with pain presence, intensity, and interference in cross-sectional analyses, but this was not consistent in the two prospective studies. Evidence level: Some</td>
</tr>
<tr>
<td>[84]</td>
<td>GAIN (n=1); HADS (n=1)</td>
<td>-Prospective (correlation) (n=1; in text)</td>
<td></td>
</tr>
<tr>
<td>Medium: (n=11)</td>
<td>HSC* (n=1); INTRP (n=1)</td>
<td>-Cross-sectional (OR) (n=5; Fig 4)</td>
<td></td>
</tr>
<tr>
<td>[2,29,31,44,47,50,61,66,88,89,91,104]</td>
<td>PASS (n=1); PCS* (n=4)</td>
<td>-Cross-sectional (SMD) (n=8; Fig 5/text)</td>
<td></td>
</tr>
<tr>
<td>Low: (n=6)</td>
<td>PHQ (n=1); POMS (n=1)</td>
<td>-Cross-sectional (correlation) (n=5; Fig S9)</td>
<td></td>
</tr>
<tr>
<td>[5,53,54,57,65,110]</td>
<td>PRISM (n=1)</td>
<td></td>
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<tr>
<td><strong>Post-traumatic Stress (PTSS)</strong></td>
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<tr>
<td>High: (n=1)</td>
<td>HTQ (n=1)</td>
<td>-Prospective (ANCOVA) (n=1; text)</td>
<td>PTSS associated with poorer pain outcomes in one prospective and one high quality cross-sectional study. Evidence level: Some</td>
</tr>
<tr>
<td>[79]</td>
<td>PCL-C (n=1)</td>
<td>-Cross-sectional (SMD and correlation) (n=1; text)</td>
<td></td>
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<tr>
<td>Medium: (n=2)</td>
<td>PRISM (n=1)</td>
<td>-Cross-sectional (OR) (n=1; text)</td>
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<tr>
<td>[31,107]</td>
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<tr>
<td><strong>Drug Abuse</strong></td>
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<tr>
<td>Medium: (n=9)</td>
<td>CIDI (n=1); DIS-IV (n=1)</td>
<td>-Prospective (OR) (n=4; Fig S10/text)</td>
<td>Pain and drug abuse consistently positively associated in prospective and cross-sectional studies. Evidence level: Moderate</td>
</tr>
<tr>
<td>[47,52,60,67,76,92,104,114,119]</td>
<td>GAIN (n=1); PRISM (n=1)</td>
<td>-Cross-sectional (OR) (n=8; includes 1 cross-sectional case-control; Fig 6)</td>
<td></td>
</tr>
<tr>
<td>Low: (n=5)</td>
<td>RBS (n=1)</td>
<td>-Cross-sectional (SMD) (n=1; text)</td>
<td></td>
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<tr>
<td>[39,53,54,81,110]</td>
<td>Medical file (ICD-9; n=1; unclear, n=2)</td>
<td>-Cross-sectional (correlation) (n=1; text)</td>
<td></td>
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<tr>
<td><strong>Alcohol Abuse</strong></td>
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<tr>
<td>Medium: (n=8)</td>
<td>CIDI (n=2); DIS-IV (n=1)</td>
<td>-Prospective (HR) (n=1; text)</td>
<td>Lack of association between pain and alcohol abuse in cross-sectional studies; inconsistent findings in prospective studies. Evidence Level: Inconclusive</td>
</tr>
<tr>
<td>[27,52,60,76,77,115,117,119]</td>
<td>DSM-IV (n=1)</td>
<td>-Prospective (OR) (n=3; Fig S11/text)</td>
<td></td>
</tr>
<tr>
<td>Low: (n=3)</td>
<td>NIAAA (n=2); PRISM (n=1); SCID (n=1)</td>
<td>-Cross-sectional (OR) (n=7) (Fig S12)</td>
<td></td>
</tr>
<tr>
<td>[6,39,56]</td>
<td>Other self-report (n=2)</td>
<td></td>
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<tr>
<td></td>
<td>Unclear (n=1)</td>
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<tr>
<td><strong>Sleep Disturbance</strong></td>
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<td></td>
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<tr>
<td>High: (n=1)</td>
<td>GSDS (n=1)</td>
<td>-Cross-sectional (SMD)</td>
<td>Sleep disturbance and pain</td>
</tr>
<tr>
<td>[84]</td>
<td></td>
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<tr>
<td><strong>Medium:</strong> (n=2) [2,94]</td>
<td><strong>ISI</strong> (n=1)</td>
<td><strong>PSQI</strong> (n=1)</td>
<td>(n=2; Fig S13) -Cross-sectional (correlation) (n=1; text)</td>
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<tr>
<td><strong>ART Non-adherence</strong></td>
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<tr>
<td>Medium: (n=2) [66,68]</td>
<td>ACTG-AQ (n=5)</td>
<td>Other self-report (n=2)</td>
<td>-Prospective (OR) (n=1; text) -Cross-sectional (OR) (n=4; Fig S14/15) -Cross-sectional (correlation/SEM; n=2; text)</td>
</tr>
<tr>
<td>Low: (n=5) [8,46,57,72,110]</td>
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<tr>
<td><strong>Healthcare Use</strong></td>
<td></td>
<td></td>
<td>-Prospective (OR) (n=1; text) -Cross-sectional (OR) (n=3; Fig S17/text)</td>
</tr>
<tr>
<td>Medium: (n=4) [47,61,68,119]</td>
<td>Medical records (GP/HIV/urgent care visits) (n=4) Self-report (mental health treatment) (n=2)</td>
<td></td>
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<tr>
<td>Low: (n=2) [39,53]</td>
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<tr>
<td><strong>Missed HIV Clinic Visits</strong></td>
<td>Medical records</td>
<td></td>
<td>Prospective (OR) (n=2; Fig S18)</td>
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<tr>
<td>Medium: (n=2) [68,92]</td>
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<tr>
<td><strong>Unemployment</strong></td>
<td>WPAIQ (n=1)</td>
<td>Other self-report (n=7)</td>
<td>Cross-sectional (OR) (n=8, includes 1 cross-sectional case-control; Fig S19)</td>
</tr>
<tr>
<td>High: (n=1) [79]</td>
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<td>Medium: (n=7) [2,27,61,92,114,19,121]</td>
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<tr>
<td><strong>Protective Factors</strong></td>
<td>-ACTG-AQ self-efficacy item (n=1) -CPAQ (n=1) -Life Orientation Test (n=1) -The Resilience Scale (n=1) -SE-6 (n=1)</td>
<td></td>
<td>Prospective (correlation) (n=1) -Cross-sectional (OR) (n=1) -Cross-sectional (SMD/correlation) (n=3, includes 1 cross-sectional case-control) -All discussed in text</td>
</tr>
<tr>
<td>High: (n=1) [79]</td>
<td></td>
<td></td>
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<tr>
<td>Medium: (n=2) [86,121]</td>
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<td></td>
<td></td>
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<tr>
<td>Low: (n=2) [8,101]</td>
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<tr>
<td><strong>Social Factors</strong></td>
<td>-HASI-P (n=1) (stigma) -RSS (n=1) (stigma) -SSQ-SF (n=1) (support quality/number) Self-report: negative support/support reciprocity/patient-provider engagement (not validated; n=1)</td>
<td></td>
<td>Prospective (correlation) (n=1) Cross-sectional (OR) (n=1) Cross-sectional (SMD/correlation) (n=3) -All discussed in text</td>
</tr>
<tr>
<td>Medium: (n=3) [73,74,91,122]</td>
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<tr>
<td>Low: (n=2)[72,98]</td>
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</tbody>
</table>

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**Note:** The table above summarizes the evidence for pain associated with various factors, including ART non-adherence, healthcare use, missed HIV clinic visits, unemployment, and protective factors. The evidence level ranges from Some to Inconclusive, indicating varying levels of evidence support. The references in brackets denote the studies or measures used in the analysis.
**Note**: As justified in the Methods, we chose to focus on interpreting bivariate data. However, studies reporting a multivariate model between psychosocial and pain variables are bolded in column 1 for ease of reference. †Reference for multivariate analyses related to bivariate data reported in [29] is [38].

ACTG-AQ, AIDS Clinical Trials Group Adherence Questionnaire; BDI, Beck Depression Inventory; BSI, Brief Symptom Inventory; CES-D, Centre for Epidemiological Studies Depression Scale; CIDI, Composite International Diagnostic Interview; CPAQ, Chronic Pain Acceptance Questionnaire; CSQ, Coping Strategies Questionnaire; DAPOS, Depression, Anxiety, and Positive Outlook Scale; DASS, Depression Anxiety Stress Scales; DIS-IV, Diagnostic Interview Schedule for DSM-IV (Diagnostic and Statistical Manual of Mental Disorders, 4th Edition); GAIN, Global Appraisal of Individual Needs; GSDS; General Sleep Disturbance Scale; HADS, Hospital Anxiety and Depression Scale; HASI-P, HIV/AIDS Stigma Scale—People Living with AIDS; HSC, Hopkins Symptoms Checklist; HTQ, Harvard Trauma Questionnaire; ICD-9, International Classification of Disease mental illness/substance abuse codes extracted from medical file; INTRP, Inventory of Negative Thoughts in Response to Pain; ISI, Insomnia Severity Index; NIAAA, National Institute on Alcoholism and Alcohol Abuse Guidelines; PASS, Pain Anxiety Symptom Scale; PCL-C: PTSD Checklist-Civilian; PCS, Pain Catastrophizing Scale; PHQ, Patient Health Questionnaire Depression/Anxiety Module; POMS, Profile of Mood States—Tension/Anxiety; PRISM, Psychiatric Research Interview for Substance and Mental Disorders; PSQI, Pittsburgh Sleep Quality Index; RBS, Risk Behaviour Survey; RSS, Reece Stigma Scale; SCID, Structured Clinical Interview for DSM-IV; SE-6, Self-Efficacy for Managing Chronic Disease—6 Item Scale; SSQ-SF, Social Support Questionnaire—Short Form; WPAIQ, Work Productivity and Activity Impairment Questionnaire; †Patient report of diagnosis of mental illness by a clinician.

*Where different measures of the same variable were used for between-group and within group analyses in the same study, both measures are reported.

HR, Hazard Ratio; OR, Odds Ratio; SEM, Structural Equation Model; SMD, Standardized Mean Difference.
Records identified through
database searching
(n=4458)

Additional records identified
through other sources
(n=8)

Records after duplicates removed
(n=3733)

Records screened
(n=3733)

Records excluded
(n=3367)

Full-text articles assessed
for eligibility
(n=366)

Studies included in
qualitative synthesis
(n=46 studies across 67 papers; 63 published, 3 dissertations, 1 unpublished full-text)

Studies included in
quantitative synthesis
(meta-analysis)
(n=37)

Full-text articles excluded,
with reasons (n=299)
- Participants <18 years old (n=2)
- Not HIV or cannot extract HIV subsample (n=17)
- Not chronic pain, or cannot extract measure of pain/psychosocial association for (sub-)sample with pain (n=221)
- No modifiable psychosocial factors (n=36)
- Duplicate (n=14)
- Review or letter (n=7)
- Qualitative (n=2)
NOTE: Weights are from random effects analysis

- Overall (I-squared = 83.0%, p = 0.000)

<table>
<thead>
<tr>
<th>Study</th>
<th>N</th>
<th>OR (95% CI)</th>
<th>Weight</th>
</tr>
</thead>
<tbody>
<tr>
<td>Berg et al. (2009)</td>
<td>70</td>
<td>2.17 (1.01, 4.69)</td>
<td>12.17</td>
</tr>
<tr>
<td>Ellis et al. (2010)</td>
<td>881</td>
<td>1.61 (1.12, 2.33)</td>
<td>15.81</td>
</tr>
<tr>
<td>Fellows et al (2012)</td>
<td>278</td>
<td>0.89 (0.53, 1.49)</td>
<td>14.52</td>
</tr>
<tr>
<td>Hansen et al. (2011)</td>
<td>270</td>
<td>8.79 (2.01, 38.40)</td>
<td>6.76</td>
</tr>
<tr>
<td>Merlin et al. (2017)</td>
<td>140</td>
<td>125.96 (7.50, 2114.01)</td>
<td>2.58</td>
</tr>
<tr>
<td>Robbins et al. (2013)</td>
<td>254</td>
<td>2.70 (1.40, 5.00)</td>
<td>13.42</td>
</tr>
<tr>
<td>Safo et al. (2017)</td>
<td>851</td>
<td>4.72 (3.37, 6.59)</td>
<td>16.03</td>
</tr>
<tr>
<td>Simms et al. (1992)</td>
<td>58</td>
<td>9.72 (0.54, 176.63)</td>
<td>2.46</td>
</tr>
<tr>
<td>Tsui et al. (2013)</td>
<td>699</td>
<td>2.02 (1.49, 2.74)</td>
<td>16.24</td>
</tr>
</tbody>
</table>

Overall (I-squared = 83.0%, p = 0.000)

2.65 (1.62, 4.34) Weight 100.00
<table>
<thead>
<tr>
<th>Study</th>
<th>SMD (95% CI)</th>
<th>N, mean (SD); Pain</th>
<th>N, mean (SD); No Pain</th>
<th>%</th>
</tr>
</thead>
<tbody>
<tr>
<td>Aouizerat et al. (2010)</td>
<td>0.47 (0.25, 0.70)</td>
<td>175, 19 (10.7)</td>
<td>142, 14.2 (9.5)</td>
<td>9.05</td>
</tr>
<tr>
<td>Evans et al. (1998)</td>
<td>0.73 (0.37, 1.10)</td>
<td>40, 13.3 (8.1)</td>
<td>121, 7.7 (7.5)</td>
<td>8.10</td>
</tr>
<tr>
<td>Kirkland (2012)</td>
<td>2.26 (1.90, 2.62)</td>
<td>103, 27.2 (12.9)</td>
<td>93, 4.03 (6.12)</td>
<td>8.16</td>
</tr>
<tr>
<td>Mann et al. (2015)</td>
<td>0.16 (-0.30, 0.62)</td>
<td>78, 8.63 (3.8)</td>
<td>24, 8 (4.3)</td>
<td>7.41</td>
</tr>
<tr>
<td>Parker et al. (2017)</td>
<td>0.59 (0.29, 0.89)</td>
<td>170, 18.5 (10.4)</td>
<td>59, 12.6 (9)</td>
<td>8.57</td>
</tr>
<tr>
<td>Phillips et al. (2014)</td>
<td>0.68 (0.09, 1.28)</td>
<td>19, 11.2 (4.22)</td>
<td>29, 8.38 (4.1)</td>
<td>6.36</td>
</tr>
<tr>
<td>Rosenfeld et al. (1996)</td>
<td>0.53 (0.34, 0.73)</td>
<td>274, 19.6 (10.5)</td>
<td>164, 14.3 (9.2)</td>
<td>9.20</td>
</tr>
<tr>
<td>Saylor et al. (2017)</td>
<td>0.29 (0.04, 0.54)</td>
<td>76, 12.3 (10)</td>
<td>323, 9.6 (9.1)</td>
<td>8.90</td>
</tr>
<tr>
<td>Simmonds et al. (2005)</td>
<td>0.40 (-0.00, 0.81)</td>
<td>50, 19.5 (10.9)</td>
<td>46, 15.2 (10.3)</td>
<td>7.82</td>
</tr>
<tr>
<td>Singer et al. (1996)</td>
<td>0.38 (0.06, 0.70)</td>
<td>69, 1 (.9)</td>
<td>89, .7 (.7)</td>
<td>8.46</td>
</tr>
<tr>
<td>Surratt et al. (2015)</td>
<td>0.72 (0.54, 0.90)</td>
<td>258, 6.34 (2.64)</td>
<td>245, 4.3 (3)</td>
<td>9.28</td>
</tr>
<tr>
<td>Uebelacker et al. (2015)</td>
<td>0.94 (0.66, 1.22)</td>
<td>107, 14.8 (8.09)</td>
<td>107, 8.14 (5.96)</td>
<td>8.69</td>
</tr>
<tr>
<td>Overall (I-squared = 89.2%, p = 0.000)</td>
<td>0.68 (0.42, 0.93)</td>
<td>1419</td>
<td>1442</td>
<td>100.00</td>
</tr>
</tbody>
</table>

NOTE: Weights are from random effects analysis
<table>
<thead>
<tr>
<th>Study</th>
<th>OR (95% CI)</th>
<th>Pain Events</th>
<th>No Pain Events</th>
<th>Weight</th>
</tr>
</thead>
<tbody>
<tr>
<td>Fellows et al. (2012)</td>
<td>0.74 (0.28, 1.99)</td>
<td>8/150</td>
<td>9/128</td>
<td>12.11</td>
</tr>
<tr>
<td>Jiao et al. (2016)</td>
<td>2.69 (1.94, 3.73)</td>
<td>159/252</td>
<td>150/386</td>
<td>28.85</td>
</tr>
<tr>
<td>Koepppe et al. (2010)</td>
<td>2.13 (1.64, 2.77)</td>
<td>176/324</td>
<td>285/796</td>
<td>30.76</td>
</tr>
<tr>
<td>Merlin, Cen, et al. (2012)</td>
<td>4.11 (2.00, 8.44)</td>
<td>37/76</td>
<td>15/80</td>
<td>17.38</td>
</tr>
<tr>
<td>Merlin et al. (2015)</td>
<td>6.88 (2.38, 19.94)</td>
<td>13/30</td>
<td>7/70</td>
<td>10.90</td>
</tr>
<tr>
<td>Overall (I-squared = 68.3%, p = 0.013)</td>
<td>2.56 (1.67, 3.90)</td>
<td>393/832</td>
<td>466/1460</td>
<td>100.00</td>
</tr>
</tbody>
</table>

NOTE: Weights are from random effects analysis.
<table>
<thead>
<tr>
<th>Study</th>
<th>SMD (95% CI)</th>
<th>(SD); Pain</th>
<th>(SD); No Pain</th>
<th>Weight</th>
</tr>
</thead>
<tbody>
<tr>
<td>Aouizerat et al. (2010)</td>
<td>0.36 (0.14, 0.58)</td>
<td>175, 9.8 (7.4)</td>
<td>142, 7.3 (6.4)</td>
<td>14.96</td>
</tr>
<tr>
<td>Kirkland (2012)</td>
<td>2.78 (2.38, 3.17)</td>
<td>103, 34.9 (14.6)</td>
<td>93, 3.09 (6.32)</td>
<td>14.07</td>
</tr>
<tr>
<td>Mann et al. (2015)</td>
<td>-0.23 (-0.69, 0.22)</td>
<td>78, 10.1 (3.36)</td>
<td>24, 10.9 (4.1)</td>
<td>13.65</td>
</tr>
<tr>
<td>Phillips et al. (2014)</td>
<td>0.79 (0.20, 1.39)</td>
<td>19, 23.7 (12.6)</td>
<td>31, 14.1 (11.8)</td>
<td>12.67</td>
</tr>
<tr>
<td>Rosenfeld et al. (1996)</td>
<td>0.74 (0.54, 0.94)</td>
<td>274, 1.26 (.7)</td>
<td>164, .79 (.5)</td>
<td>15.05</td>
</tr>
<tr>
<td>Singer et al. (1996)</td>
<td>0.77 (0.45, 1.10)</td>
<td>69, 1 (.8)</td>
<td>89, .5 (.5)</td>
<td>14.47</td>
</tr>
<tr>
<td>Surratt et al. (2015)</td>
<td>0.74 (0.56, 0.92)</td>
<td>258, 5.84 (3.54)</td>
<td>245, 3.3 (3.3)</td>
<td>15.12</td>
</tr>
<tr>
<td>Overall (I-squared = 95.4%, p = 0.000)</td>
<td>0.85 (0.35, 1.35)</td>
<td>976</td>
<td>788</td>
<td>100.00</td>
</tr>
</tbody>
</table>

NOTE: Weights are from random effects analysis
<table>
<thead>
<tr>
<th>Study</th>
<th>OR (95% CI)</th>
<th>Events, Pain</th>
<th>Events, No Pain</th>
<th>%</th>
</tr>
</thead>
<tbody>
<tr>
<td>Hansen et al. (2011)</td>
<td>1.82 (0.59, 5.62)</td>
<td>75/240</td>
<td>4/20</td>
<td>6.66</td>
</tr>
<tr>
<td>Jiao et al. (2016)</td>
<td>2.23 (1.59, 3.14)</td>
<td>108/252</td>
<td>97/386</td>
<td>17.55</td>
</tr>
<tr>
<td>Koeppen et al. (2010)</td>
<td>1.36 (1.04, 1.78)</td>
<td>123/324</td>
<td>247/796</td>
<td>18.66</td>
</tr>
<tr>
<td>Merlin et al. (2017)</td>
<td>2.67 (1.30, 5.45)</td>
<td>36/63</td>
<td>22/66</td>
<td>11.26</td>
</tr>
<tr>
<td>Morgello et al. (2004)</td>
<td>0.22 (0.08, 0.63)</td>
<td>26/65</td>
<td>18/24</td>
<td>7.34</td>
</tr>
<tr>
<td>Safo et al. (2017)</td>
<td>1.57 (1.02, 2.42)</td>
<td>36/208</td>
<td>76/646</td>
<td>15.95</td>
</tr>
<tr>
<td>Singer et al. (1996)</td>
<td>2.49 (1.43, 4.33)</td>
<td>69/98</td>
<td>64/131</td>
<td>13.83</td>
</tr>
<tr>
<td>Uebelacker et al. (2015)</td>
<td>1.38 (0.55, 3.41)</td>
<td>12/107</td>
<td>9/107</td>
<td>8.76</td>
</tr>
<tr>
<td>Overall (I-squared = 69.8%, p = 0.002)</td>
<td>1.59 (1.12, 2.26)</td>
<td>485/1357</td>
<td>537/2176</td>
<td>100.00</td>
</tr>
</tbody>
</table>

NOTE: Weights are from random effects analysis